



GRE

R E S E A R C H R E P O R T

Ensuring the Fairness of GRE Writing Prompts: Assessing Differential Difficulty

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July 2005

**GRE Board Report No. 02-07R
ETS RR-05-11**

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GRE Board Research Report No. 02-07R

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The report presents the findings of a
research project funded by and carried
out under the auspices of the Graduate
Record Examinations Board.

Educational Testing Service, Princeton, NJ 08541

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Abstract

Three polytomous DIF detection techniques — the Mantel test, logistic regression, and polySTAND — were used to identify GRE[®] Analytical Writing prompts (“Issue” and “Argument”) that are differentially difficult for (a) female test takers; (b) African American, Asian, and Hispanic test takers; and (c) test takers whose strongest language skills are in a language other than English. Prompts in the gender comparison showed low differential item functioning (DIF) values. Higher DIF values were found in the White versus African American comparison (for Argument prompts), while a smaller number of prompts showed moderate DIF to the disadvantage of the Hispanic (for Argument prompts), Asian American (for Issue prompts), and English as a second language groups (for Issue prompts). No DIF values were found that were large enough to warrant the removal of prompts from the item pool. Several potential causes of high DIF values for some prompts are discussed with respect to the content characteristics of these prompts.

Key words: ESL, ethnic minorities, gender, GRE, item characteristics, polytomous DIF, prompts, writing assessment

Acknowledgments

We would like to thank Ourania Rotou for preparing the GRE dataset; John Willey for carrying out the statistical analysis for the Mantel test; Youn-Hee Lim for assisting us in preparing the datasets for logistic regression analysis; Mary Fowles and Cynthia Welsh for their help regarding the GRE prompt characteristics and the prompt review; Cecilia Rosenblum, Marlene Mujica-Fill, Sylvia Ledesma, Pablo Garcia Gomez, Luis Saldivia, Jean Williams, Mel Kubota, and Jennifer Bivens-Tatum for reviewing prompt characteristics in the ethnic minority group comparisons; and Paul Holland, John Donoghue, Neil Dorans, and Steve Sireci for their advice on DIF techniques.

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Introduction

Direct assessments of writing skill (i.e., essay tests such as the GRE Analytical Writing measure) have considerable appeal because, unlike their less direct, multiple-choice counterparts, they actually require examinees to write, not merely to recognize the conventions of standard written English. Along with their appeal, however, these tests carry a special burden that does not encumber traditional multiple-choice tests. The trade-off is that direct writing assessments typically require much more time per item and, as a result, typically contain far fewer items than do multiple-choice measures. They have, in fact, sometimes been characterized as one-item (or two-item) tests. This characterization is probably not entirely fair, as the elaborate responses elicited by essay prompts are invariably richer than those yielded by typical multiple-choice items.

Nonetheless, the point is well taken that much is demanded of a limited number of questions. Because so much depends on so few questions, the “luck of the draw” comes into play to a far greater extent for such assessments than it does for multiple-choice measures. Thus, it is incumbent on test developers to craft essay prompts that are fair to all examinees and as comparable to one another as possible. The hope is that no test taker will be unfairly disadvantaged by being administered a prompt whose content is so unfamiliar or uninspiring as to hinder the demonstration of analytical writing skills.

The GRE Program already does much to ensure the fairness of the Analytical Writing measure. Currently, prospective test takers are assured that the measure is fair because:

In creating this assessment for the GRE Board, Educational Testing Service (ETS®) [has] followed a rigorous test development process that was guided by faculty committees representing different academic institutions, disciplines, and cultural perspectives.... All topics have passed stringent reviews for fairness and, in national field test trials, [have] proved accessible and appropriate for entry-level graduate students across many disciplines and various cultural groups. Reading training and scoring procedures were designed to ensure scoring accuracy, fairness, and reliability, and to enhance the overall validity of the assessment (ETS, 2004).

Problem Description

Despite considerable effort to ensure equitable assessment, the GRE program's claims of fairness are often challenged simply because of mean differences in the performances of various subgroups of examinees. It should be noted, however, that simple differences in mean scores on an item across different examinee subgroups are not evidence of bias or unfairness. In some cases, examinees from two different groups may actually differ in the ability of interest, and differences in item performance are to be expected. These results are often called *item impact* (Ackerman, 1992; Clauser & Mazor, 1998). In fact, the real fairness issue should be the extent to which differential item functioning (DIF) is present in any of the GRE prompts. DIF occurs when examinees of equal ability, but with different group membership (e.g., male, female), have unequal probabilities of success on an item (Angoff, 1993; Clauser & Mazor, 1998; Hambleton, Swaminathan, & Rogers, 1991). To minimize the likelihood of this situation, test developers attempt to craft prompts that are as nearly equivalent as possible and thus, to the extent possible, ensure that all essay prompts function similarly for all test takers. In this way, any between-group difference in performance on a prompt will be due to construct-relevant factors rather than to influences that are irrelevant to the assessment of analytical writing ability. In addition, though, it is desirable to have some procedure — differential item functioning (DIF) — to detect when this condition has not been fully met.

The identification of a satisfactory DIF procedure for essay prompts is not an easy undertaking for a number of reasons. For multiple-choice items, well-researched methods exist for detecting items that are differentially difficult for certain subgroups of test takers. However, there are currently no entirely satisfactory, well-researched comparable procedures for determining when essay prompts are differentially difficult for matched subgroups of test takers.

One of the biggest problems in this regard is the absence of a reliable, internal criterion on which test takers can be matched with respect to the overall ability or skill being measured by the test. This overall matching must be accomplished before between-subgroup performance comparisons can be made for individual items. For standardized multiple-choice measures, the total score on the test serves this function. For direct writing assessments, a comparable internal matching criterion is not usually available. Unfortunately, this is the case for the GRE Analytical Writing assessment, which contains a single essay prompt for each of two task types (Issue and Argument). In such cases, the usual (less-than-satisfactory) strategy is to use an external

matching criteria — for example, scores on multiple-choice tests that measure similar knowledge, skills, or abilities. An even less attractive possibility is to use performance on one task (e.g., Argument) as the matching variable for the other task (e.g., Issue). Problems with this approach arise because one task may measure a somewhat different aspect of the construct than the other one does. In addition, performance on a single task does not usually constitute a reliable enough measure to allow sufficiently precise matching.

A further complication is that essay responses are scored polytomously, not simply as correct or incorrect, leaving open the possibility that, in addition to ability level, the item score level may also be a possible interacting variable (Dorans & Schmitt, 1993; French & Miller, 1996). That is, because of the polytomous nature of essay scores, DIF can occur in some or all score categories.

Prior Relevant Research

Several well-established DIF detection procedures for dichotomous items have been extended to polytomous items, including

- item response theory (IRT)-based procedures (Raju, 1988; Wainer, Sireci, & Thissen, 1991)
- the extended SIBTEST (Shealy & Stout, 1993)
- the standardization approach (Dorans & Kulick, 1986),
- the Mantel test of linear association — a generalization of the Mantel-Haenszel procedure that accommodates polytomous items¹ (Agresti, 1990; Mantel, 1963; Zwick & Thayer, 1996)
- logistic regression-based procedures (French & Miller, 1996; Swaminathan & Rogers, 1990; Zumbo, 1999)
- logistic discriminant function analysis (Miller & Spray, 1993)
- analysis of covariance (Sireci, Harter, Yang, & Bhola, 2003; Sireci, Yang, Harter, & Ehrlich, 2004)

Penfield and Lam (2000) conducted a comparative evaluation of these DIF methods for polytomously scored items and recommended using a combination of methods that included the standardization method and logistic regression as the most favorable approach to DIF investigation for performance assessments (see also Potenza & Dorans, 1995, for a

comprehensive review). In this study, we used two of the methods recommended by Penfield and Lam (2000) as well as the Mantel test, a method that has been widely researched for multiple-choice items.

Welch and Miller (1995) studied different DIF detection procedures and different combinations of external and internal matching variables in investigations of DIF for essay items. One of the most important findings from this study was that, in most cases, combining a multiple-choice test score of related content with one essay score as the matching criterion reduced the apparent degree of DIF substantially (when compared to a criterion made up of only scores on the multiple-choice items), and even further when a second essay score was added. The researchers argued that the most likely explanation for this DIF pattern concerned potential construct differences between the multiple-choice and the essay items. Under such circumstances, using only multiple-choice test scores as a matching variable, or some other matching variables dominated by multiple-choice test scores, does not match examinees satisfactorily on the very construct measured by the essay items. As a result, the ability of the DIF methodology to separate DIF from impact diminishes. In this sense, the major obstacle to applying currently available DIF methodology to performance assessments is the difficulty associated with finding appropriate matching variables that enable the separation of DIF from impact (i.e., mean score differences due to true ability differences between groups).

Another problem for DIF is that not all detection techniques (like the Mantel-Haenszel procedure) are sensitive enough to identify nonuniform DIF. Nonuniform DIF exists when the probability of answering the item correctly is not *uniformly* different across all levels of ability for the two groups being compared; that is, there is interaction between ability level and group membership (Mellenbergh, 1982). (Uniform DIF exists when the probability of answering an item correctly is uniformly greater for one group than another at all levels of ability.) One class of DIF detection techniques that can be used to examine nonuniform DIF is based on item response theory (e.g., Thissen, Steinberg, & Wainer, 1993; Williams, 1997). These methods, however, also require a reliable, internal matching criterion and, in addition, may pose some special problems when applying them to the GRE Writing Measure.² Two other procedures that are deemed effective for identifying both uniform and nonuniform DIF in polytomous items are (a) a logistic regression modeling procedure (French & Miller, 1996; Lee, Breland, & Muraki, 2004; Swaminathan & Rogers, 1990) and (b) a logistic discriminant function analysis (Miller &

Spray, 1993), a method that is similar to logistic regression but where the roles of the item score and group membership indicators are reversed. In this study, only the logistic regression procedure was used to investigate nonuniform DIF, because the results are conceptually easier to understand in relation to the standard definition of DIF.

Lee et al. (2004) employed logistic regression to study whether matched examinees with non-Indo-European native language backgrounds are unduly disadvantaged on some writing prompts used in the Test of English as a Foreign Language™ (TOEFL®) when compared to examinees with European native language backgrounds. For their matching variable, Lee et al. created an English language ability score by standardizing and summing TOEFL Reading, Listening, and Structure scores. Approximately one-third of the prompts analyzed exhibited statistically significant matched group differences ($p < .05$). The authors concluded, however, that the effect size (i.e., R^2 difference between the nested models) was far too small, practically speaking, for any prompt to be classified as being differentially difficult for any of the matched groups studied. In their research, Clauser, Nungester, Mazor, and Ripkey (1996) studied both the Mantel-Haenszel procedure and logistic regression. Best results were obtained when using a logistic regression model based on group membership and four relevant subtests. The model was specified so that each of the four subtest scores was forced to enter the equation for each item.

Objectives

The aim of the research reported here was to evaluate the usefulness of several alternative DIF methods for detecting GRE essay prompts that may function differently for specified subgroups of matched GRE test takers. Each of the two prompt types used in the GRE Writing assessment (Issue and Argument) were analyzed. An attempt was also made to compare the impact on DIF estimates of using different matching variables created by combining multiple-choice test scores and essay scores.

Methods

Matching Variables

Two main matching variables were used in the analysis described here. Both entailed a composite score of the verbal section plus the “other prompt” (i.e., when the Argument prompt was studied, the other prompt was the Issue prompt, and vice versa). It should be noted that when

the simple arithmetic mean of the verbal section and other prompt scores is used as a matching variable, the verbal score is, in effect, weighted far more heavily than is the other prompt because of the difference in valid score range for the two sections (i.e., 200–800 for verbal, 1–6 for essay). Thus, for the purpose of evaluating different weighting schemes, standardized (z) scores were created first for each component so that the scores for both sections had a mean of 0 and a standard deviation of 1. These z -scores were then weighted differentially for the two components in each of the two matching variables used in this study.

In determining the weights for the verbal section and essay scores in each of the two matching variables, both construct similarity and test length of each component were taken into consideration. Matching variable 1 was defined as follows:

$$\text{Matching variable 1} = (z_{\text{verbal}} + z_{\text{other prompt}})/2$$

This is simply the mean of the z -scores of the verbal section and the other prompt score in which the two components received approximately equal weights in spite of the different number of items used in each component. It should be noted that the “other essay prompt” component consisted of a single polytomously scored item, but this component can be viewed as representing the test construct of the studied prompt more closely than the multiple-choice verbal items. This weighing scheme also produced the highest correlations between the studied prompts and the matching variable.

However, one can argue that this formulation places too much weight on a single polytomously scored prompt (i.e., the essay score component). Therefore, we also created another matching variable in which the weight of the essay score component was significantly decreased, compared to the first matching variable. Matching variable 2 was defined as follows:

$$\text{Matching variable 2} = [z_{\text{verbal}} + (z_{\text{other prompt}}/7.6)]/2$$

This is simply the mean of the z -scores of the verbal scores and the z -scores of the other prompt scores divided by 7.6. The logic for the weighting applied in this second matching variable is as follows: There is one threshold in dichotomous items (0–1), whereas there are 5 threshold levels in essay items (1–2, 2–3, 3–4, 4–5, 5–6). We argue that an essay item should

therefore be “worth” 5 multiple-choice items in the verbal section. Since there are 38 items in the verbal section, the resulting weight adjustment is 7.6 (i.e., 38/5).³

Differential Item Functioning Techniques

The Mantel test. One of the most widely used methods for studying DIF in multiple-choice items is the Mantel-Haenszel procedure (Holland & Thayer, 1988; Mantel & Haenszel, 1959). In DIF investigation, we usually compare the performance of two groups of examinees: the focal group and the reference group. The focal group is a group of primary interest (e.g., females), whereas the reference group (e.g., males) is taken as a standard against which the focal group is compared (Holland & Thayer, 1988). The Mantel-Haenszel procedure compares, for a given item, the probability of obtaining a right answer in the reference group to the probability of obtaining a right answer in the focal group, when the examinees in the two groups are matched on total test scores. Mantel’s (1963) test of conditional linear association (Agresti, 1990), which accommodates polytomous items, was used here to study DIF in the GRE Analytical Writing measure.

The Mantel test is capable of handling 2-by- j ($j > 2$) contingency tables. Assuming two groups (focal and reference) and a j -category ordered response variable (the essay scores), with observations at each of k ability levels, there are $2j$ cells at each of the k ability levels. Specifically, for the k^{th} matched set of members of the focal and reference groups at response level j on the studied item (essay prompt), there will be A_{jk} members in the focal group and B_{jk} members in the reference group. With the score Y_j assigned to the j^{th} level of response, a summary chi-square, with one degree of freedom, is computed as

$$m^2 = \frac{\left[\sum_k^K \sum_j^J A_{jk} Y_j - \sum_k^K E \left(\sum_j^J A_{jk} Y_j \right) \right]^2}{\sum_k^K \text{Var} \left(\sum_j^J A_{jk} Y_j \right)},$$

where for level k ,

$$\text{Var}\left(\sum_j^J A_{jk} Y_j\right) = \frac{n_{Fk} n_{Rk}}{T_k^2 (T_k - 1)} - \left[T \sum_j^J M_{jk} Y_j^2 - \left(\sum_j^J M_{jk} Y_j \right)^2 \right],$$

where T_k is the total number of examinees at level k , $m_{jk} = A_{jk} + B_{jk}$, and

$$E\left(\sum_j^J A_{jk} Y_j\right) = \sum_k^K n_{Fk} \frac{\sum_j^J (A_{jk} + B_{jk}) Y_j}{n_{Fk} + n_{Rk}}$$

(cf., Welch & Miller, 1995).

Logistic regression procedure. French and Miller (1996) and Zumbo (1999) demonstrated that the logistic regression procedure could be extended to studying polytomous DIF as well. Logistic regression has two main advantages over linear regression. The first is that the dependent variable need not be continuous, unbounded, or measured on an interval or ratio scale. The second is that it does not require a linear relationship between the dependent and independent variables.

The logistic regression method that was employed is the proportional odds-ratio model. A three-step modeling process based on logistic regression (Zumbo, 1999) was used as the main method of analysis for the sequential investigation of gender DIF. It should be noted that the logistic regression method was not used for racial/ethnic group DIF investigations due to sample size requirements. The R^2 differences between several nested models were used as an effect size measure, based on Jodoin and Gierl's (2001) and Zumbo's (1999) DIF classification schemes.

Since the final essay score is an average of two raters' ratings, the dependent variable in this study is in increments of 0.5, with 11 valid score categories (i.e., 1.0, 1.5, 2.0, 2.5, 3.0, 3.5, 4.0, 4.5, 5.0, 5.5, 6.0). Polytomous essay scores were dichotomized into 10 binary response variables according to the cumulative-logit dichotomization scheme (Agresti, 1990; French & Miller, 1996). For each essay prompt, these 10 binary variables were regressed on the 3 independent variables (i.e., the matching variable, the group variable, and the matching-score-by-group interaction term) in a stepwise fashion. The logistic regression equation for each dichotomized variable (i) for an essay prompt can be written as:

$$P(U_i | x, D) = \frac{\exp[g_i(x, D)]}{1 + \exp[g_i(x, D)]} = \frac{1}{1 + \exp[-(g_i(x, D))]} ,$$

where U_i represents the binary responses for a dichotomized response variable i ($U_i = 0$ or 1), x is the matching variable score that is treated as a continuous variable, and D is the design matrix of the covariate variables. In this equation, the function $g_i(x, D)$ is called a *logit*. The logit is a linear combination of the matching variable score (x), a covariate variable (D), and an interaction term (xD). If we want to analyze the DIF for M levels of an analysis group covariate (e.g., focal group = 0, reference group = 1), as in our GRE essay data, we can rewrite the logit $g_i(x, D)$ as:

$$g_i(x, D) = \beta_{0i} + \beta_1 x + \beta_2 D_m + \beta_3 x D_m ,$$

where β_{0i} is the intercept for a dichotomized variable (i); β_1 is the slope parameter associated with the matching variable score; β_2 is the parameter associated with the analysis group variable, D_m ; and β_3 is the slope parameter associated with the ability score-by-group interaction (cf., Lee et al., 2004).

Polytomous standardization (polySTAND) statistic. The Advanced Placement Program® (AP®) currently uses smoothed conditional means (that are conditioned on the scores of the multiple-choice section, free-response item section, or the composite of both) and polySTAND — an extension of the standardization approach (Dorans & Kulick, 1986) for polytomous DIF analysis. (Most of the AP tests involve a combination of dichotomously scored multiple-choice items and polytomously scored free-response items.) Graphical displays of smoothed conditional means as well as an overall index (polySTAND) are used to compare the performance of different groups on an essay (Potenza & Dorans, 1995; Tateneni, 2001).

PolySTAND is a summary statistic; it is the sum of differences in smoothed conditional means between groups 1 and 2, weighted by the relative frequency of individuals at each matching variable score level. The formula is given by:

$$\text{PolySTAND} = \frac{\sum_{i=1}^n n_i (\bar{x}_{i1} - \bar{x}_{i2})}{\sum_{i=1}^n n_i} ,$$

where n are the criterion score levels; \bar{x}_{i1} and \bar{x}_{i2} are the smoothed conditional means of groups 1 and 2 on the essay prompt at the i^{th} criterion score level; and n_i is the total sample size at the i^{th} criterion score level (cf., Tateneni, 2001).⁴

Sample and Dataset

The study relied on operational data from some 397,806 GRE General Test examinees who took the test between October 2002 and October 2003. Thus, the sample is quite representative of the full GRE test-taking population. In total, 117 Argument prompts and 109 Issue prompts were administered to these test takers. About 42,405 test takers (about 10.7% of the total sample) indicated that English was not their best language (non-EBL). For comparisons involving gender or race/ethnicity, only test takers who indicated that English was their best language (EBL) were included in the analysis. Of the 355,401 test takers who indicated that English was their best language, about 39% were male and about 60% were female, with some test takers of unidentified gender. Among the four racial/ethnic groups, the White group was the largest (about 61%) followed respectively by the African American (7%), Asian American (6%), and Hispanic American groups (2%). The dataset was cleaned for missing and out-of-range values for key variables. The analysis does not include essay scores of 0, as we believe that this score cannot be viewed as a score that represents a continuation of the 1 to 6 scale, but rather represents a qualitatively different score category (e.g., off-topic essays).

Procedure

Our initial assumption was that, because we lacked a purely internal matching variable, reliance on any single DIF detection method might yield biased results. Thus, basing the results on only one DIF detection method could further increase the risk of falsely flagging prompts. (Evidence will be presented later that this is probably a more serious problem than failing to identify prompts that *do* function differently.)

Because we did not believe that any of the DIF methods would yield unequivocal results, we combined the results of the DIF methods employed in this investigation and then focused on those prompts that exceeded values that are likely to have a practical impact on test performance (see the classification scheme below). As will be shown later, the methods were in considerable, though not perfect, agreement. The correlations among DIF values from the three different

methods ranged from .83 to .90 for the Argument prompts and from .86 to .89 for the Issue prompts.

Six different DIF values were computed for each prompt by using matching variables 1 and 2 with each of the three procedures (i.e., Mantel-Haenszel, logistic regression, and polySTAND). In each method, the DIF values for all the studied prompts based on the two matching variables were examined in terms of magnitude and direction. These prompts were then ordered from highest to lowest in absolute DIF values separately in each of the six conditions (3 procedures \times 2 matching variables), and rank numbers were assigned to each of them (ranking number 1 indicates the highest DIF). Then, an average rank was calculated from the ranks of both matching variables in each procedure. Finally, a mean rank of the procedures was calculated for those prompts that had large enough samples to be studied (or where DIF values were nonzero in the case of the logistic regression procedure).⁵

In order to make the polySTAND DIF measure comparable across different prompts and to convey a better idea of the size of its effect, the polySTAND value was divided by the combined standard deviation of the reference and focal groups. The polySTAND effect size (PS EZ) measure is independent of sample size, while the Mantel test is not. To correct for different categorizations based on different sample sizes of the prompts, the Mantel test measure (i.e., a summary chi-square statistic) was divided by the sample size of the focal group for ranking purposes.⁶ It should be also mentioned that *impact* values (standardized mean score differences between groups) were also computed for each prompt and compared with the PS EZ effect size values to examine the empirical relationship between these two indices for some of the racial/ethnic group comparisons (e.g., White vs. African American comparisons).

Although we did not attempt to flag prompts on the basis of our analysis, the following classification scheme, based on ETS’s polytomous DIF categorization used in the National Assessment of Educational Progress examination (Dorans & Schmitt, 1993), might be reasonably applied for PS EZ:

| Normal ETS categories | Flag code | Description |
|---------------------------------|-----------|-------------|
| $ \text{PS EZ} \leq .17$ | A | negligible |
| $.17 < \text{PS EZ} \leq .25$ | B | moderate |
| $ \text{PS EZ} > .25$ | C | large |

Given the absence of a purely internal matching variable (and the increased prospect of false positive results), however, this study can give no definitive answer regarding specific values that should be used. Nonetheless, a preliminary analysis suggested that DIF values tend to increase when less *construct-relevant*, non-essay components (e.g., multiple-choice verbal and quantitative sections) are added to the matching variable for the essay items, and vice versa (see also Summary and Discussion for more details). This is also consistent with previous research (Welch & Miller, 1995). Under such circumstances, one practical suggestion would be to increase the threshold for flagging GRE prompts in the following way:

| Categories for GRE prompts | Flag code |
|----------------------------|-----------|
| $ PS\ EZ \leq .25$ | A |
| $.25 < PS\ EZ \leq .35$ | B |
| $ PS\ EZ > .35$ | C |

The use of these less conservative criteria does make a difference with respect to conclusions about the prevalence of DIF in our study data. And, of course, an argument can be made for using the original, more conservative criteria. We decided, therefore, not to flag prompts until further investigation could be carried out about appropriate cut-off values (see the Summary and Discussion section), but rather to present a ranking of the prompts in each comparison, along with their respective PS EZ values, in order to give test developers more information regarding potentially problematic prompts. Using a matrix of major prompt characteristics developed by the GRE content team for the purpose of this study, we investigated whether or not certain prompt features were associated with higher DIF values (see the appendix).

Furthermore, we combined the psychometric analysis with a qualitative content review as follows. For the Argument prompts, we assembled two groups of 3 prompts each that included those with the highest DIF values, and two groups of 3 prompts each that showed no DIF (12 prompts in total) for each of the two racial/ethnic group comparisons (i.e., White vs. African American, White vs. Hispanic American).

For the Issue prompts, however, this review was performed only for the White versus Asian American comparison, because most of prompts flagged in this comparison were Issue prompts. We presented these sets to a group of ETS assessment experts that included at least one reviewer from the focal ethnic group (for example, reviewers from a Latino/Hispanic background reviewed DIF and non-DIF prompts in the White versus Hispanic group comparison). In order to discourage post hoc hypothesis formation, the reviewers were not told which prompts were identified as showing DIF. We decided not to present individual prompts to reviewers at this stage, because a trial review revealed that this task would be too difficult and too time-consuming. Every person in the group was asked first to read all the prompts in the four groups. Then, each one made a decision regarding which two prompt groups were the ones likely to show DIF based on the characteristics they identified. The decisions were then compared with the psychometric findings. Afterwards, the reviewers discussed the characteristics they believed distinguished the DIF prompts from the non-DIF prompts.

Some Notes on Conditions Affecting the Analysis

Sample size restrictions. The sample size was sufficient to study nearly all prompts for the male-female comparison. However, because of sample size restrictions, fewer prompts could be investigated for other focal groups.

Due to the large sample size requirements for logistic regression, this method could be employed only for the male-female comparison. For the White versus Ethnic group comparisons, and for the English as a best language (EBL) versus non-EBL comparison, only the polySTAND procedure and the Mantel test were used. Also, in these comparisons, some prompts were deleted from the analyses because they had not yet been administered to sufficiently large numbers of examinees.

Table 1 shows the minimum sample size requirements for the focal group that were imposed for the polySTAND procedure and the Mantel test.

Table 1***Minimum Sample Size Requirements and Percentage of Analyzed Prompts***

| Condition | Size of smaller subgroup | Percentage of prompts analyzed |
|-----------------------------|------------------------------|--------------------------------|
| Gender | | |
| Argument | Male ≥ 200 | 95.7% |
| Issue | Male ≥ 200 | 97.2% |
| Ethnic groups | | |
| <i>Argument</i> | | |
| African American vs. White | African American ≥ 200 | 67.5% |
| Hispanic American vs. White | Hispanic American ≥ 150 | 57.3% |
| Asian American vs. White | Asian American ≥ 150 | 74.4% |
| <i>Issue</i> | | |
| African American vs. White | African American ≥ 200 | 58.7% |
| Hispanic American vs. White | Hispanic American ≥ 150 | 45.9% |
| Asian American vs. White | Asian American ≥ 150 | 56.0% |
| EBL vs. non- EBL | | |
| Argument | non-EBL ≥ 150 | 92.3% |
| Issue | non-EBL ≥ 150 | 91.7% |

Levels of the matching variable. The number of categories (or intervals) into which the matching variable is divided should be fairly large. If the number is small (*fat matching*), the detection of DIF is confounded with impact (Donoghue & Allen, 1993). However, sample size also determines how many categories can be effectively used.⁷ Initially, we used too many categories, resulting in the assignment of too few focal group test takers per category. As a result, the different groups' score curves were very erratic. Only when fewer categories were used for prompts that had limited exposure did the score curves appear to be more or less monotonically increasing from one level to the next. The following rule was therefore applied to determine category size in the polySTAND procedure and the Mantel test (see Table 2).

Table 2

Number of Categories for Different Sample Sizes of the Focal Group

| Sample size of focal (i.e., smaller) group | Number of categories in the matching variable |
|---|--|
| 1,201 – 1,500 | 50 |
| 901 – 1,200 | 40 |
| 501 – 900 | 30 |
| 150 – 500 | 10 |

Fat matching. To evaluate the impact of so-called fat matching, we carried out the following comparison in the October 2002 to April 2003 dataset using 38 Argument prompts in the male-female comparison with sample sizes for the male subgroup ranging from 516 to 694. These test takers were separated into 30 categories (as indicated in the table above) as well as into only 10 categories. The resulting PS EZ values were then compared. In this way, we hoped to better understand how comparisons in which the sample size was too small to allow dividing the matching variable into more than 10 categories would affect the results.

The PS EZ difference between the average 10-category (-.059) and the average 30-category conditions (-.026) was -.033. The maximum observed difference was -.061. The absolute size of this difference is small, but so are the differences in this comparison in general. This is especially true for the White versus racial/ethnic minorities comparison, where sample and category sizes were small and overall score differences between groups were somewhat larger. Because impact and DIF are confounded to some extent, one can expect fat matching to increase impact values for these comparisons as well.

Impact of using a quasi-internal matching variable. To evaluate the effect of not having a purely internal matching variable, we created and evaluated an additional matching variable that was less construct-relevant than the two matching variables that were routinely used here. Instead of the z-score of the Verbal section, the z-score of the Quantitative section of the GRE was combined with the weighted z-score of the Issue prompt (for the analysis of Argument prompts using matching variable 2 in the October 2002 to April 2003 dataset).

The objective was to see if DIF values would increase (leading to more false positives) or decrease (thus providing less power for detecting DIF when present) for the gender and racial/ethnic group comparisons under this condition.

Comparing the DIF values obtained by using this matching variable with those obtained by matching variables employed for this study (which, though not perfect, are clearly closer to representing the measured construct), we found that the PS EZ values were higher in all of the prompts in this study that had this specially constructed matching variable. This suggests that the effect of using a less construct-relevant, non-essay test score as part of the matching variable for studying the essay prompts will result in higher DIF values (and therefore likely more false positive values) as compared to a more appropriate combination of non-essay verbal test score and essay prompt and, by inference, as compared to a hypothetical ideal matching variable consisting of various writing measures.

Results

Male-Female Comparison for Argument Prompts

Table 3 shows the summary of the DIF analysis of the Argument prompts between males and females for the 10 prompts that exhibited the highest overall mean rank based on ranked DIF values from the three DIF procedures. Apart from the rankings in all three procedures and the overall ranking, the mean PS EZ value is displayed for each of the prompts. A negative PS EZ value means that the focal group (in this case women) is favored. Three things can be observed from Table 3: (a) The 10 highest PS EZ values are all relatively small, (b) all values indicate that prompts favor women, and (c) the first two prompts on the list display the greatest agreement among the three methods.

The expected essay score curves for the same matching variable scores for the highest-ranking prompt are presented graphically in Figure 1 using the values from the logistic regression (LR) procedure. It should be mentioned here that the matching variable used to create the LR- and polySTAND-based graphs here and later was matching variable 1 $[(z_{\text{verbal}} + z_{\text{other prompt}}) / 2]$. As shown in Figure 1, there were nonuniform effects for this prompt for the LR procedure. The female group is predicted to score higher than the male group for most of the matching score scale, but especially at the higher score levels on the prompt. The vertical distance between the two curves was the largest (0.29) at a matching variable score point of

about 2, while the two lines crossed at a matching variable score point of around -3. When the differences between the two curves were averaged across the actual score range of about -2.8 to 2.4 in the data, the average difference was about 0.18 (favoring females).

Table 3

Summary of DIF Results for the 10 Argument Prompts With the Highest DIF Values for the Male-Female Comparison

| Prompt # | Total N | Males | Overall rank | Rank in Mantel | Rank in LR | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|-------|--------------|----------------|------------|---------------|-----------|-------------|
| A107 | 3,363 | 1,313 | 1 | 1 | 1.5 | 2 | 1.5 | -0.206 |
| A18 | 3,886 | 1,573 | 2 | 2 | 3 | 3 | 2.7 | -0.177 |
| A109 | 4,189 | 1,604 | 3 | 4.5 | 6.5 | 5.5 | 5.5 | -0.155 |
| A98 | 2,715 | 1,055 | 4 | 5 | 6.5 | 6 | 5.8 | -0.153 |
| A65 | 2,989 | 1,185 | 5 | 8 | 4.5 | 6 | 6.2 | -0.151 |
| A20 | 4,691 | 1,874 | 6 | 6 | 7 | 9.5 | 7.5 | -0.142 |
| A12 | 2,722 | 1,026 | 7 | 6 | 12 | 6 | 8 | -0.146 |
| A74 | 3,938 | 1,600 | 8 | 7 | 13 | 14.5 | 11.5 | -0.124 |
| A72 | 2,546 | 1,052 | 9 | 11 | 15.5 | 14 | 13.5 | -0.126 |
| A36 | 3,544 | 1,315 | 10 | 9 | 19 | 14 | 14 | -0.127 |

Note. Prompts are ordered according to their overall mean rank in all three procedures in the left part of the table. The right part of the table displays the PS EZ values for each of the prompts.

In contrast to the logistic regression-based expected essay score curves shown in Figure 1, Figure 2 shows connected points for observed essay score values for males and females at every level of the matching variable (in this case, 50 levels). As shown in Figure 2, the female group scored higher at most of the matching score levels, and such a pattern was more clearly observed at the higher score levels. The conditioned observed mean essay score difference between the male and female groups was the largest (about 0.78) at a matching score level of 44. In sum, both the logistic regression and standardization methods showed a similar DIF pattern for this particular prompt.

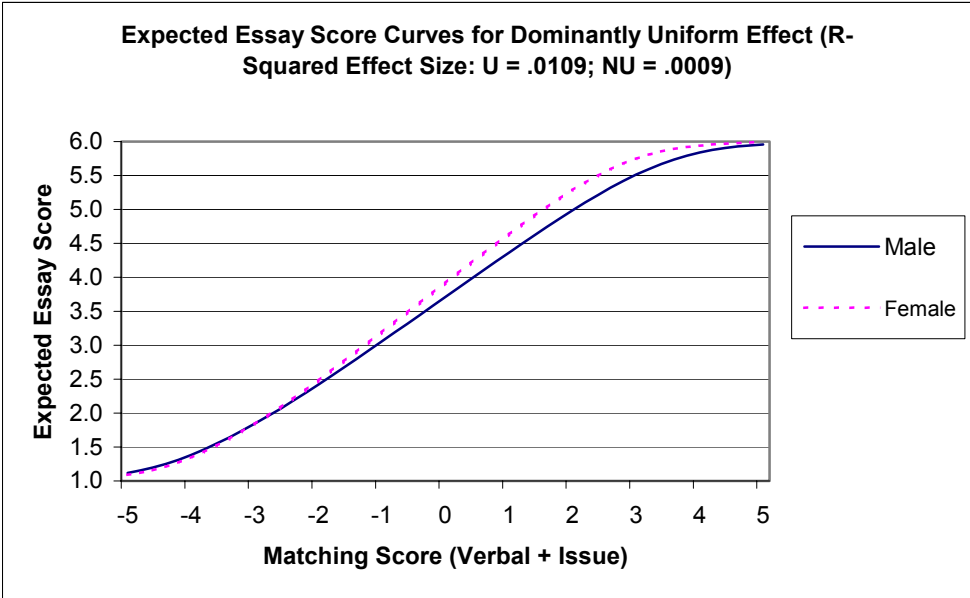


Figure 1. Logistic regression of expected essay score per matching variable 1 score for an argument prompt (A107).

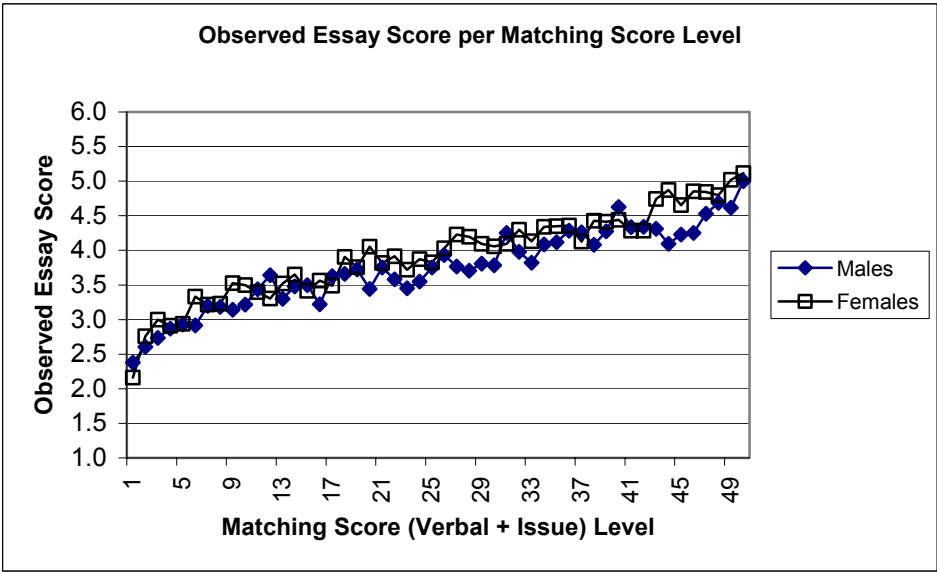


Figure 2. PolySTAND of observed essay score per matching variable 1 score for an argument prompt (A107).

Prompt content analysis. The first step in the content review was to examine the matrix of prompt characteristics. Table 4 gives an example of the five highest-ranking prompts and a small selection of characteristics in the matrix. A close content inspection of these prompts revealed that all of these five prompts require test takers to analyze the logical fallacies related to “alternative explanations and solutions” and “causations versus correlation/reasonableness,” while four of the five prompts dealt with “pragmatic topics.” However, these are very common features, and a correlation of PS EZ values with the presence of these features revealed only very small values (around .15). In contrast, the topic “health/safety,” which is present in the four highest-ranking prompts, is not as common, and the correlation is .29 — fairly high in comparison to other correlations found in this investigation. However, DIF values were too small to signal the need for concern about prompts with this characteristic.

Table 4

Some Characteristics of the Five Highest-Ranking Prompts for the Male-Female Comparison of the Argument Prompts

| Prompt # | HS | SSHC | PT | AES | Ambiguity | CCR | TUE |
|--|------|------|------|------|-----------|------|------|
| A107 | 1 | 0 | 1 | 1 | 1 | 1 | 0 |
| A18 | 1 | 1 | 1 | 1 | 0 | 1 | 1 |
| A109 | 1 | 0 | 1 | 1 | 0 | 1 | 0 |
| A98 | 1 | 0 | 0 | 1 | 0 | 1 | 0 |
| A65 | 0 | 1 | 1 | 1 | 0 | 1 | 0 |
| Proportion of prompts that share this characteristic | .218 | .136 | .436 | .464 | .018 | .482 | .218 |

Note. HS = Health and Safety; SSHC= Social Science, History, and Culture; PT = Pragmatic Topics; AES= Alternative Explanations and Solutions; CCR = Causation versus Correlation/Reasonableness; TUE = Typicality/Unrepresentative Evidence.

Male-Female Comparison for the Issue Prompts

Table 5 summarizes the DIF analysis of the Issue prompts between males and females for the 10 prompts that exhibited the highest overall mean rank based on ranked DIF values from the three DIF procedures. Besides the rankings for all three procedures and the overall ranking, the PS EZ value is displayed in the table. It should also be mentioned that when overall mean ranks were the same for two or more prompts (e.g., I10, I97), the absolute values of PS EZ were used as an additional criterion to rank order the prompts. Most Issue prompts showed low DIF values, all favoring the females. The only prompt that stands out, and where the three DIF detection techniques strongly concur, is prompt number I35.

Table 5

Summary of DIF Results for the 10 Issue Prompts With the Highest DIF Values for the Male-Female Comparison

| Prompt # | Total N | Males | Overall rank | Rank in Mantel | Rank in LR | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|-------|--------------|----------------|------------|---------------|-----------|-------------|
| I35 | 7,120 | 2,533 | 1 | 3 | 1 | 1 | 1.7 | -0.242 |
| I10 | 5,754 | 2,008 | 2.5 | 4 | 2 | 7 | 4.3 | -0.189 |
| I97 | 3,059 | 1,161 | 2.5 | 6 | 3.5 | 3.5 | 4.3 | -0.164 |
| I92 | 6,257 | 2,316 | 4 | 7 | 3.5 | 6.5 | 5.7 | -0.162 |
| I12 | 2,212 | 814 | 5 | 8 | 7 | 4.5 | 6 | -0.156 |
| I22 | 4,453 | 1,540 | 6 | 12 | 6.5 | 2 | 6.8 | -0.148 |
| I87 | 5,988 | 2,185 | 7 | 11 | 6 | 9 | 8.7 | -0.148 |
| I51 | 2,697 | 970 | 8 | 8 | 16 | 11.5 | 11.8 | -0.152 |
| I49 | 4,866 | 1,889 | 9 | 8.5 | 6.5 | 12.5 | 12.5 | -0.131 |
| I69 | 2,644 | 936 | 10 | 12 | 12.5 | 14.5 | 13 | -0.147 |

Note. Prompts are ordered according to their overall mean rank in all three procedures in the left part of the table. The right part of the table displays the PS EZ values for each of the prompts.

Figure 3 shows expected essay score curves for the male and female groups obtained from the logistic regression-based procedure for prompt I35. Again, nonuniform DIF is present in this prompt. In contrast to the Argument prompt (A107) discussed previously, the group differences were larger at the lower matching score levels. As shown in Figure 3, the female group is predicted to score higher than the male group on this prompt, especially at the lower score levels. The two lines crossed at a matching variable score point of 1.6. When the differences between the two curves were averaged across the actual score range of about -2.6 to 2.4 in the data, the average distance was about 0.21 (favoring females).

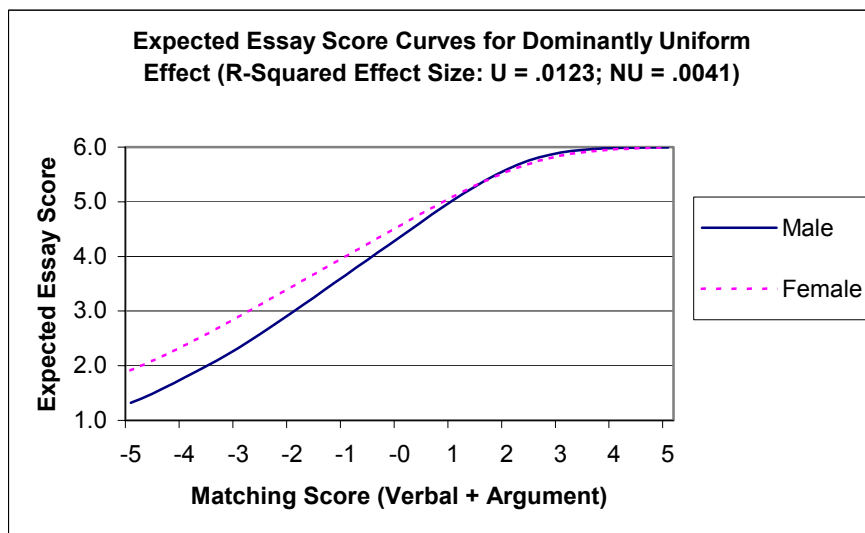


Figure 3. Logistic regression of expected essay score per matching variable 1 score for an issue prompt (I35).

Figure 4 shows connected points for observed score values for males and females at every level of the matching variable. As shown in Figure 4, the female group scored higher on most of the matching score levels, and such a pattern was more clearly observed at the lower score levels. The conditioned observed mean essay score difference between the male and female groups was the largest (about 0.60) at matching score level 4. In sum, both the logistic regression-based and the standardization approach-based methods again show a similar DIF pattern for this particular prompt.

With respect to the prompt content analysis, a correlation of .26 was found between the education/schools topic and the PS EZ values, indicating that prompts involving this topic had higher DIF values (benefiting women) than did prompts that involved other topics.

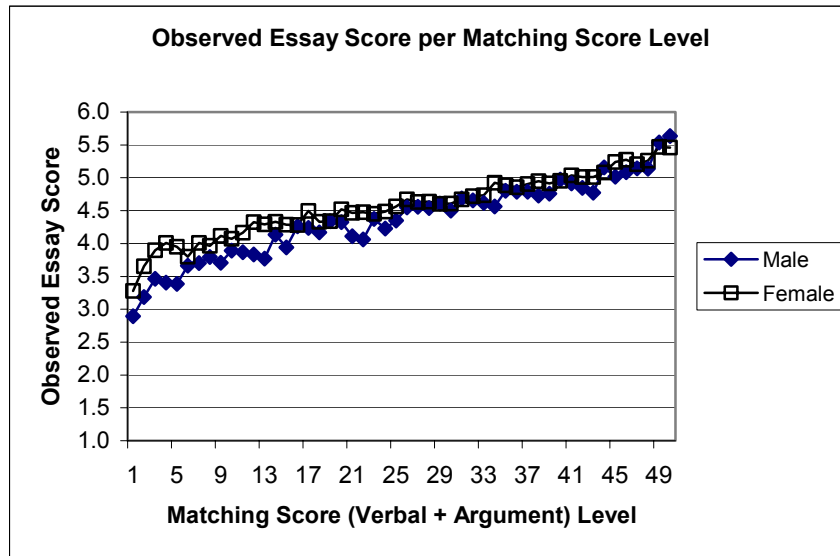


Figure 4. PolySTAND of observed essay score per matching variable 1 score for an issue prompt (I35).

White Versus African American Comparison for the Argument Prompts

Table 6 shows the summary of the DIF analysis of the Issue prompts between the White and African American groups for the 10 prompts that exhibited the highest overall mean rank based on ranked DIF values from two of the DIF procedures (i.e., Mantel, polySTAND). In addition to the rankings for the two procedures and the overall ranking, the PS EZ value is displayed. It should be recalled that the logistic regression procedure was not used for all of the racial/ethnic group comparisons due to the small sample sizes for the focal groups. For this reason, the overall mean ranks for all of the racial/ethnic group comparisons are based on only the Mantel-Haenszel and polySTAND procedures.

Table 6***Summary of DIF Results for the 10 Argument Prompts With the Highest DIF Values for the White Versus African American Comparison***

| Prompt | Total N | African American | Overall rank | Rank in Mantel | Rank in PS EZ | Mean rank | PS EZ value |
|--------|---------|------------------|--------------|----------------|---------------|-----------|-------------|
| A19 | 3,867 | 340 | 1.5 | 1 | 6 | 3.5 | 0.413 |
| A1 | 2,492 | 232 | 1.5 | 4 | 3 | 3.5 | 0.382 |
| A79 | 3,509 | 306 | 3 | 9 | 3.5 | 6.3 | 0.381 |
| A2 | 3,281 | 293 | 4.5 | 5 | 10 | 7.5 | 0.394 |
| A70 | 2,761 | 224 | 4.5 | 9.5 | 5.5 | 7.5 | 0.379 |
| A85 | 2,247 | 219 | 6 | 5.5 | 13.5 | 9.5 | 0.342 |
| A10 | 2,263 | 222 | 7 | 18.5 | 1.5 | 10 | 0.433 |
| A71 | 3,227 | 283 | 8 | 6.5 | 16 | 11.3 | 0.317 |
| A6 | 3,794 | 333 | 9 | 6.5 | 17.5 | 12 | 0.353 |
| A14 | 2,240 | 209 | 10 | 10 | 17 | 13.5 | 0.309 |

Note. Prompts are ordered according to their overall mean rank.

As shown in Table 6, high DIF value comparisons were found here for all of the 10 prompts (all favoring the White group), with all 10 prompts exceeding .30 on the PS EZ measure, and 2 of them exceeding .40. In a separate comparison, we found that the highest impact values (adjusted by the respective standard deviation) were also observed for this particular paired racial/ethnic group comparison. The mean standardized impact was .88 in all studied prompts. The correlation between PS EZ and impact values was .58. This number is similar to correlations found among White/African American impact and DIF measures from six SAT[®] I test forms administered from October 2002 to June 2003 (N. Dorans, personal communication, January 23, 2004).

Figure 5 displays PS EZ and impact values along with the matching score differences between the two groups for all prompts used in this comparison. It should be mentioned that in Figure 5 these prompts were ordered from high to low in terms of their PS EZ values. It is apparent from this graph that there are no outliers, that is, prompts that have much higher PS EZ values than other adjacent prompts. Rather, we observe a very gradual decline from high DIF

prompts to low DIF prompts. Because of the absence of outliers (and because no cut-off value has been established for this kind of DIF study, where no internal matching criterion is available), we cannot say with any certainty how problematic some of these relatively high DIF values may be. We therefore suggest a further study of prompt characteristics that could be associated with higher DIF values. Only a preliminary review could be done within the framework of this investigation, and almost no correlation was found between the DIF values and the pre-established prompt characteristics. The only exception was the logical fallacy category “overgeneralization/insufficient evidence,” which had a correlation of .20 with PS EZ values. Argument prompts that contained this logical flaw were relatively more difficult for African American examinees.

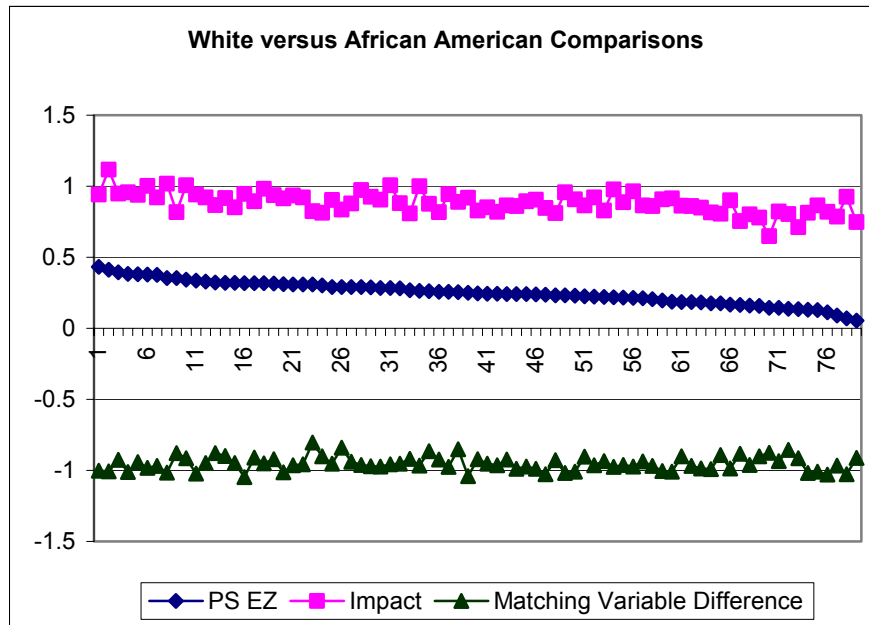


Figure 5. PolySTAND and impact values adjusted by their respective standard deviation for all 79 prompts in the White versus African American comparison (in descending order of polySTAND EZ values).

Note. For better display, matching variable difference, adjusted by their respective standard deviation, are presented on a negative scale.

However, there might be other features that have not been captured by the present classification scheme that could be present in the prompts showing relatively high DIF values. Given the results of this investigation, GRE test development experts may want to conduct further content analyses of the prompt texts and develop alternative classifications.

It is more likely that less tangible characteristics are responsible for higher DIF in certain prompts, and that these characteristics will defy easy categorization. Or possibly, certain combinations of characteristics may play a role. One relatively easy and worthwhile avenue to follow would be a categorization of sentence complexity and the number of different points made in the text (different from text length), which increases processing load and perhaps affects prompt difficulty on a differential basis for different examinee subgroups.

The identification of groups of DIF prompts via content analysis was mixed. Of four reviewers, two identified both groups correctly, while two identified only one group correctly (i.e., no better than chance). The two who identified both groups correctly were both reviewers from racial/ethnic minorities (African American and Hispanic), while the other two reviewers were White.

Reviewers noted that the Argument prompts with the highest DIF values were ones that contained a greater number of statements. Moreover, these statements were relatively complex. (As one reviewer noted, “One can get lost in the topic.”) In contrast, the non-DIF groups of prompts were said to have a very light reading load. For example, about one particular prompt, a reviewer noted that “one point leads to the next, they build on each other.”

White Versus African American Comparison for the Issue Prompts

In contrast to their Argument counterparts, the Issue prompts shown in Table 7 all exhibited much smaller PS EZ DIF values for the White versus African American comparison, with none exceeding .30 and only six exceeding .25, even though all of these prompts favored the White group. Therefore, no further review was undertaken for prompts in this comparison.

Table 7***Summary of DIF Results for the 10 Issue Prompts With the Highest DIF Values for the White Versus African American Comparison***

| Prompt # | Total N | African American | Overall rank | Rank in Mantel | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|------------------|--------------|----------------|---------------|-----------|-------------|
| I93 | 3,115 | 273 | 1 | 1 | 5 | 3 | 0.277 |
| I60 | 3,838 | 359 | 2.5 | 2 | 4.5 | 3.3 | 0.295 |
| I43 | 2,368 | 270 | 2.5 | 5.5 | 1 | 3.3 | 0.291 |
| I34 | 2,139 | 242 | 4 | 4.5 | 6 | 5.3 | 0.262 |
| I15 | 1,831 | 238 | 5.5 | 5.5 | 8 | 6.8 | 0.239 |
| I94 | 2,957 | 324 | 5.5 | 11 | 2.5 | 6.8 | 0.291 |
| I13 | 4,140 | 449 | 7 | 13 | 5.5 | 9.3 | 0.265 |
| I26 | 2,962 | 284 | 8 | 3.5 | 17.5 | 10.5 | 0.203 |
| I22 | 3,328 | 407 | 9 | 11.5 | 11.5 | 11.5 | 0.227 |
| I71 | 4,317 | 386 | 10 | 13.5 | 13 | 13.3 | 0.211 |

Note. Prompts are ordered according to their overall mean rank.

White Versus Hispanic Comparison for the Argument Prompts

The 10 highest PS EZ values in this comparison, shown in Table 8, are smaller than those in the White versus African American group comparison for the Argument prompt type. The average impact of the 67 studied prompts was also lower (.47) than that of those in the White versus African American group comparison, favoring the focal (White) group. The extent of DIF found was comparable to that found in the Issue prompt comparison in the White versus African American comparison. Nevertheless, the decision was made to submit low and high DIF prompts for expert review to gather information about characteristics that could negatively affect this group. Also, the Argument prompts seemed to lend themselves better to in-depth content analysis than did the Issue prompts because they provide more concrete contexts for analysis and thinking.

Table 8***Summary of DIF Results for the 10 Argument Prompts With the Highest DIF Values for the White Versus Hispanic Comparison***

| Prompt # | Total N | Hispanic American | Overall rank | Rank in Mantel | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|-------------------|--------------|----------------|---------------|-----------|-------------|
| A85 | 2,247 | 151 | 1 | 5 | 1.5 | 3.3 | 0.308 |
| A108 | 4,419 | 231 | 2 | 2.5 | 4.5 | 3.5 | 0.264 |
| A83 | 2,936 | 188 | 3 | 4.5 | 4.5 | 4.5 | 0.281 |
| A6 | 3,794 | 195 | 4 | 2 | 8 | 5 | 0.265 |
| A106 | 2,874 | 176 | 5 | 6 | 5.5 | 5.8 | 0.241 |
| A68 | 4,516 | 265 | 6.5 | 4 | 10 | 7 | 0.239 |
| A78 | 2,877 | 172 | 6.5 | 5 | 9 | 7 | 0.234 |
| A38 | 3,245 | 178 | 8 | 10 | 6 | 8 | 0.235 |
| A31 | 3,320 | 188 | 9 | 11.5 | 6.5 | 9 | 0.268 |
| A76 | 2,244 | 160 | 10 | 13 | 7.5 | 10.3 | 0.231 |

Note. Prompts are ordered according to their overall mean rank.

The results of the reviewers' analysis were similar to the ones described before: Two reviewers identified the groups correctly, while two other reviewers identified only one of the DIF groups. Of the two who identified both groups correctly, one was White and one was a member of a racial/ethnic minority group (African American). Of the other two reviewers, one was also White and the other was a member of a racial/ethnic minority group (Hispanic American). No substantial correlations were found between DIF values and other predefined prompt characteristics. While the five prompts with the highest PS EZ values were all in the category "causation versus correlation," the correlation between the PS EZ values and the classification codes for the "causation versus correlation/reasonableness" category over all prompts was only .12.

The main characteristic that was identified as being common to high-DIF prompts was a high quantitative content, which was hypothesized to confuse lower ability examinees from the Hispanic American group: "The numbers in the prompt texts make you think you have to do

something with them,” one reviewer noted. Moreover, when there was quantitative content in the non-DIF prompts, the task seemed simpler.

Again, a higher reading load was mentioned for the prompts exhibiting higher DIF. In contrast, the prompts with low DIF values dealt more with social topics, and it was argued that social commitment levels are high among members of this racial/minority group, which could help offset other effects that might lead to a disadvantage for this group.

White Versus Hispanic Comparison for the Issue Prompts

PS EZ DIF values in this comparison, shown in Table 9, were generally lower than those obtained for the Argument prompts. With the highest PS EZ values around .20, no further investigation was undertaken.

Table 9

Summary of DIF Results for the 10 Issue Prompts With the Highest DIF Values for the White Versus Hispanic Comparison

| Prompt # | Total N | Hispanic American | Overall rank | Rank in Mantel | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|-------------------|--------------|----------------|---------------|-----------|-------------|
| I9 | 2,593 | 167 | 1 | 2 | 2 | 2 | 0.215 |
| I89 | 2,456 | 160 | 2 | 5.5 | 1 | 3.3 | 0.211 |
| I93 | 3,040 | 198 | 3 | 2 | 5.5 | 3.8 | 0.179 |
| I65 | 4,862 | 383 | 4 | 5 | 4 | 4.5 | 0.180 |
| I52 | 2,993 | 207 | 5 | 2 | 8.5 | 5.3 | 0.138 |
| I54 | 2,279 | 156 | 6 | 5 | 6 | 5.5 | 0.152 |
| I43 | 2,292 | 194 | 7 | 7 | 6 | 6.5 | 0.156 |
| I18 | 2,548 | 161 | 8 | 9 | 10.5 | 9.8 | 0.122 |
| I26 | 2,858 | 180 | 9 | 14 | 6.5 | 10.3 | 0.154 |
| I95 | 4,136 | 272 | 10 | 10 | 12.5 | 11.3 | 0.122 |

Note. Prompts are ordered according to their overall mean rank.

White Versus Asian American Comparison for the Argument Prompts

Table 10 shows the summary of the DIF analysis of the Argument prompts between the White and Asian American groups for the 10 prompts that exhibited the highest overall mean rank based on ranked DIF values from the two DIF procedures. The White versus Asian American comparison for the two prompt types is the only one in which the highest DIF values had (in general) different signs depending on the type of prompt. For the Argument prompts, as indicated by negative PS EZ values here, Asian Americans test takers had an advantage when compared to White test takers on 9 out of 10 prompts with the highest absolute PS EZ values. The average of all studied prompts (with both signs), however, is close to 0 (-.04). Therefore, no further review was carried out for the Argument prompts for this comparison.

Table 10

Summary of DIF Results for the 10 Argument Prompts With the Highest DIF Values for the White Versus Asian American Comparison

| Prompt # | Total N | Asian American | Overall rank | Rank Mantel | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|----------------|--------------|-------------|---------------|-----------|-------------|
| A90 | 2,111 | 161 | 1 | 1 | 1.5 | 1.3 | -0.316 |
| A33 | 2,817 | 240 | 2 | 2 | 2.5 | 2.3 | -0.272 |
| A55 | 3,070 | 251 | 3 | 3 | 3.5 | 3.3 | -0.204 |
| A32 | 2,973 | 244 | 4 | 5.5 | 2 | 3.8 | -0.207 |
| A10 | 2,263 | 174 | 5 | 5.5 | 4 | 4.8 | -0.202 |
| A38 | 3,245 | 264 | 6 | 6.5 | 4.5 | 5.5 | -0.150 |
| A42 | 2,463 | 194 | 7 | 7 | 5.5 | 6.3 | -0.185 |
| A21 | 3,243 | 253 | 8 | 8 | 5 | 6.5 | -0.126 |
| A80 | 2,320 | 179 | 9 | 11 | 6.5 | 8.8 | -0.121 |
| A93 | 4,967 | 399 | 10 | 11.5 | 7 | 9.3 | 0.150 |

Note. Prompts are ordered according to their overall mean rank.

White Versus Asian American Comparison for the Issue Prompts

Table 11 displays the results of the analysis of the Issue prompts between the White and Asian American groups for the 10 prompts that exhibited the highest overall mean rank based on ranked DIF values from the two DIF procedures. For the Issue prompts, all 10 prompts exhibiting high levels of DIF had positive PS EZ values, indicating that Asian American test takers are disadvantaged, which is in contrast to the results for the Argument prompts. While only two prompts had DIF values exceeding .30, the effect was very consistent over all prompts, with an average of .19 for all studied prompts.

Table 11

Summary of DIF Results for the 10 Issue Prompts With the Highest DIF Values for the White Versus Asian American Comparison

| Prompt # | Total N | Asian American | Overall rank | Rank in Mantel | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|----------------|--------------|----------------|---------------|-----------|-------------|
| I52 | 3,055 | 269 | 1 | 1.5 | 1.5 | 1.5 | 0.312 |
| I22 | 3,114 | 193 | 2 | 2 | 2 | 2 | 0.313 |
| I89 | 2,502 | 206 | 3 | 3 | 4 | 3.5 | 0.269 |
| I14 | 3,692 | 339 | 4 | 3.5 | 4.5 | 4 | 0.287 |
| I7 | 2,265 | 169 | 5 | 7.5 | 6.5 | 7 | 0.274 |
| I93 | 3,159 | 317 | 6 | 5 | 9.5 | 7.3 | 0.264 |
| I67 | 2,605 | 219 | 7 | 10 | 7 | 8.5 | 0.286 |
| I13 | 3,980 | 289 | 8 | 9 | 11.5 | 10.3 | 0.261 |
| I78 | 2,243 | 154 | 9 | 11 | 11.5 | 11.3 | 0.249 |
| I92 | 4,337 | 352 | 10 | 9.5 | 15 | 12.3 | 0.255 |

Note. Prompts are ordered according to their overall mean rank.

The reviewers' decisions regarding DIF versus non-DIF prompts in this comparison were no better than chance. The task of identifying characteristics in the Issue prompts seemed to be much more difficult, possibly because most prompts contain only a single statement. A post hoc analysis of the characteristics common to high-DIF prompts suggested that the topics were generally lofty, philosophical, abstract, and tended to use abstract phrases like "destiny of

society” or “the condition of humanity” — concepts that may have different cross-cultural connotations.

EBL Versus Non-EBL Comparison for the Argument and Issue Prompts

The results for comparison of these groups again depended strongly on the type of prompt: While the Argument prompts showed very low DIF (Table 12), the PS EZ values for the Issue prompts (Table 13) were of moderate size and demonstrated a disadvantage for the non-EBL group. They were also consistent with the average PS EZ values over all 100 studied prompts (.15). Due to the heterogeneous structure of this focal group, no review was carried out at this point, although we suggest forming a group that is made of test developers of European, Asian, and Hispanic backgrounds, whose best language is one other than English, to review the Issue prompts in the future.

Table 12

Summary of DIF Results for the 10 Argument Prompts With the Highest DIF Values for the EBL Versus Non-EBL Comparison

| Prompt # | Total N | Asian American | Overall rank | Rank in Mantel | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|----------------|--------------|----------------|---------------|-----------|-------------|
| A88 | 1,926 | 203 | 1 | 6 | 5.5 | 5.8 | -0.112 |
| A65 | 3,340 | 351 | 2 | 10.5 | 3.5 | 7 | 0.114 |
| A102 | 1,902 | 183 | 3 | 10 | 10 | 10 | 0.120 |
| A29 | 3,690 | 350 | 4 | 18 | 6.5 | 12.3 | 0.108 |
| A21 | 4,393 | 471 | 5 | 18 | 8.5 | 13.3 | 0.102 |
| A16 | 2,437 | 235 | 6 | 4.5 | 24 | 14.3 | -0.077 |
| A27 | 3,819 | 404 | 7 | 15 | 16 | 15.5 | 0.094 |
| A11 | 3,274 | 357 | 8 | 18.5 | 14 | 16.3 | 0.102 |
| A23 | 1,820 | 191 | 9 | 13 | 23.5 | 18.3 | 0.074 |
| A50 | 2,388 | 246 | 10 | 19.5 | 17.5 | 18.5 | 0.087 |

Note. Prompts are ordered according to their overall mean rank.

Table 13***Summary of DIF Results for the 10 Issue Prompts With the Highest DIF Values for the EBL Versus Non-EBL Comparison***

| Prompt # | Total N | Non-EBL | Overall rank | Rank in Mantel | Rank in PS EZ | Mean rank | PS EZ value |
|----------|---------|---------|--------------|----------------|---------------|-----------|-------------|
| I56 | 2,075 | 252 | 1 | 2.5 | 1 | 1.8 | 0.300 |
| I18 | 4,089 | 422 | 2 | 1.5 | 2.5 | 2 | 0.269 |
| I87 | 6,623 | 635 | 3 | 2 | 2.5 | 2.3 | 0.282 |
| I92 | 6,981 | 724 | 4 | 4.5 | 7.5 | 6 | 0.251 |
| I22 | 4,932 | 479 | 5 | 7.5 | 8 | 7.8 | 0.246 |
| I103 | 2,610 | 259 | 6 | 7.5 | 8 | 7.8 | 0.237 |
| I35 | 7,875 | 755 | 7 | 4.5 | 11.5 | 8 | 0.240 |
| I50 | 5,437 | 569 | 8 | 11.5 | 7 | 9.3 | 0.240 |
| I21 | 1,648 | 161 | 9 | 15 | 4.5 | 9.8 | 0.245 |
| I60 | 6,228 | 687 | 10 | 7.5 | 14.5 | 11 | 0.221 |

Note. Prompts are ordered according to their overall mean rank.

Summary and Discussion

This study was designed to investigate the comparability of both GRE Argument and GRE Issue prompts for different test-taker subgroups, using several alternative DIF procedures. For these DIF analyses, test takers were matched on a combination of their scores on the Verbal section of the GRE General Test and their scores on the other prompt (Issue or Argument). Our first finding was that the DIF methods investigated here (polySTAND, the Mantel test, and logistic regression) were in substantial, though not perfect, agreement with one another in terms of the extent to which they identified prompts having large DIF values. Thus, future screenings might be accomplished in a more cost-effective way by using only one of these procedures. If this were to become a realistic option, we recommend using the polySTAND procedure, whose metric may be more intuitively understood and whose values are, in contrast to the Mantel test, uninfluenced by differences in sample sizes. Because it requires larger sample sizes, the logistic regression procedure could only be employed in this study for the gender comparison. However,

this procedure may become a more feasible DIF research tool for the GRE after further data collection, since it is not only capable of answering questions about nonuniform DIF (across different score levels), but can also be used to study possible interaction between groups (e.g., between racial/ethnic groups and gender groups or between gender groups and groups determined by the examinees' major field of study).

We also found, as expected, that the matching variables that were available for studying DIF in essay prompts were not entirely adequate, since they do not constitute a truly internal matching criterion. Moreover, as one of the reviewers of this report noted, because GRE Verbal and Writing scores measure different constructs, they cannot be expected to order subgroups of test takers in the same way. This reviewer thus questioned whether it made sense psychometrically to use GRE Verbal scores as part of the matching criterion, and he was concerned that using this criterion might result in the flagging of some prompts simply because they were less highly related to verbal ability. This point is indisputable, of course, and suggests the need for an extremely thorough review of any flagged prompts, and the development of compelling rationales, before any action is taken.

To prove that a less construct-relevant matching variable is likely to result in inaccurately high DIF values rather than in decreased detection power, we decreased the purity of the matching variable further by substituting the GRE quantitative ability score for the GRE verbal ability score. This substitution indeed resulted in considerably higher DIF values, again reinforcing the need for caution when making any decisions about the possible lack of fairness of flagged prompts.

Of the two matching variables used in the analysis (both of which entailed a composite score of the GRE Verbal section plus the other prompt), the matching variable in which the other prompt score received more weight than in the Verbal matching variable, resulted in lower DIF values. This is also consistent with the findings from Welch and Miller's (1995) study on writing prompt DIF.

In order to better understand how another factor might influence DIF values, for some gender comparisons we increased the numbers within each score category (by decreasing the number of levels of the matching variables from 30 to 10), and evaluated the effect of this modification. Our interest in the cruder, 10-level matching related to the fact that many prompts in the racial/ethnic group comparisons relied on relatively small sample sizes, which permitted us

to divide the matching variables into only 10 levels. The effect of fat matching (i.e., larger samples within score categories) was higher DIF values for the gender comparisons, indicating that DIF values would likely have been lower for the ethnic group comparisons if larger sample sizes had been available for “thinner matching.”

Thus, under optimal conditions for DIF analysis (that is, with a pure matching criterion and larger sample sizes), DIF values would most likely have been lower than those computed in this study, although it is not known by how much. Because our matching criterion was not entirely adequate, a possibility for eventual follow-up research would entail developing a better matching variable (for example, by asking GRE test takers to write multiple essays on various GRE writing prompts, thus providing a truer internal matching criterion). The DIF values found in the presence of such a matching criterion could be compared to those obtained using the two matching variables used here. The difference of DIF values in this study and the proposed follow-up could provide insight as to how much the DIF values found in our more economical approach were inflated as a result of our using a less construct-related matching criterion. In the event that the differences are consistent over several prompts and groups, and taking into consideration other factors like impact and sample size, one could use this information to define new flagging values for prompts that lack a purely internal matching criterion.

Because of the reasons detailed above, we decided against flagging prompts based on specific cut-off values. Instead, we ranked all prompts based on averaged ranks over different DIF methods and conducted a preliminary analysis of prompt characteristics that could be associated with higher DIF values.

The largest DIF values to the disadvantage of the focal group were found for the African American group for the Argument prompts. Moderate values were observed for the African American, Asian American, and the non-EBL groups for the Issue prompts and for the Hispanic group for the Argument prompts. To the advantage of the focal group, moderate DIF values were found for the Asian American group for the Argument prompts, and low (although consistent) values were observed in the gender comparison (favoring the female group) for both prompt types.

Even though prompts showed a range of DIF values, no outliers (i.e., prompts that exhibited substantially higher DIF values than the others) were found. Our analyses did not, therefore, suggest the need to delete any prompts from the currently used pools of GRE

Argument and Issue prompts. However, we suggest that further investigation may be necessary to ensure that some *combinations* of Issue and Argument prompts (i.e., high-ranked Argument prompts paired with high-ranked Issue prompts) are avoided, especially for African American test takers.

Finally, this investigation took preliminary steps to link DIF values to prompt characteristics, but further investigation in this area is clearly indicated. Pre-established prompt features showed only low to moderate correlations with DIF values. For example, with respect to gender comparisons for the Argument prompts, DIF values correlated .29 with whether or not topics dealt with health and safety issues (health/safety topics being even more easy for women than other topics). Most characteristics, however, showed no relationship at all with DIF values. However, there might be other features that have not been captured by the existing classification scheme that could account for relatively high DIF for certain groups for the prompts we studied. Based on the DIF results provided in this investigation, GRE test development experts might wish to develop alternative classifications of prompts that can be used to further analyze prompt texts.

It is more likely that less tangible, harder-to-classify characteristics are responsible for higher DIF in certain prompts. As suggested by our expert reviewers, one relatively easy and possibly worthwhile avenue to explore would be a categorization of sentence complexity and an analysis of the number of different points made in the text (in contrast to overall text length), which increases processing load. These points were mentioned as one possible factor that might explain the relatively higher DIF values in the Argument prompt for the African American and Hispanic groups. For the latter group, the high quantitative content of some prompts was hypothesized to be another characteristic that might contribute to higher DIF values.

Finally, we suggest that the kind of review we have begun here might be broadened, perhaps by including more reviewers with ethnic minority backgrounds. These additional reviewers and a broader classification scheme might in turn provide additional insights that could be incorporated in future reviews of new prompts before they are employed operationally. We also note that the content of all GRE prompts is routinely reviewed for the potential introduction of bias against a particular group. Notes from this content or sensitivity review could also be incorporated in this review process in the future.

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Notes

- ¹ It was referred to in the proposal as the “Generalized Mantel-Haenszel” procedure, but a reviewer noted that this was an imprecise term that has been wrongly repeated in some literature.
- ² One can argue that IRT-based methodology can also be used by obtaining theta estimates from a calibration of only two essay prompts (i.e., the studied prompt plus the other prompt) assigned to each examinee in the writing section, and using the estimate as an internal matching variable. It should be noted, however, that the examinees do not take the same pair of Issue and Argument prompts on the computer-based GRE. Rather, the testing software selects Argument prompts from a pool of prompts and assigns one prompt to each examinee. The examinees also have the option of choosing one prompt out of the two Issue prompts from a pair of the prompts assigned to each examinee. For this reason, using an IRT-based methodology for prompt DIF investigation is not straightforward.
- ³ One can say that this weighting scheme seems rather arbitrary, and a more elaborate methodology (e.g., IRT item information statistics) could also have been used to decide the optimal worth of a single polytomously scored essay item, compared to a single dichotomously scored multiple-choice item. It should be pointed out, however, that the main reason for using the second matching variable was to examine the impact on DIF of decreasing the weight of the more construct-relevant (essay) component to the matching variable, whatever the weighting scheme is. In that sense, the psychometric stringency in deciding an optimal weight for the essay item is not an important issue in this study.
- ⁴ No smoothing was carried out in this investigation.
- ⁵ One can argue that rank-ordering prompts based on DIF values in each method does not take into account actual variability among prompts in terms of DIF values. Thus, while the ranks for the DIF values for all the prompts analyzed will be very much clustered, the DIF values themselves may be disperse. Nevertheless, it should be noted that the main reason for computing the mean ranks of the prompts over different methods is to identify prompts that had the large DIF values consistently across different methods, for content expert review.
- ⁶ It should be noted that the pooled (combined) standard deviation in the case of the polySTAND measure was used to make the resulting PS EZ values comparable over different prompts. The resulting PS EZ values show how much a prompt is “DIFed” in terms of a pooled standard

deviation for each item. In other words, this is not done to control for different sample sizes (which does not affect the polySTAND procedure) but for different variability of scores in different prompts. In contrast, the adjustment in the Mantel test is done to control for different sample sizes of the prompts that do have an effect on the chi-square statistic. We used the sample sizes for the smaller (focal) group rather than the total group to make such an adjustment for the following two reasons: (a) This would be consistent with the categorization scheme for matching variables, which is based on the smaller group, not the total group, and (b) this scheme seemed to produce better results (i.e. the ranks of polySTAND procedure and the Mantel test seemed to be better matched).

⁷ The data were categorized based on percentile groups, with each group containing approximately the same number of cases. The procedure “categorize variables” in SPSS (SPSS, 1999) was used to do it. For this reason, the score range for each category depends on the prompt, not on the way it was categorized.

Appendix

Content Feature Coding Scheme for GRE Writing Prompts

For each prompt, the presence of a certain feature is represented by a 1 in the matrix, and its absence by a 0. The following elements are part of that matrix:

Content (both prompts)

Business/Economics
Education/Schools
Employment
Environment/Natural Science
Ethics
Fine Arts
Health/Safety
Humanities/Lang/Lit
Interdisciplinary
Knowledge/Truths/Learning
Politics/Government.
Popular Culture/Media
Psychology
SS/History/Cultures
Technology

Focus (both prompts)

Conceptual
Pragmatic
Both

Argument-Reasoning

Alternative explanations and solutions
Ambiguity (vagueness, shift in meaning)
Assumptions/complex question/begging the question
Causation vs. correlation/reasonableness
Circularity/tautology
Conclusion/inference invalidly drawn
Efficacy of proposal
False dichotomy/either-or
False equivalence/illogical comparisons
Internal contradictory/incompatible assert
Non sequitur
Overgeneralization/insufficient evidence
Strawman
Typicality/unrepresentative evidence

Issue-Reasoning

Description/Definition
Cause/Effect
Comparison
Problem/Recommendation
Problem/Recommendation +
Solution/Rationale



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