This study was designed to explain how Profile Analysis via Multidimensional Scaling (PAMS) could be viewed as a structural equations model (SEM). The study replicated the major profiles extracted from PAMS in the context of the latent variables in SEM. Data involved the Basic Theme Scales of the Strong Campbell Interest Inventory (Campbell and Hanse, 1985). Data were collected from 1,308 males, who were clients of a vocational assessment clinic. Findings show that the profile patterns identified by PAMS can be used to examine the relationships between the profile patterns and other criterion variables, but the results from a PAMS approach can only be used as an exploratory tool. The profile patterns can be examined in confirmatory factor analysis as shown by this study. Also, when major profiles identified in PAMS are replicated in SEM, one can examine whether significant test results of the coordinates in PAMS match those of factor loadings in SEM, since SEM analyses always provide asymptotic standard errors corresponding to estimates of parameters. (Contains 2 figures, 3 tables, and 14 references.) (SLD)
Connecting SEM Analysis and Profile Analysis via MDS

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This paper was designed to explain how the Profile Analysis via Multidimensional Scaling (PAMS) model could be viewed as a structural equations model (SEM). In the previous presentation (Davison, Kim, & Ding, 2001), the PAMS model was used in an exploratory MDS analysis for identifying major profiles. This paper presented a parallel analysis of the same data but attempted to develop a confirmatory approach by means of the structural equations model under the assumptions of the PAMS model. One may speculate that there is a connection between latent variables in SEM and dimensions in MDS, and easily conclude that the latent variables in SEM correspond MDS dimensions. MDS dimensions are actually major profiles in PAMS, but PAMS emphasizes interpretation of patterns in dimensions. In the paper, we replicated the major profiles extracted from PAMS in the context of the latent variables in SEM.

A line of research (e.g., Rounds, Davison, & Dawis, 1979; Tracy & Rounds, 1993) by James Rounds suggests that it is possible to use structural equations modeling to develop a confirmatory approach for a MDS analysis. Rounds has been studying the structure of vocational interests. Although Rounds has not exploited the profile pattern interpretation of his results, he began using multidimensional scaling to explore the core dimensional profile patterns that underlie vocational interests (Rounds, Davison, & Dawis, 1979). Using a relationship between factor analysis and multidimensional scaling developed by Davison (1985), Rounds has developed a confirmatory analysis of covariance structures (e.g., Tracy & Rounds, 1993).
The combination of the PAMS model and Rounds' work implies that by considering the linear components of measurement around profile patterns, rather than latent factors, it is possible to connect the analysis of covariance structures to the study of profile patterns, if a researcher intents to study profiles defined over variables that have been standardized in the research sample.

To specify the PAMS model with \( K \) major profile patterns as a structural equations model, one must specify a structural equations model with \( 1+K \) latent variables. The first additional latent variable corresponds to the level parameter in the PAMS model and the last \( K \) latent variables correspond to the \( K \) major profile patterns in PAMS. As introduced in the previous presentation (Davison et al., 2001), the level parameter accounts for each individual's profile height and determines how much the individual's profile elevated or depressed from the major profiles. The level parameter is estimated by the individual's total score over subscales.

Then one may ask why SEM analysis includes one more additional latent variable than PAMS. To answer, it is necessary to examine the relation between parameter estimates in the factor model and the PAMS model. Davison (1985) compared MDS analysis with principal components analysis (PCA), using the same test intercorrelations. In his Monte Carlo studies, Davison replicated \( 1+K \) factors in PCA with \( K \) dimensions in a MDS analysis when the data being analyzed were simulated to include one general factor plus \( K \) group factors, and showed that \( K \) group factors in PCA corresponded \( K \) dimensions in MDS. Again, notice that dimensions in MDS are major profiles in PAMS. The first principal factor in PCA can be interpreted as Spearman's \( g \) factor or general factor in human intelligence tests or as trivial response bias (Hanson, Prediger, & Schussel, 1977), such as individuals' particular response patterns, in interest/attitude assessment scales, and this general factor has usually high but equivalent
loadings across stimuli. When Davison (1985) reported the relationship between PCA and MDS, he did not develop the PAMS model, and could not include psychometric property of the general factor regarding MDS until in 1996. Now we propose that this general factor in the factor model corresponds to the level parameter in the PAMS model.

In consistent with the constraints of the PAMS model, all parameters on these $K$ latent variables are allowed to vary freely with the constraint that the parameters must sum to zero on each latent variable. This constraint has the solution uniquely identified. Further, the last $K$ latent variables are constrained to be uncorrelated. The first latent variable accounts for individual differences in profile level. Along this latent variable, all observed variables are constrained to have equal weights since the first latent variable corresponds to the first principal factor in the factor model (see pp. 95-96, Davison, 1985). The first latent variable need not be uncorrelated to the other $K$ latent variables, but it is left free to be correlated with them. In this paper, exploratory major profile patterns defined by the PAMS model are fitted to the data.

Method

The data used in this study involves the Basic Theme Scales from the Strong-Campbell Interest Inventory (Campbell & Hansen, 1985). The scales consist of the Realistic (REAL), Investigative (INVES), Artistic (ART), Social (SOCIAL), Enterprising (ENT), and Conventional (CONV) interest scales. Holland (1973) proposed a two-dimensional, hexagonal model for the six scales and this two-dimensional solution was interpreted as People vs. Things and Data vs. Ideas by Prediger (1982) and Tracey & Rounds (1993). The data for the study were collected by Rene Dawis and David J. Weiss from 1308 males who were clients of the Vocational
Assessment Clinic at the University of Minnesota. Table 1 shows the intercorrelations of the six interest scales.

For an exploratory PAMS, a nonmetric MDS was used to analyze these correlations and resulted in the two-dimensional solution. Since $\delta_{rr'} = \left(1 - r_{rr'}\right)^{1/2}$ (see p.105, Davison, 1993), where $r$ and $r'$ refer to subscales, dissimilarity between two tests is inversely related with correlation. A confirmatory factor analysis, based on the same intercorrelations used in the exploratory PAMS approach, was performed through LISREL on the six subscales of the Strong-Campbell Interest Inventory. The hypothesized model is presented in Figure 1 where circles represent latent variables, and rectangles represent observed variables (or subscales).

A three-factor model of “Interest,” Factor 1, General factor (which corresponds to the level parameter in the PAMS model and two group factors, Factor 2 and Factor 3, is hypothesized. Factor 2 and Factor 2 correspond the first and the second major profiles in PAMS, respectively. All six observed variables, which are REAL, INVES, ART, SOCIAL, ENT, and CONV, serve as indicators of Factor 1, General factor. INVES, ART, ENT, and CONV serve as indicators of Factor 2. REAL, INVES, SOCIAL, and ENT serve as indicators of Factor 3. According to the assumptions of the PAMS model, last two factors are constrained to be uncorrelated with one another, but the first (general) factor is left free to be correlated with the other two factors.
Results

Model Estimation

Unweighted least-squares estimation was employed to estimate all models. The hypothesized model (see Figure 1) was tested without inclusion of error covariances among observed variables (or subscales) but was not supported for the model, \( \chi^2 (11, N = 1308) = 240.57, P\text{-value} = .00, \text{RMSEA} = .13, \text{AIC} = 264.57, \) and \( \text{GFI} = .96. \)

The model modifications were examined in an attempt to develop a better fitting, and possibly more parsimonious model. On the basis of the modification indices, the error covariances were added between ART and REAL, between SOCIAL and INVES, between ENT and INVES, between ENT and ART, and between CONV and ART. The error correlations for ART and REAL, for SOCIAL and INVES, for ENT and INVES, for ENT and ART, and for CONV and ART were \(-.25 (.03), .16 (.04), -.16 (.04), -.15 (.04), \) and \(-.26 (.04)\), respectively and all of them were significant at \( \alpha = .001. \) The values in parentheses are standard errors of the estimates. The modified model that allowed the error correlations was supported and provided, \( \chi^2 (6, N = 1308) = 7.90, P\text{-value} = 0.25, \text{RMSEA} = 0.02, \text{AIC} = 41.90, \text{CFI} = 1.00, \) and \( \text{GFI} = 1.00. \)

Parameter Estimation

Table 2 shows the two-dimensional solution resulting from a nonmetric MDS analysis of the correlations among Interest Subscales. Dimension 1 was similar to Prediger’s (1982) Data vs. Ideas Dimension. According to theory, the Data end of the dimension should be marked by the Enterprising and Conventional scales that fall at the positive end of Dimension 1. The Ideas end should be marked by the Artistic and Investigative scales. Both fall at the negative end of
Dimension 1, although the Investigative scale does not fall as far toward the negative end as theory would lead one to expect.

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Insert Table 2 about here.

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According to theory, the People end of Dimension 2, the People vs. Things dimension, should be marked by the Social and Enterprising scales that fall at the negative end of Dimension 2. The Things end should be marked by the Realistic and Investigative scales falling at the positive end of Dimension 2.

By inspection of each dimension’s profile pattern, Davison et al. (2001) renamed the Data vs. Things dimension as Enterprising/Conventional (or E/C) Profile in the intent of interpreting a dimensional pattern, since Enterprising and Conventional subscales marked high points in the plot of the Data vs. Things dimension, and also relabeled the People vs. Things dimension as Realistic/Investigative (or R/I) Profile since these two subscales marked high points in the plot of the People vs. Things dimension.

For SEM, three latent factors were assigned, Factor 1, Factor 2, and Factor 3, to the six Interest Subscales. Loadings of all six subscales on Factor 1 that accounts for the heights of individuals’ observed profiles in PAMS were constrained to be equal. LISREL assigned 0.58 to the six observed variables for Factor 1. According to the results of the dimensional solution in PAMS, INVES, ART, ENT, and COV were assigned to Factor 2 that corresponds to Enterprising/Conventional Profile. The second factor loadings on INVES, ART, ENT, and COV were -0.05 (.05), -0.51 (.05), 0.27(.05), and 0.29 (.05), respectively. The values in parentheses are standard errors of the estimates. Except the loading of INVES, the loadings of ART, ENT,
and COV were significant at $\alpha = .001$. REAL, INVES, SOCIAL, and ENT were assigned to Factor 3 that corresponds to Realistic/Investigative Profile. The third factor loadings on REAL, INVES, SOCIAL, and ENT were 0.33 (.03), 0.34 (.04), -0.47 (.04), and -0.20 (.04), and all these loadings were significant at $\alpha = .001$.

According to the assumptions of the PAMS model, Factor 2 and Factor 3 were constrained to be uncorrelated, but Factor 1 (level parameter factor) was left free to be correlated with the other two factors. The correlations for Factor 1 and Factor 2 and for Factor 1 and Factor 3 were 0.17 (.04) and 0.16 (.03), respectively, and both were statistically significant at $\alpha = .001$.

Figure 2 shows the path diagram between six observed variables (Interest Subscales) and three latent factors. Table 3 includes the weight of the three latent factors in SEM and scale-values of two major profiles (E/C & R/I Profiles).

Although the scale-values in PAMS were different from the factor weights in SEM since different parameterization was used for each approach, the direction and magnitude of the scale-values were replicated in SEM. Figure 3 shows similar patterns between Factors 2 and 3 in SEM and Major Profiles 1 and 2 in PAMS.
Profile Analysis via Multidimensional Scaling (PAMS) is an exploratory technique designed to extract the major profile patterns from data of interest. It also estimates each individual’s observed score profile based on the major profiles identified in PAMS. Specifically, PAMS quantifies the direction and magnitude of the match between the major profile patterns and the observed score profiles of people by regressing actual test scores of individuals to the major profiles.

The profile patterns identified by PAMS can be used to examine the relationship between the profile patterns and other criterion variables (Kim, Frisby, & Davison, 2001c) or used in regression (e.g., Kuang, 1998). Kim et al. (2001c) examined that the relationship between cognitive profile patterns and achievement test results in WJ-R data and found that there were significant relationships between the cognitive profile patterns and the achievement test results. Moreover, the level parameter was more strongly related with the achievement scores than the profile patterns. In this case the level parameter can be interpreted as Spearman’s g or general cognitive ability factor.

However, the results from a PAMS approach will only serve as an exploratory tool. The profile patterns identified in PAMS is exploratory, but those profile patterns can be examined in confirmatory analyses illustrated by this paper. The hypothesized model, according to the assumptions of the PAMS model, which did not include the error covariances between subscales was not supported. Based on the modification indices recommended in the SEM analyses, the error covariances were included for: ENT and INVES (r = .13), ART and REAL (r = .06), CONV and ART (r = -.09), SOCIAL and INVES (r = .32), and ENT and ART (r = .01), and then the model was supported (P-value for Chi-square = .25). Inclusion of the error paths among the
subscales made it possible to fit the data to the model. Note that the values in parentheses are correlation coefficients between subscales.

However, in general, the error terms should not correlated from one indicator to another. This is part of the definition of indicators of a construct. If the error terms for two or more indicators correlate, it means that the indicators measure something else in addition to the construct they are supposed to measure. If this is the case, the meaning of the construct and its dimensions may be different from what is intended.

Because of that concern, we examined intercorrelations between the subscales added with error paths. As shown above, it is interesting to note that all other correlations were trivial, except the one between SOCIAL and INVES (r = .32), although the error correlations were all highly significant. From these results, one can suspect the measurement error between the PAMS approach and the SEM method when the results of PAMS analyses were attempted to be mapped onto the SEM scheme to replicate the results. When this measurement error was considered by including the error paths of the subscales in SEM, the hypothesized factor structure (the level parameter factor plus two dimensional factors) was replicated in the SEM analysis.

While not introduced here, Kim (1999), Kim & Davison (2001a), Kim & Davison (2001b), and Kim, Craig, & Davison (2001c) have applied a bootstrapping technique to estimating standard errors of test parameters (i.e., MDS scale-values) in the PAMS model. In their simulation study, Kim & Davison (2001a) measured utility of the bootstrap procedure in estimating MDS scale-value standard errors, and reported that on average, irrespective of simulated conditions, the bootstrap error estimation included 80% of accuracy. Moreover, using the bootstrap procedure, Kim & Davison (2001b) proposed how to estimate standard errors of
correspondence indices (or person parameters) that quantify the match between individuals’ actual score profiles and major profiles in PAMS.

Estimating standard error of any parameters of interest has at least two important aspects in PAMS analyses. First, it can help identify statistically significant scale-value estimates, and then allows users to include only significant values in interpreting major profiles identified in PAMS. Considering the utility of standard error estimates, future research should include (bootstrap) standard error estimates of scale-values/coordinates in major profiles extracted in PAMS, identify significant coordinates, and include only significant ones in interpretation of the major profiles. In addition to that, when the major profiles identified in PAMS are replicated in SEM, one can examine whether significant test results of the coordinates in PAMS match those of factor loadings in SEM since SEM analyses always provide asymptotic standard errors corresponding to estimates of parameters.

This comparison allows us to check either reliability of bootstrap error estimate in PAMS or asymptotic error estimate in SEM when either of them is fixed to be a criterion. If there exits discrepancy between the bootstrap and the SEM error estimate, one may suspect that the asymptotic estimates would be underestimated. Regarding this issue, Kim (1999) and Weinberg, Carroll, and Cohen (1984) compared MDS scale-value’s asymptotic error (estimated the maximum likelihood method: ML) with bootstrap error estimates, and they consistently found that the ML estimates were significantly much more different from Monte Carlo results than the bootstrap estimates and the significant differences came mostly from underestimation.
References


Table 1

Interest Inventory Intercorrelations among Males: Minnesota Vocational Assessment Clinic Data

(N = 1308)

<table>
<thead>
<tr>
<th></th>
<th>REAL</th>
<th>INVES</th>
<th>ART</th>
<th>SOCIAL</th>
<th>ENT</th>
<th>CONV</th>
</tr>
</thead>
<tbody>
<tr>
<td>Realistic</td>
<td>1.00</td>
<td>0.52</td>
<td>0.06</td>
<td>0.18</td>
<td>0.30</td>
<td>0.37</td>
</tr>
<tr>
<td>Investigative</td>
<td>0.52</td>
<td>1.00</td>
<td>0.33</td>
<td>0.32</td>
<td>0.13</td>
<td>0.37</td>
</tr>
<tr>
<td>Artistic</td>
<td>0.06</td>
<td>0.33</td>
<td>1.00</td>
<td>0.26</td>
<td>0.01</td>
<td>-0.09</td>
</tr>
<tr>
<td>Social</td>
<td>0.18</td>
<td>0.32</td>
<td>0.26</td>
<td>1.00</td>
<td>0.38</td>
<td>0.34</td>
</tr>
<tr>
<td>Enterprising</td>
<td>0.30</td>
<td>0.13</td>
<td>0.01</td>
<td>0.38</td>
<td>1.00</td>
<td>0.36</td>
</tr>
<tr>
<td>Conventional</td>
<td>0.37</td>
<td>0.37</td>
<td>-0.09</td>
<td>0.34</td>
<td>0.46</td>
<td>1.00</td>
</tr>
</tbody>
</table>

Note. Data were collected by David J. Weiss and Rene Dawis, University of Minnesota, Vocational Assessment Clinic.
Table 2

MDS Scale-values from Two-Dimensional Nonmetric Solution based on Intercorrelations of Interest Subscales

<table>
<thead>
<tr>
<th>Subscales</th>
<th>Dimension 1</th>
<th>Dimension 2</th>
</tr>
</thead>
<tbody>
<tr>
<td>Realistic</td>
<td>0.11</td>
<td>1.21</td>
</tr>
<tr>
<td>Investigative</td>
<td>-0.25</td>
<td>0.96</td>
</tr>
<tr>
<td>Artistic</td>
<td>-2.14</td>
<td>-0.05</td>
</tr>
<tr>
<td>Social</td>
<td>-0.07</td>
<td>-1.14</td>
</tr>
<tr>
<td>Enterprising</td>
<td>1.09</td>
<td>-.94</td>
</tr>
<tr>
<td>Conventional</td>
<td>1.26</td>
<td>-0.03</td>
</tr>
</tbody>
</table>
Table 3

Factor Loadings in SEM and Scale-values in PAMS

<table>
<thead>
<tr>
<th>Subscales</th>
<th>Factor 1 (Level Parameter)</th>
<th>Factor 2 (E/C Profile)</th>
<th>Factor 3 (R/I Profile)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Realistic</td>
<td>0.58 (0.00)</td>
<td>0.00 (0.11)</td>
<td>1.01 (1.21)</td>
</tr>
<tr>
<td>Investigative</td>
<td>0.58 (0.00)</td>
<td>-0.21 (-0.25)</td>
<td>1.04 (0.96)</td>
</tr>
<tr>
<td>Artistic</td>
<td>0.58 (0.00)</td>
<td>-2.15 (-2.14)</td>
<td>0.00 (-0.05)</td>
</tr>
<tr>
<td>Social</td>
<td>0.58 (0.00)</td>
<td>0.00 (-0.07)</td>
<td>-1.44 (-1.14)</td>
</tr>
<tr>
<td>Enterprising</td>
<td>0.58 (0.00)</td>
<td>1.14 (1.09)</td>
<td>-0.61 (-0.94)</td>
</tr>
<tr>
<td>Conventional</td>
<td>0.58 (0.00)</td>
<td>1.22 (1.26)</td>
<td>0.00 (-0.03)</td>
</tr>
</tbody>
</table>

Note. Factor 2 and Factor 3 loadings were multiplied by 4.21 and 3.07, respectively, to be consistent with magnitude of scale-values in PAMS analyses.
Figure Captions

Figure 1. The Hypothesized Model in SEM by The Assumptions of the PAMS model

Figure 2. The Modified Model Allowed With Error Covariances Between Interest Subscales

Figure 3. Factor Profiles (Squares) in SEM Superimposed On Major Profiles (Circles) in PAMS
Chi-Square=240.57, df=11, P-value=0.00000, RMSEA=0.126
Chi-Square=7.90, df=6, P-value=0.24555, RMSEA=0.016
Connecting SEM and PAMS

Figure 3. Factor Profiles (Squares) in SEM Superimposed On Major Profiles (Circles) in PAMS

Note. Squares represent factor weights in SEM and circles represent scale-values in PAMS.