The Use of the Johnson-Neyman Confidence Bands and Multiple Regression Models To Investigate Interaction Effects: Important Tools for Educational Researchers and Program Evaluators.

When investigating the impact of predictor variables on an outcome variable or measuring the effectiveness of an educational program, educational researchers and program evaluators cannot ignore the possible influences of interaction effects. The purpose of this paper is to present a procedure that educational researchers can follow in order to increase their understanding of the nature of the interaction effect between a treatment variable and a continuous independent variable. This technique involves the use of three separate analytical techniques implemented in three steps. First, the interaction effect is statistically tested using a multiple regression model. Second, the interaction effect is plotted, and if the interaction effect is disordinal, the intersection point of the regression lines is calculated. Third, the Johnson-Neyman confidence limits are calculated. A list of the computer commands that can be used in conjunction with the Statistical Package for the Social Sciences (SPSS) PC+ computer software to calculate the Johnson-Neyman confidence limits is provided. In addition, this three-step analytical procedure is applied to a set of efficacy data that was collected in a study of the FOCUS instruction model developed by G. Russell (1992) to illustrate how it can be used by researchers and program evaluators. An appendix presents the computer program for the calculation of the confidence limits. (Contains 2 figures, 2 tables, and 16 references.) (Author/SLD)
The Use of the Johnson-Neyman Confidence Bands and Multiple Regression Models to Investigate Interaction Effects: Important Tools for Educational Researchers and Program Evaluators

John W. Fraas
Ashland University

Isadore Newman
The University of Akron

Paper presented at the Annual Meeting of the Eastern Educational Research Association,

February, 1997, Hilton Head, South Carolina

The authors of this paper thank Dr. Gary Russell for granting permission to use a portion of the data from his FOCUS study.
Abstract

When investigating the impact of predictor variables on an outcome variable or measuring the effectiveness of an educational program, educational researchers and program evaluators cannot ignore the possible influences of interaction effects. The purpose of this paper is to present a procedure that educational researchers can follow in order to increase their understanding of the nature of the interaction effect between a treatment variable and a continuous independent variable. This technique involves the use of three separate analytical techniques implemented in three steps. First, the interaction effect is statistically tested using a multiple regression model. Second, the interaction effect is plotted, and if the interaction effect is disordinal, the intersection point of the regression lines is calculated. Third, the Johnson-Neyman confidence limits are calculated. A list of the computer commands that can be used in conjunction with the SPSS/PC+ computer software to calculate the Johnson-Neyman confidence limits is provided. In addition, this three-step analytical procedure is applied to a set of efficacy data that was collected in a study of the FOCUS instructional model in order to illustrate how it can be used by researchers and program evaluators.
The Use of the Johnson-Neyman Confidence Bands and
Multiple Regression Models to Investigate Interaction Effects:
Important Tools for Educational Researchers and Program Evaluators

Most educational researchers and program evaluators are aware of the need to investigate the possible existence of interaction effects. When an interaction effect is being examined, a researcher or an evaluator must answer two questions. First, what analytical technique can be used to test for the presence of an interaction effect? Second, what analytical technique can provide the maximum amount of information regarding the interaction effect when, in fact, it exists? Researchers and evaluators often consider the first question. The second question, however, appears to be a consideration less often. To obtain an in-depth understanding of the interaction effect, the researcher or evaluator must utilize an analytical technique that can provide such information. That is, the researcher must avoid a Type VI error (Newman, Deitchman, Burkholder, Sanders, & Ervin, 1976), which occurs when the analytical technique does not provide the appropriate or necessary information.

In this paper, we present a three-step analytical procedure for examining an interaction effect between a treatment variable and a continuous independent variable. The first step in this analytical procedure, which was discussed in detail by McNeil, Newman, and Kelly (1996, pp. 127-140), requires the researcher to design models that are capable of statistically testing the interaction effect. The technique used in the second step, which was previously presented by Fraas and Newman (1977), Newman and Fraas (1979) and Pedhazur (1982, pp. 468-469), requires the researcher or program evaluator to calculate the point of intersection between the two regression lines. The third step requires that the Johnson-Neyman confidence bands be
calculated. This technique has been discussed by Johnson and Neyman (1936), Rogosa (1980, 1981), Chou and Huberty (1992), and Chou and Wang (1992).

In this paper, we are stressing the importance of using these techniques together in a three-step analytical procedure. The use of this analytical procedure will provide researchers and program evaluators with the type of information that will increase their understanding of the nature of the interaction effect being examined. To illustrate the type of information that is produced by this three-step analytical procedure, we have analyzed the personal and teaching efficacy levels of teachers who were exposed to an instructional model developed by Russell (1992), which is referred to as FOCUS.

Analytical Technique Applied to Efficacy Scores

Even though Russell (1992) believed that the exposure to the FOCUS model would increase the participants' levels of personal-teaching and teaching efficacy, he was not willing to assume that those increases would be constant across the participants' pre-term efficacy levels. That is, when comparing the post-treatment personal-teaching-efficacy and teaching-efficacy scores of the teachers who were exposed to the FOCUS model to teachers who were not exposed to the model, the differences may not be consistent across the ranges of the pre-term efficacy scores. Thus, to understand the possible influence of the FOCUS model on the personal-teaching-efficacy and teaching-efficacy scores of teachers, it was essential, not only to test for the existence of pre-term efficacy scores by group interaction effects, but also to gain insight into the nature of these interaction effects, if in fact, they did exist.
Subjects

Sixty-eight teachers who were enrolled in graduate level classes offered by the Education Department of Ashland University were included in the evaluation of the FOCUS model. Ashland University is located in north-central Ohio, which contains rural, suburban, and urban school systems. The courses, which required 36 hours of instruction, were offered during a summer term. Twenty-nine of the 68 teachers were not exposed to the FOCUS model. These 29 teachers, who taught in grade levels that ranged from kindergarten to the twelfth grade, served as the Control Group. The other 39 teachers were exposed to the FOCUS model during the same academic summer term. These 39 teachers, who also taught in grade levels that ranged from kindergarten through the twelfth grade, were designated as the treatment group. This treatment group was referred to as the FOCUS Group.

Instruments

Various instruments are used to measure the level of a teacher’s sense of efficacy. In this evaluation project, the Teacher Efficacy Scale, which was devised by Gibson and Dembo (1984), was used. This selection was consistent with the view expressed by Ross (1994) who stated in his extensive review of the teacher-efficacy research that:

Future researchers should treat the [teacher efficacy] construct as a multi-dimensional entity rather than a singular trait, examining personal and general teaching efficacy separately rather than aggregating them . . . [and they] should measure teacher efficacy with the most frequently used instruments to facilitate comparisons between studies (p. 27).
Each educator who participated in this study completed the Teacher Efficacy Scale at the beginning and end of the summer academic term. This instrument required each participant to rate each of 16 statements on a 1 (strong disagree) to 6 (strongly agree) scale. The ratings obtained from the first nine statements were summed to obtain a personal-teaching-efficacy score for each teacher. A high score on these nine statements was interpreted to mean that the teacher had a high level of personal-teaching efficacy. And a low score would indicate that the teacher had a low level of personal-teaching efficacy.

The other seven statements were used to measure a teacher’s teaching-efficacy score. The total score on these seven statements for each teacher was subtracted from 42. This procedure produced a teaching efficacy score that would be high for a teacher who had a high level of teaching efficacy. The score would be low for a teacher who had a low level of teaching efficacy.

Gibson and Dembo (1984) reported in their study that an analysis of internal consistency reliability values produced Cronbach’s alpha coefficient values of .78 and .75 for the personal-teaching-efficacy scores and teaching-efficacy scores, respectively. In addition, Gibson and Dembo stated that a multitrait-multimethod analysis supported both convergent and discriminant validity of the instrument.

**Hypotheses**

The null and research hypotheses that were statistically tested were as follows:

1. $H_0$: The interaction effect between the pre-term personal-teaching-efficacy scores and group membership does not account for some of the variation in the post-term personal-teaching-efficacy scores.
1H₁: The interaction effect between the pre-term personal-teaching-efficacy scores and group membership does account for some of the variation in the post-term personal-teaching-efficacy scores.

2H₀: The interaction effect between the pre-term teaching-efficacy scores and group membership does not account for some of the variation in the post-term teaching-efficacy scores.

2H₁: The interaction effect between the pre-term teaching-efficacy scores and group membership does account for some of the variation in the post-term teaching-efficacy scores.

Step 1: Statistical Tests of the Interaction Effects

Step 1 of the three-step analytic procedure was implemented for the efficacy data by statistically testing multiple linear regression models that were designed to measure the interaction effects. As part of this hypothesis testing procedure, the data utilized in each model were tested for possible outlier values with tests of Cook's distance measures (Neter, Wasserman, & Kutner, 1985). Any person who had a value that would distort the regression analysis was reviewed to determine whether the data for that person should be eliminated.

The model that was designed to test 1H₀ contained three independent variables. The teachers' post-term personal-teaching-efficacy scores served as the dependent variable for this model. One of the independent variables included in this model consisted of the teachers' pre-term personal-teaching-efficacy scores. This variable was labeled Pre-Term PTE. The second independent variable included in this model was the Group variable. This Group variable consisted of the values of zero and one. A value of one indicated that the teacher was in the
FOCUS Group, and a zero value meant that the teacher was in the Control Group. The third variable included in this model was formed by multiplying the Pre-Term PTE variable by the Group variable. The inclusion of this variable, which was labeled (Pre-Term PTE)X(Group), allowed us to use the regression model to calculate the difference between the slopes of the Control and FOCUS groups' regression lines.

The t-test value of the (Pre-Term PTE)X(Group) variable regression coefficient was used to test H₀. Since this study involved two dependent variables—personal-teaching efficacy and teaching efficacy—the alpha level for the t test of this regression coefficient value was set at .025, which is equal to .05 divided by 2. The chance of committing a type I error was reduced by using this alpha value (Newman & Fry, 1972).

Before the regression model was used to test H₀, Cook's distance measures were tested to determine whether the data for any teacher would tend to distort the regression results. The test results of Cook's distance measures indicated that none of the teachers was identified as having data that would be considered as outlier values. Thus, the data for all 68 teachers were included in an analysis of the regression model.

The results obtained from the analysis of the regression model are contained in Table 1. The t test of regression coefficient for the (Pre-Term PTE)X(Group) variable (t = -2.44, p = .0175) indicated that the difference between the slopes of the regression lines of the FOCUS and Control groups was statistically significant at the .025 level, that is, H₀ was rejected. Thus, the differences between the post-term personal-teaching-efficacy scores of the FOCUS and Control groups were not constant across the range of pre-term personal-teaching-efficacy scores.
The teaching efficacy scores served as the dependent variable in the regression model that was used to test $2H_0$. Similar to the previous regression model, this model included three independent variables. One of these independent variables was composed of the teachers' pre-term teaching-efficacy scores. This variable was labeled Pre-Term TE. A second independent variable included in the model was the Group variable. The third independent variable included in the model was generated by multiplying the Pre-Term TE variable by the Group variable. This variable, which was labeled $(\text{Pre-Term TE}) \times \text{(Group)}$, was used to estimate the difference between the slopes of the regression lines for the Control and FOCUS groups.

As part of the statistical testing process for $2H_0$, the data were tested for outlier values. The test results of Cook's distance measures indicated that the data recorded for one teacher may distort the results obtained from the regression analysis. After reviewing the data for this teacher, the data were deleted from the regression analysis. Thus, before $2H_0$ was statistically tested, the number of teachers in the FOCUS Group was reduced to 38. The values generated by the analysis of the regression model used to test $2H_0$ are listed in Table 2.

The $t$ test of the interaction effect between the Pre-Term TE variable and the Group variables ($t = 2.742, p = .008$) indicated that this interaction effect was statistically significant at the .025 level. Thus, the differences between the post-treatment teaching-efficacy scores of the FOCUS and Control groups were not constant across the range of pre-term personal-teaching-efficacy scores.
Step 2: Calculation of the Point of Intersection

The second step of the three-step analytical procedure was implemented by, first, graphing the interaction effect. If the interaction effect is disordinal, the point of intersection between the two regression lines would be calculated. If the interaction effect is ordinal, that is, the regression lines do not intersect in the relevant range, the researcher would proceed to Step 3.

The interaction effect between the Pre-Term PTE variable and the Group variable is diagramed in Figure 1. Since the interaction effect was disordinal, the point at which the two regression lines intersected was calculated as follows:

1. The value of zero was substituted for the Group variable in the regression equation contained in Table 1 to obtain the regression line for the Control Group.

\[
Y = 6.362 - .538*(\text{Pre-Term PTE})*(\text{Group}) + .852*(\text{Pre-Term PTE}) + 25.124*(\text{Group})
\]

\[
Y = 6.362 - .538*(\text{Pre-Term PTE})*(0) + .852*(\text{Pre-Term PTE}) + 25.124*(0)
\]

\[
Y = 6.362 + .852*(\text{Pre-Term PTE})
\]

2. The value of one was substituted for the Group variable in the regression equation contained in Table 1 to obtain the regression line for the FOCUS Group.

\[
Y = 6.362 - .538*(\text{Pre-Term PTE})*(\text{Group}) + .852*(\text{Pre-Term PTE}) + 25.124*(\text{Group})
\]

\[
Y = 6.362 - .538*(\text{Pre-Term PTE})*(1) + .852*(\text{Pre-Term PTE}) + 25.124*(1)
\]

\[
Y = 31.486 + .314*(\text{Pre-Term PTE})
\]
3. The two regression lines were set equal to each other and the researcher solved the equation for Pre-Term PTE.

\[ 6.362 + .852 \times \text{Pre-Term PTE} = 31.486 + .314 \times \text{Pre-Term PTE} \]

\[ .538 \times \text{Pre-Term PTE} = 25.124 \]

Pre-Term PTE = 46.7

As indicated by the results of this calculation and the graph of the disordinal interaction effect contained in Figure 1, the post-term personal-teaching-efficacy scores of the teachers in the FOCUS Group were higher than the post-term personal-teaching-efficacy scores of the teachers in the Control Group when their pre-term personal-teaching-efficacy scores were less than 47. The post-term personal-teaching-efficacy scores of the teachers in the Control Group, however, were higher than the post-term personal-teaching-efficacy scores of the teachers in the FOCUS Group when their pre-term personal-teaching-efficacy scores were greater than or equal to 47.

The interaction effect between the Pre-Term TE variable and the Group variable, which is diagramed in Figure 2, was also disordinal. Using the values produced by the regression analysis contained in Table 2, the point at which the two regression lines for the post-term teaching-efficacy scores intersected was calculated in the same manner as was the intersection point for the personal-teaching-efficacy scores. The calculations were as follows:
1. The value of zero was substituted for the Group variable in the regression equation contained in Table 2 to obtain the regression line for the Control Group.

\[
Y = 19.800 + .703\text{(Pre-Term TE)}\text{(Group)} + .153\text{(Pre-Term TE)} - 14.569\text{(Group)}
\]

\[
Y = 19.800 + .703\text{(Pre-Term TE)} \times (0) + .153\text{(Pre-Term TE)} - 14.569\times(0)
\]

\[
Y = 19.800 + .153\text{(Pre-Term TE)}
\]

2. The value of one was substituted for the Group variable in the regression equation contained in Table 2 to obtain the regression line for the FOCUS Group.

\[
Y = 19.800 + .703\text{(Pre-Term TE)}\text{(Group)} + .153\text{(Pre-Term TE)} - 14.569\text{(Group)}
\]

\[
Y = 19.800 + .703\text{(Pre-Term TE)} \times (1) + .153\text{(Pre-Term TE)} - 14.569\times(1)
\]

\[
Y = 5.231 + .856\text{(Pre-Term TE)}
\]

3. The two regression lines were set equal to each other and the researcher solved the equation for Pre-Term TE.

\[
19.800 + .153\text{(Pre-Term TE)} = 5.231 + .856\text{(Pre-Term TE)}
\]

\[
.703\text{(Pre-Term TE)} = 14.569
\]

\[
\text{Pre-Term TE} = 20.7
\]

The post-term teaching-efficacy scores of the teachers in the Control Group were greater than the post-term teaching-efficacy scores of the teachers in the FOCUS Group when their pre-term teaching-efficacy scores were below 21. In addition, the post-term teaching-efficacy scores of the teachers in the FOCUS Group were greater than the post-term teaching-efficacy scores of
the teachers in the Control Group when their pre-term teaching-efficacy scores were greater than or equal to 21.

After the intersection point is calculated in a study that investigates an interaction effect between a continuous independent variable and a treatment variable, it is important to note the percentage of the study's participants who have scores above and below the intersection point. For the efficacy data of the 67 teachers who were included in both analyses, the percentages of interest were as follows:

1. Seventy-two percent of the teachers had pre-term efficacy scores that corresponded to points on the regression lines where the teachers had higher post-term personal-teaching-efficacy scores and higher post-term teaching-efficacy scores when exposed to the FOCUS model. That is, 72% of the teachers had pre-term personal-teaching-efficacy scores that were lower than the intersection point of 46.7 and pre-term teaching-efficacy scores that were higher than the intersection point of 20.7.

2. Three percent of the teachers had pre-term efficacy scores that corresponded to points on the regression lines where the teachers had lower post-term personal-teaching-efficacy scores and lower post-term teaching-efficacy scores when exposed to the FOCUS model. That is, 3% of the teachers had pre-term personal-teaching-efficacy scores that were higher than the intersection point of 46.7 and pre-term teaching-efficacy scores that were lower than the intersection point of 20.7.

3. Nineteen percent of the teachers had pre-term efficacy scores that corresponded to points on the regression lines where the teachers had higher post-term personal-teaching-efficacy scores and lower post-term teaching-efficacy scores when exposed to the FOCUS model. That is,
19% of the teachers had pre-term personal-teaching-efficacy scores that were lower than the intersection point of 46.7 and pre-term teaching-efficacy scores that were lower than the intersection point of 20.7.

4. Six percent of the teachers had pre-term efficacy scores that corresponded to points on the regression lines where the teachers had lower post-term personal-teaching-efficacy scores and higher post-term teaching-efficacy scores when exposed to the FOCUS model. That is, 6% of the teachers had pre-term personal-teaching-efficacy scores that were higher than the intersection point of 46.7 and pre-term teaching-efficacy scores that were higher than the intersection point of 20.7.

Two points should be noted with respect to these percentages. The first point is that 72% of the study’s 67 participants who were included in both regression analyses had pre-term efficacy scores that would place them on the regression lines at points where the post-term efficacy scores were higher on both of the efficacy scales when exposed to the FOCUS model. The second point, which may modify the conclusions that are drawn from the analyses of the interaction effects in Step 2, is that the differences between the post-term efficacy scores of the FOCUS and Control groups may be statistically significant only for certain ranges of the pre-term efficacy scores. Thus, before conclusions are drawn with respect to who benefits and who does not benefit from being exposed to the FOCUS model, it is essential to determine the ranges of pre-term efficacy scores in which the differences between the post-term efficacy of the teachers in the FOCUS Group and the teachers in the Control Group are statistically significant. Step 3 of this three-step analytical procedure is designed to determine these statistically significant ranges.
Step 3: Calculation of the Johnson-Neyman Confidence Bands

The third step of the three-step analytical procedure requires that the Johnson-Neyman confidence limits be calculated for each statistically significant interaction effect. It should be noted that some researchers have argued that the Johnson-Neyman regions of significance are nonsimultaneous ones (Potthoff, 1964 and Rogosa, 1980, 1981). Based on empirical results by Chou and Huberty (1992) and Chou and Wang (1992), it appears that the Johnson-Neyman technique can be used to make simultaneous inferences provided that the slope homogeneity assumption is statistically tested and rejected. Since $1H_0$ and $2H_0$ were rejected, it was appropriate to calculate Johnson-Neyman (1936) confidence bands for the nonsignificance regions for the efficacy scores.

The program that was used to calculate the Johnson-Neyman confidence bands, which was used in conjunction with the SPSS/PC+ 4.0 software (SPSS Inc., 1990), is listed in the Appendix. The program, which calculates the Johnson-Neyman significance bands as suggested by Pedhazur (1982, pp. 169-171), requires that 12 values be provided. A description of the required values, as well as their labels, are as follows:

1. The symbol $ss1$ represents the pre-term sum of squares value for the Control Group.
2. The symbol $ss2$ represents the pre-term sum of squares value for the FOCUS Group.
3. The symbol $n1$ represents the sample size of the Control Group.
4. The symbol $n2$ represents the sample size of the FOCUS Group.
5. The symbol $sumresid$ represents the residual sum of squares value of the regression model.
6. The symbol mean\(1\) represents the mean of the pre-term scores of the Control Group.

7. The symbol mean\(2\) represents the mean of the pre-term scores of the FOCUS Group.

8. The symbol slope\(1\) represents the slope of the regression line for the Control Group.

9. The symbol slope\(2\) represents the slope of the regression line for the FOCUS Group.

10. The symbol intl represents the intercept point of the regression line for the Control Group.

11. The symbol int2 represents the intercept point of the regression line for the FOCUS Group.

12. The symbol fcrit represents the critical F value with 1 and \(N - 4\) degrees of freedom.

The sum of squares values, the sample sizes, and the mean values were obtained from the printout generated by the DESCRIPTIVE subprogram of the SPSS/PC+ (1990) computer software, with each of the two groups being analyzed separately. The residual sum of squares value, the slope values, and the intercept-point values were obtained from the printouts generated by the REGRESSION subprogram of the SPSS/PC+ computer software. The critical F value can be obtained from an F-Distribution Table. Note that for the evaluation of the personal efficacy scores, the numerator degrees of freedom \((df_n)\) and the denominator degrees of freedom \((df_d)\) values were 1 and 64 (68-4), respectively. For the teaching efficacy scores, the \(df_n\) and the \(df_d\) values were 1 and 63 (67-4), respectively. The confidence level was set at .95 for each set of limits. The data line of the program listed in the Appendix, which utilized the freefield format, contains the data used to generate the Johnson-Neyman confidence limits for the personal-teaching-efficacy scores. The data line used for the analysis of the teaching efficacy scores was as follows: 567.30 745.82 29 38 1334.32 23.24 24.71 .15 .86 19.80 5.23 4.00.
The upper limit for the 95% confidence bands for the personal-teaching-efficacy scores was 81.8, which was above the maximum score of 54 points on the personal-teaching-efficacy section of the Teacher Efficacy Scale. The lower limit was 40.7. Based on these limits, which are included in Figure 1, it can be concluded that the post-term personal-teaching-efficacy scores for the teachers in the FOCUS and Control groups were not statistically significantly different when their scores were greater than or equal to 41. The post-term personal-teaching-efficacy scores of the teachers in the Focus Group were statistically significantly higher than the corresponding scores of the teachers in the Control Group, however, when their pre-term scores were less than 41.

The lower limit of the 95% Johnson-Neyman confidence limits for the regression lines diagramed in Figure 2 was equal to 9.97, which was less than three points above the minimum score of 7 that a teacher could receive on the teaching-efficacy section of the Teacher Efficacy Scale. It should be noted that none of the teachers included in this analysis had a pre-term teaching-efficacy score below 13. Thus, none of the teachers had a score below the lower limit of the nonsignificance region. The upper limit of the nonsignificance region of the Johnson-Neyman 95% confidence limits for the pre-term teaching efficacy scores was 23.8. Thus, the post-term teaching-efficacy scores of the teachers in the FOCUS and Control groups were not statistically significantly different when their pre-term teaching-efficacy scores were less than 24. The post-term teaching-efficacy scores of the teachers in the FOCUS Group, however, were statistically significantly higher than the post-term teaching-efficacy scores of the teachers in the Control Group when their pre-term teaching-efficacy scores were equal to or greater than 24.
Group were statistically significantly higher than the post-term efficacy scores of the teachers in the Control Group on both efficacy scales.

2. Twenty-seven percent of the teachers had pre-term efficacy scores on both of the efficacy scales that were located inside the Johnson-Neyman confidence limits. That is, 27% of the teachers had pre-term personal-teaching-efficacy scores that were greater than or equal to 41 and pre-term teaching-efficacy scores that were less than 24. Thus, 27% of the teachers had pre-term efficacy scores that corresponded to the points on the regression lines where the post-term personal-teaching-efficacy scores and the post-term teaching-efficacy scores of the teachers in the FOCUS Group and the teachers in the Control Group were not statistically significantly different.

3. Twenty-one percent of the teachers had pre-term efficacy scores that were located below the lower limit of the Johnson-Neyman confidence limits on the personal-teaching-efficacy scale but inside or below the lower limit of the confidence limits for the teaching-efficacy scores. That is, 21% of the teachers had pre-term personal-teaching-efficacy scores that were less than 41 and pre-term teaching-efficacy scores that were less than 24. It should be noted, again, that none of the teachers had pre-term teaching-efficacy scores below the lower confidence limit of 9.97. Thus, 21% of the teachers had pre-term efficacy scores that corresponded to the points on the regression lines where the post-term personal-teaching-efficacy scores of the teachers in the FOCUS Group were statistically significantly higher than the scores of the teachers in the Control Group but the post-term teaching-efficacy scores of the two groups were not statistically significantly different.

4. Twenty-one percent of the teachers had pre-term efficacy scores that were located inside or above the upper limit of the Johnson-Neyman confidence limits on the personal-
That is, 21% of the teachers had pre-term personal-teaching-efficacy scores that were less than 41 and pre-term teaching-efficacy scores that were less than 24. It should be noted, again, that none of the teachers had pre-term teaching-efficacy scores below the lower confidence limit of 9.97. Thus, 21% of the teachers had pre-term efficacy scores that corresponded to the points on the regression lines where the post-term personal-teaching-efficacy scores of the teachers in the FOCUS Group were statistically significantly higher than the scores of the teachers in the Control Group but the post-term teaching-efficacy scores of the two groups were not statistically significantly different.

4. Twenty-one percent of the teachers had pre-term efficacy scores that were located inside or above the upper limit of the Johnson-Neyman confidence limits on the personal-teaching-efficacy scale but above the upper confidence limit on the teaching-efficacy scale. That is, 21% of the teachers had pre-term personal-teaching-efficacy scores that were greater than or equal to 41 and pre-term teaching-efficacy scores that were greater than or equal to 24. Again, it should be noted that the upper confidence limit for the pre-term personal-teaching-efficacy scores, which was 81.8, exceeded the maximum score of 54. Thus, 21% of the teachers had pre-term personal-teaching-efficacy scores and pre-term teaching-efficacy scores that corresponded to the points on the regression lines where the post-term teaching-efficacy scores of the teachers in the FOCUS Group were statistically significantly higher than the scores of the teachers in the Control Group but the post-term personal-teaching-efficacy scores of the two groups were not statistically significantly different.

It is essential to note that 73% of the study's 67 participants had pre-term efficacy scores that corresponded to points on the regression lines where the post-term efficacy scores of the
teachers in the FOCUS Group were statistically significantly higher than the scores of the teachers in the Control Group on at least one of the efficacy scales. In addition, the remaining 27% of the teachers had pre-term efficacy scores that corresponded to points on the regression lines where the post-term efficacy scores of the two groups were not statistically significantly different.

Implications Based on the Results of the Three-Step Analytical Procedure.

It is important to understand what each step in this three-step analytical procedure reveals about the interaction effects. The results of Step 1 indicate that both interaction effects were statistically significant. A more in-depth understanding of these interaction effects, however, is obtained by reviewing the information generated by Steps 2 and 3 of this three-step analytical procedure.

The graphs containing the interaction effects and the points of intersection between the regression lines for the personal-teaching-efficacy scores and the teaching-efficacy scores, which were completed in Step 2, revealed that both interaction effects were disordinal and the regression lines for the personal-teaching efficacy scores and the teaching-efficacy scores intersected at 46.7 and 20.7, respectively. These graphs and the intersection points appear to suggest that, with respect to their post-term efficacy scores, certain teachers would benefit from being exposed to the FOCUS model, while exposure to the FOCUS model would be detrimental to other teachers. In addition, these points of intersection could possibly be used to identify which teachers would and would not benefit from exposure to the FOCUS model. Before such a conclusion is reached, however, it is important to realize that the differences between the post-term efficacy scores of the teachers in the FOCUS and Control groups, who have pre-term scores near the intersection points, could simply be due to noise or random variation. That is, the post-term scores of the
students in the two groups are statistically significantly different only for pre-term scores that are located some distance above and below the intersection points. Thus, before one should draw a conclusion with respect to the nature of these interaction effects, it is essential to review the information provided by the Johnson-Neyman confidence limits calculated in Step 3.

The significance region between the two regression lines that were designed to analyze the post-term personal-teaching-efficacy scores included only the pre-term personal-teaching-efficacy scores that were less than 41. In addition, the significance region between the two regression lines that were designed to analyze the post-term teaching-efficacy scores included only the pre-term teaching-efficacy scores that were greater than or equal to 24. Thus, as indicated by the interaction effects contained in Figures 1 and 2, whenever the post-term efficacy scores of the two groups were statistically significantly different, the post-term efficacy scores of the Focus Group exceeded the post-term efficacy scores of the Control Group.

Thus, a majority of teachers (73%) had pre-term efficacy scores that placed them in ranges along the regression lines that indicated that the post-term efficacy scores of the teachers in the Focus Group, on at least one of the efficacy scales, were statistically significantly higher than the post-term efficacy scores of the teachers in the Control Group. It is important to also note that in spite of the fact that the interaction effects were disordinal, the reverse statement is not true. That is, none of the teachers had pre-term efficacy scores in the ranges along the regression lines that indicated that the post-term efficacy scores of the Focus Group were statistically significantly lower than the post-term efficacy scores of the Control Group on either of the two efficacy scales. The remaining 27% of the teachers had pre-term efficacy scores in the ranges
along the regression lines that indicated that the post-term efficacy scores of the FOCUS and Control groups were not statistically significantly different on either of the two efficacy scales. Based on this information, one would not use the intersection points between the regression lines to determine who would and who would not benefit from being exposed to the FOCUS model. Rather, it would be more appropriate, keeping in mind research design limitations, to suggest that, based on pre-term efficacy levels, exposing the teachers to the FOCUS model would be beneficial to the majority of teachers and it would not be detrimental to any one group of teachers. Researchers or program evaluators would reach this conclusion only by using this three-step analytical procedure.

Summary

It is important for educational researchers and program evaluators to increase their understanding of the interaction effects that may be present in their data. We believe that a more in-depth understanding of an interaction effect between a continuous independent variable and a treatment variable can be obtained if the researcher or program evaluator follows the three-step analytical procedure that was presented in this paper.

Two points should be noted regarding this three-step analytical procedure. First, the use of a multiple regression model to statistically test the interaction effect, which is undertaken in Step 1, is an essential analytical procedure when investigating an interaction effect. This test of the homogeneity of the slopes of the regression lines allows the researcher to not only to determining if the interaction effect is statistically significant, but it also permits simultaneous inferences to be made from the Johnson-Neyman confidence bands, which are calculated in the third step of this analytical procedure.
Second, the calculation of the intersection point between the two regression lines in Step 2 provides a researcher or program evaluator with information that could be used to identify groups of people who would benefit from being exposed to the treatment being investigated. It is important to realize, however, that the difference between the post-term scores of the students in the two groups who have pre-term scores that are located near this intersection point could be simply due to noise or random variation. That is, the post-term scores of the students in the two groups are statistically significantly different only for pre-term scores that are located some distance above and below that intersection point. The calculation the Johnson-Neyman confidence limits in Step 3 allows the researcher or program evaluator to determine the pre-term scores at which the post-term scores of the two groups are statistically significantly different. This information may lead the researchers or program evaluators to modify conclusions that were based solely on information provided by the analytical techniques contained in the first two steps of this process.

As was demonstrated by the analyses of the personal-teaching-efficacy and teaching-efficacy scores that were presented in this paper, following the three-step analytical procedure can provide essential information not only regarding whether an interaction effect does, in fact, exist but also with respect to the nature of the interaction effect. Such information can be invaluable to researchers and program evaluators.
References


Newman, I., Deitchman, R., Burkholder, J., Sanders, R., & Ervin, L. (1976). Type VI error: Inconsistency between the statistical procedure and the research question. Multiple Linear Regression Viewpoints. 6 (4), 1-19.


Appendix

Computer Program for the Calculation of the Johnson-Neyman Confidence Limits

Data list free/
   ss1 ss2 n1 n2 sumresid mean1 mean2 slope1 slope2 int1 int2 fcrit
Begin data.
1434.21 1821.59 29 39 2495.58 39.31 38.90 .85 .31 6.36 31.49 3.99
End data.
Compute term1 = (fcrit/(n1+n2-4))*sumresid.
Compute terma = term1*(-1).
Compute a = ((term1)*((1/ss1)+(1/ss2))+(slope1-slope2)**2.
Compute b = (term1*((mean1/ss1)+(mean2/ss2))+((int1-int2)*(slope1-slope2)).
Compute c = (terma)*(((n1+n2)/(n1*n2))+((mean1**2)/ss1)+
   ((mean2**2)/ss2)+((int1-int2)**2).
Compute RegionU = ((b*(-1))+sqrt(b**2-(a*c)))/a.
Compute RegionL = ((b*(-1))-sqrt(b**2-(a*c)))/a.
List RegionU RegionL.
Table 1

Regression Results for the Post-Term Personal-Teaching-Efficacy Scores

<table>
<thead>
<tr>
<th>Variable</th>
<th>Regression Model</th>
<th>1 Test</th>
<th>p Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>(Pre-Term PTE) X (Group)</td>
<td>-.538</td>
<td>-2.44</td>
<td>.018</td>
</tr>
<tr>
<td>Pre-Term PTE</td>
<td>.852</td>
<td>5.17</td>
<td>&lt;.000</td>
</tr>
<tr>
<td>Group</td>
<td>25.124</td>
<td>2.87</td>
<td>.006</td>
</tr>
<tr>
<td>Constant</td>
<td>6.362</td>
<td>.97</td>
<td>.338</td>
</tr>
</tbody>
</table>

R² = .370

Adjusted R² = .341

N = 68

Residual Sum of Squares = 2495.58

*Note.* The values for the Group variable are zero and one for teachers in the Control and FOCUS groups, respectively.
Table 2

Regression Results for the Post-Term Teaching-Efficacy Scores

<table>
<thead>
<tr>
<th>Variable</th>
<th>Regression Coefficient</th>
<th>t Test Value</th>
<th>p Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>(Pre-Term TE)X(Condition)</td>
<td>.703</td>
<td>2.742</td>
<td>.008</td>
</tr>
<tr>
<td>Pre-Term TE</td>
<td>.153</td>
<td>.790</td>
<td>.433</td>
</tr>
<tr>
<td>Condition</td>
<td>-14.569</td>
<td>-2.339</td>
<td>.023</td>
</tr>
<tr>
<td>Constant</td>
<td>19.800</td>
<td>4.331</td>
<td>&lt;.000</td>
</tr>
</tbody>
</table>

R² = .347

Adjusted R² = .316

N = 67

Residual Sum of Squares = 1334.318

Note. The values for the Group variable are zero and one for teachers in the Control and FOCUS groups, respectively.
Figure Captions

Figure 1. Pre-Term Personal-Teaching-Efficacy-Scores-by-Group Interaction.

Figure 2. Pre-Term Teaching-Efficacy-Scores-by-Group Interaction.
I. DOCUMENT IDENTIFICATION:

Title: The Use of the Johnson-Neyman Confidence Bands and Multiple Regression Models to Investigate Interaction Effects: Important Tools for Educational Researchers and Program Evaluators

Author(s): John W. Fraas  Isadore Newman

Corporate Source: Ashland University

Publication Date: February, 1997

II. REPRODUCTION RELEASE:

In order to disseminate as widely as possible timely and significant materials of interest to the educational community, documents announced in the monthly abstract journal of the ERIC system, Resources in Education (RIE), are usually made available to users in microfiche, reproduced paper copy, and electronic/optical media, and sold through the ERIC Document Reproduction Service (EDRS) or other ERIC vendors. Credit is given to the source of each document, and, if reproduction release is granted, one of the following notices is affixed to the document.

If permission is granted to reproduce the identified document, please CHECK ONE of the following options and sign the release below.

X Sample sticker to be affixed to document

Check here

"PERMISSION TO REPRODUCE THIS MATERIAL HAS BEEN GRANTED BY
Sample
TO THE EDUCATIONAL RESOURCES INFORMATION CENTER (ERIC):"

Level 1

Sample sticker to be affixed to document

"PERMISSION TO REPRODUCE THIS MATERIAL IN OTHER THAN PAPER COPY HAS BEEN GRANTED BY
Sample
TO THE EDUCATIONAL RESOURCES INFORMATION CENTER (ERIC):"

Level 2

or here

Permitting reproduction in other than paper copy.

Sign Here, Please

Documents will be processed as indicated provided reproduction quality permits. If permission to reproduce is granted, but neither box is checked, documents will be processed at Level 1.

"I hereby grant to the Educational Resources Information Center (ERIC) nonexclusive permission to reproduce this document as indicated above. Reproduction from the ERIC microfiche or electronic/optical media by persons other than ERIC employees and its system contractors requires permission from the copyright holder. Exception is made for non-profit reproduction by libraries and other service agencies to satisfy information needs of educators in response to discrete inquiries."

Signature: John W. Fraas
Position: Trustee's Professor
Organization: Ashland University
Telephone Number: (419) 289-5930
Date: 3/21/97

Address: 220 Andrews Hall
Ashland University
Ashland, Ohio 44805
III. DOCUMENT AVAILABILITY INFORMATION (FROM NON-ERIC SOURCE):

If permission to reproduce is not granted to ERIC, or if you wish ERIC to cite the availability of this document from another source, please provide the following information regarding the availability of the document. (ERIC will not announce a document unless it is publicly available, and a dependable source can be specified. Contributors should also be aware that ERIC selection criteria are significantly more stringent for documents which cannot be made available through EDRS).

<table>
<thead>
<tr>
<th>Publisher/Distributor:</th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>Address:</td>
<td></td>
</tr>
<tr>
<td>Price Per Copy:</td>
<td>Quantity Price:</td>
</tr>
</tbody>
</table>

IV. REFERRAL OF ERIC TO COPYRIGHT/REPRODUCTION RIGHTS HOLDER:

If the right to grant reproduction release is held by someone other than the addressee, please provide the appropriate name and address:

Name and address of current copyright/reproduction rights holder:

Name:

Address:

V. WHERE TO SEND THIS FORM:

Send this form to the following ERIC Clearinghouse:

AERA/ERIC
American Institute for Research
3333 K Street, NW
Washington, DC 20007

If you are making an unsolicited contribution to ERIC, you may return this form (and the document being contributed) to:

ERIC Facility
1301 Piccard Drive, Suite 300
Rockville, Maryland 20850-4305
Telephone: (301) 258-5500