

Preliminary investigation of the psychological sense of school membership scale with primary school students in a cross-cultural context

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Abstract

The Psychological Sense of School Membership (PSSM) scale has been used for more than 20 years to measure students' sense of school belonging, yet its psychometric properties have had limited examination with pre-adolescent children. This study investigated the

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utility and psychometrics of the PSSM in three primary school samples from the United States, China, and the United Kingdom. Exploratory factor analysis revealed good fit for a unidimensional factor structure in the US sample, which was subsequently confirmed in all three samples. Partial invariance across all three samples and full invariance across pairwise samples (United States and United Kingdom; United Kingdom and China) was found. Path analyses revealed significant positive relations of the PSSM total belonging score with gratitude and prosocial behavior, and significant negative relations with symptoms of distress. Future directions and implications are discussed.

Keywords

cross-cultural, primary school age, Psychological Sense of School Membership Scale, school belonging, school membership

Research over the past several decades shows that school climate influences students' psychological and academic well-being. As such, researchers and practitioners have been interested in assessing school climate with the goal of improving student functioning. School climate is conceptualized as a multidimensional construct that includes a school's culture, norms, values, and expectations (National School Climate Council, 2007). The dimensions of school climate have been measured in several ways, including assessing school belonging or psychological sense of school membership (Thapa, Cohen, Guffey, & Higgins-D'Alessandro, 2013).

This article aims to provide psychometric support for a widely used measure of school belonging — Psychological Sense of School Membership (PSSM; Goodenow, 1993). Although used with adolescent populations, few studies have investigated its use with primary school children (Cheung & Hui, 2003). Investigating the use of the PSSM with primary school children is particularly important as perceptions of school membership and belonging begin at an early age (Lerner et al., 2018; Whitney & Smith, 1993) and may differ from adolescents. Additionally, despite its use in cross-national research (e.g., Cowden, Govender, Asante, Reardon, & George, 2016; Gaete, Montero-Marin, Rojas-Barahona, Olivares, & Araya, 2016; Ye & Wallace, 2014), the PSSM has not yet been subject to needed investigations regarding its invariance across sociocultural contexts. This article addresses these limitations by examining the PSSM in primary schools across three countries: The United States, China, and the United Kingdom. This is a first step in a larger cross-national investigation to provide clarity on the use of the PSSM to measure students' sense of school belonging across the world.

School belonging: Defining terms, relations to other constructs, and assessment

School belonging has been conceptualized in multiple ways, including 'the extent to which students feel personally accepted, respected, included, and supported by

others in the school social environment' (Goodenow, 1993, p. 80) and 'feel[ing] close to, a part of, and happy at school' (Libbey, 2007, p. 52). Although school belonging is generally defined as feeling accepted by or a part of a school community, the construct of school belonging has also been related to, or used interchangeably with, a number of different terms including school connectedness, student engagement, and school bonding (Furlong, Froh, Muller, & Gonzalez, 2014). The Center for Disease Control (2009) views school connectedness as the 'belief by students that adults and peers in the school care about their learning as well as them as individuals' (p. 3). Student engagement is conceptualized by three types of engagement: cognitive, affective, and behavioral (Fredericks, Blumenfeld, & Paris, 2004). School bonding refers to the connection students feel to their school, school personnel, and academics (Maddox & Prinz, 2003). These constructs all are related to school belonging as a psychological construct; that is, students' mindsets about how they fit in and are a part of the broader school community.

Regardless of the term, school belonging, engagement, connectedness, and bonding have been linked with academic persistence and effort (Tinto, 1997), better educational outcomes (Goodenow, 1993), prosocial behaviors (Solomon, Battistich, Kim, & Watson, 1996), fewer mental health problems (Gaete, Rojas-Barahona, Olivares, & Araya, 2016), reduced bullying victimization (Bosworth, Espelage, & Simon, 1999), gratitude (Froh, Sefick, & Emmons, 2008), and lower rates of delinquency and dropout (Finn & Rock, 1997). Additionally, school belonging has been shown to be a protective factor against high-risk behaviors, including substance use and suicidality (Marraccini & Brier, 2017). These relations document the developmental advantages for students who feel a sense of belonging in their school community.

International studies show that school belonging affects students in similar ways in different cultural contexts. For example, Gaete and colleagues investigated school belonging's relation with mental health problems in students in Chile and found that school belonging was associated with reduced emotional, conduct, hyperactivity, and peer problems (Gaete et al., 2016). Khawaja, Allan, and Schweitzer (2017) found similar results for students in Australia, linking school connectedness to a variety of positive outcomes. These studies, among others, support the findings that school belonging has significant benefits for students.

A substantial body of research has investigated school climate and school belonging with adolescent or early adolescent populations (Cowden et al., 2016; You, Ritchey, Furlong, Shochet, & Boman, 2011). In contrast, few studies have examined the effects of school belonging on younger children in primary schools. Academic habits and social learning begin in primary school, lending support to the idea that interventions fostering school belonging in the primary grades could have profoundly positive consequences. Students at this age are beginning to view themselves as part of the educational institution and as people who belong to a school. In addition, primary school students generally stay with one teacher throughout the school day, implicating teacher support as particularly influential at this developmental stage. It is during this early developmental period that students begin to form

life-long mindsets related to school experiences, such as their identity as a 'student', a 'learner', and a 'citizen' of a school community (Lerner et al., 2018). For these reasons, further investigation of school belonging among primary school students, including validated measures for this age group, is needed.

Psychological Sense of School Membership (PSSM) Scale

Goodenow (1993) created a scale designed to measure adolescents' sense of school membership or belonging, which has been widely used across a variety of contexts. The PSSM includes 18 items (13 positively and five negatively worded items) measured on a five-point Likert-type scale (1 = not at all true and 5 = completely true). The items were initially validated on one suburban and two urban populations and were found to be positively correlated with academic success and motivation (Goodenow, 1993). Other studies have found that the PSSM is negatively correlated with emotional distress and behavioral problems (Shochet, Dadds, Ham & Montague, 2006).

Although Goodenow originally described the scale as one-dimensional, solely measuring the construct of school belonging, the scale's dimensionality was not investigated during the initial study. Subsequent studies have found varying factor structures across different samples, particularly with the five negatively worded items. Several studies found that the negatively worded items formed their own factor (Gaete et al., 2016; Hagborg, 1994; Ye & Wallace, 2014; You et al., 2011), which may be due to a method effect (DiStefano & Motl, 2009). Two of these studies also found that the scale had a three-factor structure, but, depending on the specific study, some items were not retained due to cross-loadings (e.g., Hagborg, 1994; You et al., 2011). Gaete and colleagues (2016) provided evidence that a one-factor solution was possible after removing the negatively worded items, which had strong cross-loadings. Similarly, O'Farrell and Morrison (2003) found a unidimensional structure after removing 13 of the 18 items, which were found to have cross-loadings. These studies involving adolescent samples provide evidence that there is ambiguity surrounding the factor structure and necessity of including all 18 PSSM items.

Few studies have examined the PSSM's factor structure with primary school children. In one previous, independent study, Cheung and Hui (2003) translated the PSSM into Chinese and modified the response scale to a six-item response scale for use with students in primary school. A principal components analysis provided evidence for a two-factor structure, with 13 positively worded items forming a *school belonging* factor and the five negatively worded items forming a *feeling of rejection* factor. No other studies were found that used confirmatory factor analysis to confirm the factor structure of the PSSM.

Current study

The current study investigated the factor structure of the PSSM using three primary school samples: (a) United States, (b) China, and (c) United Kingdom.

These samples were chosen to provide information that could better inform future studies involving cross-cultural and cross-national samples. Due to the inconsistencies of the factor structure of the PSSM in past studies and the lack of investigation in primary schools, this study examined the factor structure of the PSSM on primary school populations in cross-cultural contexts.

Since several previous studies found that the negatively worded PSSM items converge on a separate factor, we examined the possible contribution of negatively worded items for primary school age children. Ye and Wallace (2014) proposed that the negatively worded items form a unique factor, but fundamentally as a method artifact. This contrasts with the original PSSM formulation that included negatively worded items because, for Goodenow (1993), acceptance and rejection were considered to be opposite ends of a school belonging continuum — if an individual does not feel accepted, they must feel rejected and vice versa. An alternative conceptualization is provided by peer relationship studies that suggest acceptance and rejection are distinct, but related constructs (Bukowski, Sippola, Hoza, & Newcomb, 2000). Bukowski et al. (2000) state that 'acceptance refers to the number of strong positive links a child has with other members of the group; rejection refers to the number of negative links a child has with members of the group' (p. 12). Hence, while these two constructs are interrelated, they might not be polar opposites. Consequently, it is not expected that negatively and positively worded PSSM items would load onto one factor, even if items were reverse coded.

The literature on school belonging is muddled and it is unclear if it is one continuum with acceptance and rejection on opposite ends of the continuum, or whether two separate continua of acceptance and rejection form the construct. This study investigated the psychometric properties of the PSSM in cross-cultural primary school populations to examine its factor structure and further clarify this construct of school belonging in the crucial developmental context of primary school. In addition, this study examined whether a subset of items could efficiently measure primary school students' perceptions of school belonging to support research and educational practices in primary school populations across the world.

Method

Participants

All students were between the ages of 8-years-old and 13-years-old. Mean ages are as follows: United States = 10.23 (SD = 0.87), China = 10.38 (SD = 1.05), United Kingdom = 9.19, (SD = 1.00).

The US sample consisted of 782 primary school students aged 8-years-old (0.1%), 9-years-old (22%), 10-years-old (37.8%), 11-years-old (35.5%), 12-years-old (2.9%), and 13-years-old (0.8%) from seven schools in a small urban school district in California. The sample was made up of 50.8% female and 49.0% male students. At home, the majority of the sample spoke English (71.5%). Other students spoke Spanish (6.9%), English and Spanish (15.8%), or another language (5.8%) at home.

All students were proficient English speakers and were able to read and respond to survey questions. The ethnicity breakdown, as reported by students, is: Caucasian/White (40.5%), Hispanic/Latino/a/Mexican (21.6%), Native American or American Indian (8.6%), Asian American (5.4%), Black or African American (2.7%), Pacific Islander (1.3%), Other ethnicity (16%), missing response (3.9%).

The Chinese sample consisted of 1,178 students aged 8-years-old (1.1%), 9-years-old (19.1%), 10-years-old (33.1%), 11-years-old (33.5%), 12-years-old (11%), and 13-years-old (1.3%), from five schools in Changsha and Yongzhou, Hunan Provinces in China. This sample included both suburban and urban school districts. The sample consisted of 60.5% male students and 39.5% female students.

The UK sample consisted of 522 students aged 8-years-old (30.8%), 9-years-old (30.5%), 10-years-old (27.2%) and 11-years-old (11.5%), from three primary schools in an urban area of London. The sample was made up of 48.7% male and 51.3% female students. About half (51%) of the students spoke English at home, with the remaining students indicating that they spoke another language at home, including Hindi, Arabic, and Polish. All students were proficient English speakers and were able to read and respond to survey questions.

Procedure

Researchers across the three countries discussed the need for a measure of school belonging in primary schools. Following a literature review and discussions regarding instruments for use in the various contexts, the researchers decided to embark on a cross-national study to examine the PSSM. The researchers initially determined the desire for a unidimensional, brief, and efficient measure of school belonging, which was consistent with Goodenow's originally stated rationale for the PSSM.

The participating school sites were part of a convenience sample based on the researchers' connections with different schools. For the US and UK samples, the survey was administered in English. For the Chinese sample, the PSSM was translated into simplified Chinese by two researchers and independently back translated into English by two other researchers. All four researchers were fluent in English and Chinese. Further, four teachers from elementary schools in China reviewed the Chinese version to evaluate its readability for elementary school students; they confirmed the readability and no additional modifications were made.

Local policies and procedures related to consent for surveys were followed. The US sample required active consent procedures, while the Chinese and UK samples allowed passive consent procedures. The procedures for each country are described in the Online Supplemental Material.

Instruments

Psychological Sense of School Membership (PSSM). The PSSM (Goodenow, 1993) was the focal instrument for this study. Two modifications were made to the scale to be more developmentally appropriate for the younger students in this study. First, the

original five-point response scale was modified to a simplified six-point response scale in a yes/no format (1 = no, never, 2 = no, almost never, 3 = yes, sometimes,4 = yes, often, 5 = yes, very often, and 6 = yes, always) to reduce the cognitive burden for primary school students. This modification is consistent with another PSSM study which also used a six-point response scale for primary school students (Cheung & Hui, 2003). Second, the items were modified into question form rather than the original statement form. Questions as opposed to statements are considered more developmentally appropriate (Woolfolk, 2004). Most of the PSSM item content remained unchanged: (e.g., original item = 'People at this school are friendly to me'; modified item = 'Are people at this school friendly to you?'). However, two of the originally negatively worded items were converted into positive wording after the modification: (original item = 'Sometimes I don't feel as if I belong here'; modified item = 'Do you belong at your school?'; original item = 'Teachers here are not interested in people like me'; modified item = 'Are teachers interested in students like you?'). These items were changed to avoid a double negative thought process (e.g.: 'Do you feel as if you don't belong at your school?'). Thus, the survey used in this study consisted of 15 positively worded and three negatively worded questions.

In summary, all 18 items were modified to a question format and the response scale was changed to match the question format. The modifications were originally conducted in English. The survey was then translated into simplified Chinese and back translated into English. The Online Supplemental Material includes the English and Chinese PSSM items.

Convergent Validity Instruments. See Online Supplemental Material for additional details.

Social and Emotional Health Survey – Primary (SEHS-P). Four gratitude items and four prosocial items from the SEHS-P (Furlong, You, Renshaw, O'Malley, & Rebelez, 2013) were used to assess convergent validity because of school belonging's positive associations with gratitude (Froh et al., 2008) and prosocial behavior (Solomon et al., 1996). The items used the same six-point response scale as the PSSM. The internal consistency across the three samples were: Gratitude ($\alpha = 0.68-0.69$) and prosocial behavior ($\alpha = 0.80-0.82$).

Me and My School Questionnaire (M&MS). For the US and UK samples, the M&MS (Deighton et al., 2013) was used as a measure of divergent validity as school belonging has been found to be inversely related to distress. The M&MS has 16 items that assess emotional and behavioral difficulties and uses a three-point response format: never, sometimes, always. For this study, the alpha internal consistency for the total score for US and UK samples were 0.76 and 0.77, respectively.

Strengths and Difficulties Questionnaire (SDQ). For the Chinese sample, the SDQ total difficulties score was used as a measure of divergent validity as a measure of

psychological distress (Goodman, 1997). The SDQ uses a three-item response scale: not true, somewhat true, certainly true. The alpha internal consistency for the SDQ total difficulties score was 0.84.

Statistical analyses

Prior to investigating the factor structure, assumptions of normality were tested by analysing Q-Q plots, boxplots, and descriptive statistics (see Table 1). The same item in the US and UK samples (16. Do you wish you were in a different school?) and one item in the Chinese sample (17. Are you proud to be a part of your school?) displayed high skewness and kurtosis values (acceptable skewness and kurtosis values are between -3.0 and +3.0; Tabachnick & Fidel, 2012). These items were included in the analysis using maximum likelihood with robust standard errors, which helps minimize standard errors. All other items met assumptions.

Next, the US sample was split in half and an exploratory factor analysis (EFA) of all 18 items was conducted using Mplus 6.1 (Muthén & Muthén, 2010). Analyses were first conducted on the US sample given that the PSSM was originally developed in the United States and the interest was to see if this model would fit well with other cultures. Given the lack of literature regarding the psychometrics of the PSSM in primary schools, an EFA was conducted on the first randomly generated split sample. An oblique Geomin rotation was chosen to find a simple solution due to the fact that the survey items are similar and may correlate with one another, which is common in psychological research (Fabrigar, Wegener, MacCallum, & Straham, 1999). Then, a confirmatory factor analysis (CFA) was run on the other half of the US sample, the Chinese sample, and the UK sample to confirm the factor solution found in the EFA.

Next, measurement invariance tests across all three samples were completed to evaluate if the solution found for the US sample was replicated in the UK and Chinese samples. Establishing measurement invariance involves fitting a series of CFA models, each with different levels of invariance and comparing model fit of the models (Meredith, 1993). The first step is configural invariance to ensure that the factor model fits each group independently. Next, item loadings are constrained to be equal across groups while all other model parameters are free. Then, item intercepts and loadings are fixed to test for scalar invariance. Scalar invariance is an important step in measurement invariance testing; if it holds, latent mean comparisons can be made. Nested models are compared using chisquare difference test ($\Delta \chi^2$; Chen, 2007), in addition to the change in CFI (Δ CFI) and change in RMSEA, where values of $\Delta CFI \le 0.01$ and $\Delta RMSEA \le 0.015$ support invariance (Chen, 2007; Cheung & Rensvold, 2002). Lastly, path analyses to examine convergent and discriminant validity using positive and negative mental health indicators were conducted using items from the SEHS-P, M&MS, and SDQ.

Table 1. Descriptive statistics for retained PSSM items for the US, Chinese, and UK samples.

PSSM Question	Sample	M (SD)	Skewness (SE)	Kurtosis (SE)
I. Do you feel like yo	ou are a real þa	rt of your school?		
, , ,	US .	4.61 (1.43)	-0.72 (0.09)	-0.56 (.18)
	Chinese	4.62 (1.54)	-0.72 (0.07)	-0.64 (00.14)
	UK	4.82 (1.42)	-0.94 (0.11)	-0.28 (0.22)
5. Are most teachers	at your school	interested in you?		
	US	4.19 (1.32)	-0.27 (0.09)	-0.69 (0.18)
	Chinese	3.17 (1.44)	0.33 (0.07)	-0.55 (0.14)
	UK	3.94 (1.39)	0.07 (0.11)	-0.91 (0.22)
6. Do you belong at	your school?			
	US	5.12 (1.25)	-I.38 (0.09)	1.12 (0.18)
	Chinese	4.19 (1.74)	-0.43 (0.07)	-1.13 (0.14)
	UK	5.06 (1.45)	-I.36 (0.II)	0.64 (0.22)
7. Is there a teacher	or other adult a	it school that you co	n talk to if you have o	problem?
	US	4.84 (1.44)	-0.99(0.09)	-0.18 (0.18)
	Chinese	3.02 (1.69)	0.46 (0.07)	-0.92 (0.14)
	UK	5.01 (1.48)	-1.19 (0.11)	0.46 (0.22)
8. Are people at sch	ool friendly to yo	u?		
	US	4.84 (1.14)	-0.72 (0.09)	-0.40 (0.18)
	Chinese	4.23 (1.42)	-0.23 (0.07)	-0.89(0.14)
	UK	4.70 (1.35)	-0.72 (0.11)	-0.53 (0.22)
10. Are you included	in a lot of school	ol activities?		
	US	4.60 (1.32)	-0.69 (0.09)	-0.47 (0.18)
	Chinese	3.51 (1.71)	0.18 (0.07)	-1.19 (0.14)
	UK	4.37 (1.48)	-0.41 (0.11)	-I.00 (0.22)
II. Are you treated	with as much re	spect as other stude	ents?	
	US	4.72 (1.33)	-0.83 (0.09)	0.16 (0.18)
	Chinese	3.76 (1.66)	-0.06 (0.07)	-1.11 (0.14)
	UK	4.41 (1.56)	-0.51 (0.11)	-0.89 (0.22)
13. Can you be your	self at school?			
	US	4.84 (1.40)	-I.03 (0.09)	0.04 (0.18)
	Chinese	4.44 (1.65)	-0.65 (0.07)	-0.82(0.14)
	UK	4.76 (1.56)	-0.99 (0.11)	-0.26 (0.22)
17. Are you proud to	be a part of yo	our school?		,
•	US	4.87 (1.34)	-0.92 (0.09)	0.27 (0.18)
	Chinese	3.81 (1.84)	-0.14 (0.07)	-1.37 (0.14)
	UK	5.14 (1.34)	-1.39 (0.11)	0.75 (0.22)

Note: The minimum and maximum responses for each item are I and 6, respectively, for all three samples. See Online Supplemental Material for a list of all 18 PSSM items.

Results

Exploratory and confirmatory factor analyses

With the goal of a brief measure with a unidimensional factor structure, a series of EFAs were run on the first US randomly generated split sample, followed by CFAs separately for the split samples US, Chinese, and the UK. The first EFA was run with all 18 items and then subsequent EFAs were run following item deletion based on Howard's (2016) method: Items that loaded onto primary factors at 0.40 and above were retained, given that they did not load onto any alternate factors above 0.30 and that the difference between the primary and alternate factor loadings was at least 0.20. All other items were deleted.

Model fit criteria was based on Hu and Bentler (1999) who state that CFI values above 0.95 indicate good fit and values above 0.90 indicate adequate fit. Similarly, Browne and Cudeck (1989) found that RMSEA and SRMR values less than 0.05 suggest good fit and values up to 0.08 suggest reasonable fit. Using these parameters and the Howard (2016) method for item deletion, the iterative EFA process resulted in item deletion of seven items with cross-loadings or low factor loadings, one item, which had repetitive wording to another item, and one remaining negatively worded item for a total of nine items deleted (see Online Supplemental Material for information about how the number of factors was determined).

A final EFA was completed using the remaining nine items, which produced a one-factor solution that had good fit, $\chi^2(27) = 71.55$, p < 0.01; RMSEA = 0.06, SRMR = 0.04, CFI = 0.95. A CFA was conducted on the second split portion of the US sample and confirmed that the remaining nine items fit well onto a one- $\chi^2(27) = 74.22$, solution, p < 0.01; RMSEA = 0.07, SRMR = 0.03, CFI = 0.96, indicating that the nine items measured a unidimensional latent trait. Factor loadings remained strong ($\lambda = 0.52-0.79$; see Table 2). Additionally, for the final model, eigenvalues were greater than one and parallel analysis indicated a one-factor solution. Lastly, CFAs were performed to examine if the final US sample factor structure fit for the Chinese and UK samples. Analyses indicated good fit and high factor loadings (Chinese sample: $\chi^2(27) = 231.67$, p < 0.01; SRMR = 0.04, CFI = 0.94; UK: $\chi^2(27) = 65.37$, p < 0.01; RMSEA = 0.08, RMSEA = 0.05, SRMR = 0.03, CFI = 0.99; see Table 2 for factor loadings).

Factorial invariance

Measurement invariance testing was conducted in three steps: configural, metric, and scalar. Configural invariance analysis revealed good model fit (i.e., RMSEA = 0.07, SRMR = 0.04, CFI = 0.96), supporting configural invariance across the US, Chinese, and UK samples. Next, metric invariance was tested and confirmed (Δ CFI < 0.01, Δ RMSEA < 0.015), indicating that the constraints did not lead to a meaningful change in fit between Model 2 and Model 1 and providing support for metric invariance across all three countries. Lastly, scalar invariance was tested. Fit indexes, Δ CFI, and Δ RMSEA indicated that there were meaningful

	Sampl	es		
	US EFA	US CFA	Chinese CFA	UK CFA
I. Do you feel like you are a real part of your school?	0.79	0.78	0.64	0.79
5. Are most teachers in your school interested in you?	0.51	0.60	0.55	0.78
6. Do you belong at your school?	0.77	0.79	0.69	0.62
7. Is there a teacher or some other adult at school that you can talk to if you have a problem?	0.52	0.52	0.48	0.89
8. Are people at school friendly to you?	0.68	0.66	0.62	0.84
10. Are you included in a lot of school activities?	0.64	0.54	0.63	0.60
11. Are you treated with as much respect as other students?	0.58	0.58	0.69	0.71
13. Can you be yourself at school?	0.61	0.64	0.61	0.64
17. Are you proud to be a part of your school?	0.75	0.79	0.66	0.90

Table 2. Factor loadings for exploratory and confirmatory factor analyses for each sample with geomin rotation for one-factor solution.

Note: All loadings are significant at the p < 0.05 level. See Online Supplemental Material for a list of all 18 PSSM items.

changes between Model 3 and Model 2 and that full invariance across all three countries could not be achieved. Since there was evidence of metric invariance across all three countries, the countries were then tested for scalar invariance pairwise. Pairwise scalar invariance analyses revealed scalar invariance between the US and UK samples (Δ RMSEA < 0.015), as well as between Chinese and UK samples (Δ CFI < 0.01, Δ RMSEA < 0.015). While the delta CFI value (Δ CFI = 0.016) for the US and UK pair was slightly larger than the desired value of 0.01 for invariance, the delta RMSEA indicated invariance, and, therefore, full invariance was assumed. Scalar invariance between the US and Chinese samples was not confirmed (Δ CFI > 0.01, Δ RMSEA > 0.015). These analyses provide evidence of partial measurement invariance across all three samples and full measurement invariance across two of the three pairs of countries (i.e., US and UK; China and UK). The analyses also suggest that the modified PSSM construct differed in the US and Chinese samples. Model fit indexes can be seen in Table 3.

Additional psychometric analyses

Reliability and validity checks were performed for the final nine items of the PSSM separately in all three samples: US sample ($\alpha = 0.87$), Chinese sample ($\alpha = 0.85$), UK sample ($\alpha = 0.83$). These results support high internal consistency for the final version of the PSSM.

To examine convergent and discriminant validity, path analyses were conducted using positive (gratitude and prosocial behavior) and negative (internal and

Table 3. Model fit indices for measurement invariance.

Invariance tests	χ^2	дþ	$df = \Delta \chi^2$	abla q t	RMSEA (90% CI)	SRMR	CFI	ΔCFI	ARMSEA
Configural	436.102 ($p < 0.001$)	18		1	0.074 (0.067, 0.080)	00.035	00.958	Ι	
Metric	571.002 ($\rho < 0.001$)	65	134.900**	91	0.078 (0.072, 0.084)	0.062	0.944	0.014	0.004
Scalar	1190.879 ($p < 0.001$)	49	***18.619	91	0.109 (0.103, 0.114)	0.102	0.874	0.067	0.031
Scalar (U.S. & U.K.)	287.001 ($\rho < 0.001$)	20	284.001**	2	0.070 (0.062, 0.079)	0.062	96.0	910.0	800.0
Scalar (U.S. & Chinese)	1032.196 ($p < 0.001$)	20	461.194**	2	0.120 (0.114, 0.127)	0.112	0.824	0.120	0.042
Scalar (U.K. & Chinese)	385.377 ($\rho < 0.001$)	20	185.625**	2	0.073 (0.066, 0.080)	0.054	0.949	0.005	0.005

1000 > 0**

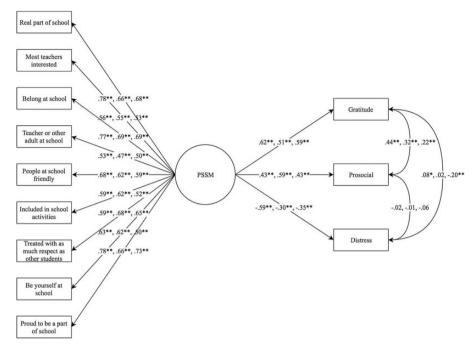


Figure 1. Psychological Sense of School Membership convergent and discriminant validity model for the US, Chinese, and UK samples, respectively. Path coefficients are completely standardized (path coefficients should be interpreted separately as full measurement invariance was not achieved across all three samples, *p < 0.05. **p < 0.001.)

external distress) mental health indicators. The path analyses revealed significant positive relations of the PSSM total belonging score with the SEHS-P gratitude mean score and prosocial mean score, and significant negative relations with the M&MS mean score and SDQ total difficulties mean score. All models had adequate fit to the data (see Figure 1).

Discussion

This study investigated the factor structure of the PSSM on primary school populations and examined measurement invariance for three cross-cultural groups. Results of this study were similar to previous studies with adolescent students (Gaete et al., 2016; Hagborg, 1994; Ye & Wallace, 2014; You et al., 2011) and supported previous research, which showed that a reduced number of PSSM items can be used to measure students' sense of school belonging. Some may raise questions about whether this is a new instrument; however, this was not our intent. Rather, the intent was to find an efficient tool that can be used to measure primary students' sense of school belonging across different cultures. The findings from this

study are consistent with Goodenow's (1993) original objective of creating a unidimensional measure of school belonging.

Exploratory factor analyses concluded that several items had cross-loadings and low factor loadings. Following systematic item deletion, the remaining positively worded items formed a nine-item unidimensional measure for school belonging. CFAs revealed good model fit for all three populations. Measurement invariance analyses confirmed partial invariance across the US, Chinese, and UK samples, as well as full invariance across two pairs of countries: US and UK, and UK and China. Path analyses provided evidence for convergent and discriminant validity of the PSSM.

School personnel in each of these countries (i.e., US, China) may consider using this nine-item PSSM measure to analyse school belonging for primary school populations. However, additional research is needed prior to recommending this tool be used to compare students' sense of school belonging internationally. The measurement invariance results reveal that there are were some differences in how the school belonging construct is being measured across all three countries. These differences might be due to cultural variations in the conceptualization of school belonging or other procedural variations such as survey format (online versus paper). The reasons for the differences are beyond the scope of this study and may benefit from additional analyses in future research. However, the results of pairwise scalar invariance across the US and UK samples, and across the UK and Chinese samples show promising results that a cross-national measure of school belonging for primary school students may be possible following additional research exploring the cross-cultural differences in school belonging.

The original content of the PSSM included some items that focused on students' sense of rejection. These items have not held up well in subsequent analyses, and have often formed a second factor. If the literature on peer relationships and the information on the method effects of negatively worded items are considered, school belonging should be a positive construct of belonging to a school in terms of peer relationships, teacher relationships, and the larger school community. Rejection may need to be considered separately when measuring for school belonging through the use of a rejection measure (e.g., Social Peer Rejection Measure; Lev-Wiesel, Sarid, & Sternberg, 2012). The results of this study suggest that the negatively worded items on the PSSM measure a separate construct than do the positively worded items. Thus, the PSSM may benefit from only including positively worded items, as seen in this study.

This study contains limitations with respect to the structural and external validity evidence, and also provides insight into areas for future needed research. Although measurement invariance analyses were conducted across the different countries, there is no knowledge about invariance across different ages, ethnicities, and genders. It should also be noted that modifications of the measure were made. Although these modifications appeared to be developmentally appropriate, further analysis would be beneficial in order to confirm that the modifications work well with additional primary school populations. Additionally, with respect to the procedures, the

US sample used an online survey format, while the UK and Chinese samples used a paper administration format. This difference may have influenced measurement invariance results, such that the US sample had more differences with the Chinese samples than with UK sample. Further, due to the fact that the delta CFI value for the US and UK pair was slightly large, future studies should also validate if there is indeed full invariance between the two countries. Finally, the measure of emotional distress used in the path analyses differed for the Chinese sample, which may have affected the path analyses and, subsequently, the discriminant validity analysis of the measure. Further research is needed to explore the role of rejection in students' perceptions of belonging at school and analyse whether the construct of school belonging is unidimensional or multidimensional.

School belonging's many influences on academics, mental health, and overall well-being of students make it a critical construct for investigation (Bosworth et al., 1999; Finn & Rock, 1997; Gaete et al., 2016). As such, valid and appropriate measures of school belonging are necessary for all age groups, including primary school students. This study provided a preliminary investigation of a version of the PSSM modified for use with primary school students that was found to have high internal consistency and strong convergent and divergent validity evidence. This measure also had invariance across multiple national samples, providing support for further exploring its use in a variety of contexts across the world.

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Supplemental material

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References

Bosworth, K., Espelage, D. L., & Simon, T. R. (1999). Factors associated with bullying behavior in middle school students. *The Journal of Early Adolescence*, *19*, 341–362. doi:10.1177/0272431699019003003

Browne, M. W., & Cudeck, R. (1989). Single sample cross-validation indices for covariance structures. *Multivariate Behavioral Research*, 24, 445–455. doi:10.1207/s15327906mbr2404 4

- Bukowski, W. M., Sippola, L., Hoza, B., & Newcomb, A. F. (2007). Pages from a sociometric notebook: An analysis of nomination and rating scale measures of acceptance, rejection, and social preference. New Directions for Child and Adolescent Development, 2000(88), 11–26. doi:10.1002/cd.23220008804
- Centers for Disease Control and Prevention. (2009). School connectedness: Strategies for increasing protective factors among youth. Atlanta, GA: US Department of Health and Human Services.
- Chen, F. F. (2007). Sensitivity of goodness of fit indexes to lack of measurement invariance. Structural Equation Modeling: A Multidisciplinary Journal, 14, 464–504. doi:10.1080/10705510701301834
- Cheung, G. W., & Rensvold, R. B. (2002). Evaluating goodness-of-fit indexes for testing measurement invariance. Structural Equation Modeling, 9, 233–255. doi:10.1207/ S15328007SEM0902 5
- Cheung, H. Y., & Hui, S. K. F. (2003). Mainland immigrant and Hong Kong local students' psychological sense of school membership. *Asia Pacific Education Review*, 4, 67–74. doi:10.1007/bf03025553
- Cowden, R. G., Govender, K., Asante, K. O., Reardon, C., & George, G. (2016). Validation of the perceived sense of school membership scale. *Journal of Psychoeducational Assessment*, 15, 1–7. doi:10.1177/0734282916678495
- Deighton, J., Tymms, P., Vostanis, P., Belsky, J., Fonagy, P., Brown, A., et al. (2013). The development of a school-based measure of child mental health. *Journal of Psychoeducational Assessment*, 31, 247–257. doi:10.1177/0734282912465570
- DiStefano, C., & Motl, R. (2009). Further investigation method effects associated with negatively worded items on self-report surveys. Structural Equation Modeling: A Multidisciplinary Journal, 13, 440–464. doi:10.1207/s15328007sem1303_6
- Fabrigar, L. R., Wegener, D. T., MacCallum, R. C., & Strahan, E. J. (1999). Evaluating the use of exploratory factor analysis in psychological research. *Psychological Methods*, *3*, 272–299. doi:10.1037/1082-989X.4.3.272
- Finn, J. D., & Rock, D. A. (1997). Academic success among students at risk for school failure. *Journal of Applied Psychology*, 82, 221–234. doi:0021-9010/97/
- Fredericks, J. A., Blumenfeld, P. C., & Paris, A. H. (2004). School engagement: Potential of the concept, state of the evidence. *Review of Educational Research*, 74, 59–109. doi:10.3102/00346543074001059
- Froh, J. J., Sefick, W. J., & Emons, R. A. (2008). Counting blessings in early adolescents: An experimental study of gratitude and subjective well-being. *Journal of School Psychology*, 46, 213–233. doi:10.1016/j.jsp.2007.03.005
- Furlong, M. J., Froh, J., Muller, M., & Gonzalez, V. (2014). The role of student engagement in engaged living and psychological and social well-being: The centrality of connectedness/relatedness. In D. J. Shernoff & J. Bempechat (Eds.), *National Society for the Study of Education Yearbook—engaging youth in schools: Empirically-based models to guide future innovations*. New York, NY: Columbia Teachers College.
- Furlong, M. J., You, S., Renshaw, T. L., O'Malley, M. D., & Rebelez, J. (2013). Preliminary development of the Positive Experiences at School Scale for elementary school children. *Child Indicators Research*, 6, 753–775. doi:10.1007/s12187-013-9193-7

- Gaete, J., Montero-Marin, J., Rojas-Barahona, C. A., Olivares, E., & Araya, R. (2016). Validation of the Spanish version of the Psychological Sense of School Membership (PSSM) scale in Chilean adolescents and its association with school-related outcomes and substance use. *Frontiers in Psychology*, 7, 1–11. doi:10.3389/fpsyg.2016.01901
- Goodenow, C. (1993). The psychological sense of school membership among adolescents: Scale development and educational correlates. *Psychology in the Schools*, *30*, 79–90. doi:10.1002/1520-6807(199301)30:1
- Goodman, R. (1997). The Strengths and Difficulties Questionnaire: A research note. Child Psychology and Psychiatry and Allied Disciplines, 38, 581–586. doi:10.1111/j.1469-7610.1997.tb01545.x
- Hagborg, W. J. (1994). An exploration of school membership among middle- and highschool students. *Journal of Psychoeducational Assessment*, 12, 312–323. doi:10.1177/ 073428299401200401
- Howard, M. C. (2016). A review of exploratory factor analysis decisions and overview of current practices: What are we doing and how can we improve? *International Journal of Human-Computer Interaction*, 32, 51–62. doi:10.1080/10447318.2015.1087664
- Hu, L., & Bentler, P. M. (1999). Cutoff criteria for fit indexes in covariance structure analysis: Conventional criteria versus new alternatives. Structural Equation Modeling: A Multidisciplinary Journal, 6, 1–55. doi:10.1080/107055199095400118
- Khawaja, N. G., Allan, E., & Schweitzer, R. D. (2017). The role of school connectedness and social support in the acculturation in culturally and linguistically diverse youth in Australia. *Australian Psychologist*. First published online:13 October 2017. doi:10.1111/ap.12327
- Lerner, R. M., Lerner, J. V., Geldhof, G. J., Gestsdóttir, S., King, P. E., & Sim, A. T., et al. (2018). Studying positive youth development. In J. E. Lansford & P. Banati (Eds.), *Handbook of adolescent development research and its impact on global policy* (pp. 68–82). New York, NY: Oxford University Press.
- Lev-Wiesel, R., Sarid, M., & Sternberg, R. (2012). Measuring social peer rejection during childhood: Development and validation. *Journal of Aggression, Maltreatment, and Trauma*, 22, 482–492. doi:10.1080/10926771.2013.785456
- Libbey, H. P. (2007). School connectedness: Influence above and beyond family connectedness. United States: UMI. Available online: www.cpc.unc.edu/projects/addhealth/publications/database/4145
- Maddox, S. J., & Prinz, R. J. (2003). School bonding in children and adolescents: Conceptualization, assessment, and associated variables. Clinical Child and Family Psychology Review, 6, 31–49. doi:1096-4037/03/0300-0031/0
- Marraccini, M. E., & Brier, Z. M. (2017). School connectedness and suicidal thoughts and behaviors: A systematic meta-analysis. *School Psychology Quarterly*, *32*, 5–21. doi:10.1037/spq0000192
- Meredith, W. (1993). Measurement invariance, factor analysis and factorial invariance. *Psychometrika*, 58, 525–543. doi:10.1007/BF02294825
- Muthén, L. K., & Muthén, B. O. (1998-2013). *Mplus user's guide* 6th ed. Los Angeles, CA: Muthén & Muthén.
- O'Farrell, S. L., & Morrison, G. M. (2003). A factor analysis exploring school bonding and related constructs among upper elementary students. *The California School Psychologist*, 8, 53–72. doi:10.1007/BF03340896
- Shochet, I., Dadds, M., Ham, D., & Montague, R. (2006). School connectedness is an underemphasized parameter in adolescent mental health: Results of a community

- prediction study. *Journal of Clinical Child and Adolescent Psychology*, 35, 170–179. doi:10.1207/s15374424jccp35
- Solomon, D., Battisch, V., Kim, D., & Watson, M. (1996). Teacher practices associated with students' sense of classroom as a community. Social Psychology of Education, 1, 235–267. doi:10.1007/BF02339892
- Tabachnik, B. G., & Fidel, L. S. (2012). *Using multivariate statistics* (6th ed.). Boston, MA: Pearson.
- Thapa, A., Cohen, J., Guffey, S., & Higgins-D'Alessandro, A. (2013). A review of school climate research. *Review of Educational Research*, 83, 357–385. doi:10.3102/0034654313483907
- Tinto, V. (1997). Classrooms as communities: Exploring the educational character of student persistence. *The Journal of Higher Education*, 68, 599–623. doi:10.2307/2959965
- Whitney, I., & Smith, P. K. (1993). A survey of the nature and extent of bullying in junior/middle and secondary schools. *Educational Research*, 35, 3–25. doi:10.1080/0013188930350101
- Woolfolk, A. (2004). Educational psychology (9th ed.). Boston, MA: Allyn & Bacon.
- Ye, F., & Wallace, T. L. (2014). Psychological Sense of School Membership Scale: Method effects associated with negatively worded items. *Journal of Psychoeducational Assessment*, 32, 202–215. doi:10.1177/0734282913504816
- You, S., Ritchey, K., Furlong, M. J., Shochet, I., & Boman, P. (2011). Examination of the latent structure of the Psychological Sense of School Membership Scale. *Journal of Psychoeducational Assessment*, 29, 225–237. doi:10.1177/0734282910379968

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