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AUTHOR Featherman, David L.
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ABSTRACT

In a secondary analysis of longitudinal data for white, originally second-parity couples resident in 7 large Standard Metropolitan Areas in 1957 (Princeton Fertility Study), the processes of educational, occupational, and economic achievements of a subset of 715 males are considered. Of primary concern is the extent of rearing in farm or rural as opposed to urban residential contexts as these influence educational attainment up to the initial interviews in 1957 and the successive socioeconomic achievement during the followup interval, 1957-67. For metropolitan males with farm or rural contexts of rearing, an educational handicap is incurred to their occupational and economic careers which men of urban backgrounds do not suffer. The fewer years of schooling attained by urban males of nonurban backgrounds is explicable in part by the relatively larger sizes of their families of origin. Moreover, when father's occupational status, size of the family of origin, and years of schooling completed are controlled statistically, the residential variable has no direct net effects on successive occupational and income career achievements. These findings are generated from regression and path analytical models of the process of stratification. Supplementary analyses within background residential categories consider potential effects of differential fertility on the socioeconomic achievements and potential interaction effects. (Author/EL)

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FARM, RURAL, AND URBAN BACKGROUNDS OF REARING AND SOCIOECONOMIC CAREER ACHIEVEMENTS IN METROPOLITAN STRATIFICATION SYSTEMS*

David L. Featherman

Department of Rural Sociology
The University of Wisconsin, Madison 53706

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Abstract

In a secondary analysis of longitudinal data for white, originally second-parity couples resident in seven large SMAs in 1957 (Princeton Fertility Study), the processes of educational, occupational, and economic achievements of the males are considered. Of primary concern is the extent of rearing in farm or rural as opposed to urban residential contexts as these influence educational attainment up to the initial interviews in 1957 and the successive socioeconomic achievements during the follow-up interval, 1957-67.

For these metropolitan males, those with farm or rural contexts of rearing incur an educational handicap to their occupational and economic careers which men of urban backgrounds do not suffer. The fewer years of schooling attained by urban males of non-urban backgrounds is explicable in part by the relatively larger sizes of their families of origin, even among men of equivalent socioeconomic backgrounds. Moreover, when father's occupational status, size of the family of origin, and years of schooling completed are controlled statistically, the residential variable has no direct, net effects on successive occupational and income career achievements. These findings are generated from regression and path analytic models of the process of stratification. Supplementary analyses within background residential categories consider potential effects of differential fertility on the socioeconomic achievements and potential interaction effects.

The objective of this paper is to assess the impact of farm or rural and urban residential contexts during child-rearing on the process of stratification (social mobility) within the metropolitan sector of the population. More specifically, the primary aspect of metropolitan stratification systems after which the ensuing discussion inquires is the permeability (Svalastoga, 1965) of the occupational structure, or the degree to which the social destination of a son is dependent upon the social status of his family of origin. In highly permeable occupational structures, the father-son occupational correlation approaches zero, and social mobility is extensive; social mobility is maximal, however, in much less permeable societies where the father-son occupational correlation is nearly -1.00. Herein, the treatment of the process of occupational achievement is less a discussion of social mobility than it is an inquiry about the permeability of metropolitan occupational structures for white urban males from differing residential backgrounds. The paper studies the achievement of both occupational and economic (income from salaries and wages) statuses over the first half of the work career as a function of social and residential background factors and educational attainments.

Recent researches of socioeconomic achievement for representative samples of the American male population (Blau and Duncan, 1967; Duncan, Featherman, and Duncan, 1968) provide the basic processual models for the ordering of variables into theoretical or causal schemes (Blalock, 1961; 1969; Duncan, 1966). Blau and Duncan's (1967) path analytic models of occupational achievement as of 1962 for several age cohorts illustrate that most of the total relationship (zero order correlation) between paternal and filial occupational statuses could be explained by the educational attainment of the son, as an intervening variable between socioeconomic background and occupational destination. Years of formal education completed proves to be the most important single variable in estimating

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current occupational achievement (in 1962), even for men with statistically equivalent father's occupation, paternal education, and own first job status.

Moreover, education is the critical variable in understanding the differential achievement of farm and nonfarm males (based on whether father was employed in farm occupations at the son's age of 16) in the Blau and Duncan data. The means for current (1962) occupational statuses for men with farm and nonfarm backgrounds are 26.2 and 40.1, respectively, on the Duncan SEI occupation scale (Reiss et al., 1961). These gross differences are largely a product of socio-economic background differences (paternal occupational and educational statuses) and the educational handicap associated with farm background (Blau and Duncan, 1967: 290-292).

Others using alternative definitions of the farm and nonfarm populations support the basic conclusion of Blau and Duncan. Haller (1968) asserts that the educational component of occupational achievement explains a substantial part of farm-born and nonfarm-born achievement differentials for men. Using 1960 Census residential designations, Hathaway, Beegle and Bryant (1968: 150) conclude that rural-farm males (and to a lesser degree, rural-nonfarm males) are underrepresented in those occupational groups requiring higher educational qualifications, especially when compared to the urban residential group. Thus, regardless of the operational designation of farm and nonfarm aggregates, there appears to be substantial agreement on the explanation of gross occupational differences.

In this paper still another designation of farm, rural and urban sub-populations is applied - that of the primary residential context in which current metropolitan (as of 1957) dwellers lived when they were growing up. That large numbers of formerly farm or rural males are now located in metropolitan areas

is a function both of urbanization of areas once considered rural and of selective off-farm migration, processes treated elsewhere in detail. We wish to inquire into the origins of the relative handicaps and benefits of rural and farm or of urban rearing contexts on the processes of educational, occupational, and economic achievements of white metropolitan males. We assume that there will exist gross differences in socioeconomic achievements owing to socioeconomic background and educational disparities; these discrepancies in attainment are probably of a smaller magnitude than for the current urban and rural populations as a result of selective migration by age, education, and employment qualifications.

Data and Methodology

Data are generated from the three panels of the Princeton Fertility Study (PFS) of second-parity white couples resident in seven of the largest Standard Metropolitan Areas as of 1957. Only a subset of this stratified random sample is selected for analysis herein; viz., those males who remained eligible throughout the duration of the restudy period (1957-67) and who returned a questionnaire in supplement to the original interview of the wife in 1957. This subset includes 715 males and represents 98% of the total remaining in the sample in the follow-up period.

The achievement variables measure years of schooling, prestige of occupational titles on the 1947 NORC scale (NORC, 1947), and dollars of salaries and wages of the husband. Background variables represent paternal occupational prestige (NORC, 1947), the number of siblings in the man's family or orientation, and the residential context during rearing. On the latter variable, men reared in urban areas (self-reported) and neither of whose parents grew up in rural areas are scored one (1); men reared in farm or rural areas and whose parents were reared in a similar residential context are given a score of six (6). Other combinations

of son's and parents' rearing contexts are scored as in Table 1. The resulting variable represents something approximating the extent of farm or rural influence

[Table 1 about here]

during the years of childhood socialization, an influence with documented impact on the process of stratification (Haller, 1966; Haller and Sewell, 1967). Note in Table 1 that of the few men with fathers in farming occupations, nearly all are in the highest categories of the farm or rural rearing variable.

Although the bulk of the ensuing correlation and regression analysis is presented tabularly, we conceive of a causal process of stratification illustrated schematically in Figure 1.

[Figure 1 about here]

Variables arrayed vertically along the left margin are not analyzed causally, for the moment; they are defined as predetermined variables. Although no arrows for causation are included in Figure 1, we assume a completely recursive model, i.e., where each antecedent variable is causally prior to each successive variable as one moves from left to right. We accept as the arbitrary criterion of significance for each causal relationship (indexed below by net regression coefficients in metric or in standard forms) a factor which is twice the absolute size of the coefficient's standard error.

Social and Residential Backgrounds and the Process of Socioeconomic Achievement

In Table 2 we observe the relative contributions of background and antecedent variables to educational completion and successive occupational and economic statuses in the longitudinal careers of the PFS males. From the pattern of

partial regression coefficients in standard form (Beta coefficients), we can describe the process of stratification in a manner similar to that offered by Blau and Duncan (1967), who employed a cross-sectional design and synthetic cohort models.

[Table 2 about here]

The net effect of paternal occupational status (X) is most direct in the early years of a man's career-- in the determination of the extent of formal schooling ($\beta_{U,X:F,S} = .302$) and the occupational statuses in the first years of family building, say prior to the age of 30 ($\beta_{W,X:F,S,U} = .098$; $\beta_{Y1,X:,F,S,U,W} = .077$);¹ no significant direct, net effects of X are apparent for any of the income variables. Thereafter, the gross relationship between paternal status and son's socioeconomic attainments is essentially an indirect one, determined by decomposing each zero-order correlation into the direct and indirect expression of X on the dependent achievement variables.² An exception to this pattern is noticed for occupational status at the end of the restudy period (Y3), where the Beta coefficient for the direct effect of X (.084) is nearly as large as corresponding coefficients for occupations earlier in the career, some 10 to 15 years prior. In previous analysis, these "lagged effects" of social background status on the course of the occupational career have not been explained by occupational inheritance, by unmeasured elements such as paternal education,

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1. The average age of men at marriage in the PFS subset is 21.9 years, and the average ages at panels I, II, and III are 29.6, 32.6, and 37.7, respectively.
 2. Correlations are decomposed according to the basic theorem of path analysis as described by Duncan: $r_{ij} = \sum_q \rho_{iq} r_{jq}$, "where i and j denote two variables in the system and the index q runs over all variables from which paths lead directly to X_i " (Duncan, 1966:5).

or by motivational products issuing from the family of socialization (Featherman, 1969). Aside from these as yet unexplained lagged direct effects, father's occupation influences the son's career most directly in the early years; but mainly the effects of X are indirect, through years of schooling (primarily) and through successive intervening occupational achievements.³

The other two background variables - size of the family of orientation, or number of siblings (S) and the extent of farm or rural influence during rearing (F)-- have virtually no statistically significant net, direct effects on the socioeconomic career. However, Table 2 does indicate that years of schooling are fewer for men from larger family sizes, regardless of paternal occupational status and residential context of rearing ($\beta_{U,S;X,F} = -.236$). If we hypothesize that both X and F are correlated but causally antecedent to S, and that all three background statuses are causally prior to U, then we can construct a path diagram (Duncan, 1966) to illustrate the direct and indirect effects of these three background statuses on socioeconomic achievement (say at panel I).

Figure 2 shows that metropolitan males with greater exposure to farm or rural (rather than urban) residential influences during rearing tend to come from slightly lower than average social status origins ($r_{XF} = -.071$).⁴ Regardless of paternal occupational status, however, men of farm or rural backgrounds have larger numbers of siblings, both as a function of their residential background

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3. Calculated by subtracting the coefficient for the direct effect (Beta coefficient or path coefficient) from the total correlation, as per the basic theorem of path analysis.
 4. This correlation is a smaller negative value than would arise were occupations coded in units of Duncan's SEI scale (Reiss et al., 1961: 264) rather than in prestige score units, which elevate the status of farm occupational titles above that estimated by mean levels of education and income.

(primarily) and their slightly lower paternal occupational statuses

($r_{SF} = .231 = p_{SF} + p_{SX}^r r_{XF} = .222 + .009$). Extent of farm or rural residential background does not affect educational completion directly (p_{UF} if non-significant), but men from larger families are handicapped in the competition for years of

[Figure 2 about here]

schooling regardless of paternal occupational status ($p_{US} = -.236$). Decomposing the correlation between residential background and educational attainment demonstrates that the educational handicap of farm or rural origins may be due to the association of these residential rearing contexts and a larger structure of the family of origin ($r_{UF} = -.119 = p_{UF} + p_{UX}^r r_{XF} + p_{US} p_{SF} + p_{US} p_{SX}^r r_{XF}$) = $(-.044) + (-.021) + (-.052) + (-.002$). Not only is the effect of the residential context variable on education expressed indirectly through family size, but its effects on occupation and income are also indirect, since there are no significant paths from F to $Y1$ or $I1$.

Thus the previously reported differentials in socioeconomic achievement by residential background appear to be a function of educational completion which in turn may be explained in part by the larger family sizes among the farm and rural reared and the negative relationship between number of siblings and years of schooling. In these data for currently metropolitan residents, no direct effects of residential background or number of siblings on occupational or economic achievements are observed; all effects are indirect through education.

Education and Socioeconomic Achievements

Both Table 2 and Figure 2 illustrate the major role of formal schooling as an intervening variable between social and residential backgrounds and successive

socioeconomic achievements. In Figure 2, it is the strong paths $p_{Y1,U}$ (.593), $p_{I1,U}$ (.180), and $p_{I1,Y1}p_{Y1,U}$ (.129) which transmit the effects of background to occupational and economic statuses at the initial interview. For example, the effect of father's occupation on that of the son at panel I, within (net of) residential background categories, is .129 directly ($p_{Y1,X}$), and .391 indirectly through education [$p_{Y1,U}(p_{UX} + p_{US}p_{SX})$]. Similar evidence for the importance of education as an intervening variable between backgrounds and successive socioeconomic attainments in the follow-up period could be calculated from Table 2. In all cases, the conclusions of Blau and Duncan (1967) about the role of education in cross-sectional data are supported in the longitudinal data.

Table 2 also illustrates the pattern of direct, net effects of education on successive occupational and economic attainments over the restudy period. We see the Beta coefficients decline in magnitude for the occupational dependent variables (.552 to .113) and the rise in the size of the net coefficients linking education and income statuses (.009 to .127), even though the rise is not perfectly monotonic. For men of statistically similar residential and social backgrounds, education is most important for the determination of occupation in the early years of the work career, and thereafter the effects become more and more indirect through prior occupational achievements, as greater experience and job tenure accrue. The lack of both experience and tenure in the earlier years of the career help explain the lower coefficient from education to income at marriage than to later income statuses. Young men in the labor force are somewhat handicapped by their inexperience and perhaps incur some discrimination not encountered by older men with equivalent formal educational qualifications. In any case, years of schooling yield greater net economic benefits later in

the careers of these metropolitan males than they do earlier, while the occupational prestige benefits (net) appear sooner in the careers and then decline.

Prior Attainments and Later Socioeconomic Achievements

Finally in Table 2 we see that occupational achievements, for white males at least, can be thought of in terms of careers, as having a history. For example, in the column for occupation at panel II (Y2) we observe significant net coefficients for both occupation at panel I (.539) and for occupation at marriage (.106). These data are consistent with the observation of Hodge (1966) and others that occupational mobility is not a simple Markovian, stochastic process. Nor can we model adequately the process of occupational achievement with simple causal chains, since the status at any time is a function of statuses more than one stage back in time; i.e., occupational achievements have histories (Featherman, 1969).

Similar analysis for income variables is less clear to interpret, the net coefficients being of lower size. Also, the prediction equations for economic statuses result in lower R^2 values than for occupational statuses. Both these facts suggest that there is less stability in the economic structure than in the occupational prestige structure of metropolitan stratification systems.

The Process of Socioeconomic Achievement Within Residential Background Contexts

The previous sections have included the residential background variable in multiple regression equations and have permitted us to look for linear relationships, between two variables within the residential variable by statistical control and between the residential variable and achievement variables. Linear relationships involving residential background are of low magnitudes, which

helps account for the non-significant net effects of F on the achieved statuses. Tables 3-5 examine possible interactions involving residential context of rearing.⁵ Table 3 is for men of urban background; Table 4, for men whose parents were reared in farm or rural contexts but who themselves were reared in urban areas; Table 5, for men with farm or rural rearing.

In these tables, patterns of effects of social background on educational completion show some fluctuation in the net regression coefficients for father's occupation. Among the urban-reared, each unit of paternal occupational status adds another .46 year of schooling (net of effects of family size), while the same change in paternal status results in only .28 year increase for men of equal family size and reared in farm or rural areas. That father's occupation has slightly greater gross impact on educational completion among the urban reared

[Tables 3 - 5 about here]

than among the non-urban-reared is seen in the zero-order correlations (r_{XU}): .38, .37, .29, for the urban-reared, for those with non-urban parents but with self reared in urban places, and for the farm or rural-reared, respectively.

Considerable irregularity in the magnitude of coefficients for the determination of occupational statuses is evident in Tables 3-5. Many of the discrepant coefficients fail the criterion of statistical significance, however, and make tenuous most interpretations of differences across residential background subgroups. Still, there does appear some modest degree of similarity among the interpretable coefficients for each of the successive occupational statuses in the three tables. We find little substantial evidence for interactions between the variables in our models for occupational achievement and the residential background classification.

5. Each table contains partial regression coefficients (unstandardized), and each represents the process of stratification within a different portion of the residential variable: Table 3 represents category 1 of Table 1; Table 4, categories 2 and 3; and Table 5, categories 4-6 of Table 1.

The most irregularity between Tables 3-5 is found for income achievements, but here the abilities of the prediction equations for the variation in economic statuses within subgroups are highly diminished (see R^2 values). Also, many more of the net regression coefficients are not statistically significant and therefore uninterpretable; few monotonic trends can be observed in the other coefficients, as we examine Tables 3-5 successively.

We do note what initially appears to be evidence of interaction in the columns for income at panel III, specifically for the antecedent occupational statuses W and Y1. However, we are cautioned by indications of multicollinearity among W, Y1, Y3 and the dependent variable in Table 5, as suggested in the correlation matrix (not reported here) for men of farm or rural rearing; such multicollinearity weakens any substantive interpretation of the large positive (.955.) and negative (-.661.) values for Y1 and W in Table 5 (column 9).

There is one possible interaction involving income at panel III and number of siblings (S). Men of farm or rural rearing (Table 5, column 9) earn \$476 less than expected as a net result of each additional sibling, whereas the panel III earnings of men from urban residential backgrounds (Tables 3 and 4, column 9) are virtually unaffected (net effects) by the size of their families of origin. The zero-order correlations ($r_{T3,S}$) for (a) men of urban, (b) non-urban parents but self urban, and (c) farm or rural backgrounds are -.13, -.15, and -.29, respectively. Thus there is the possibility that the number of siblings in one's family of origin directly influences later career earnings (net of paternal status, education, and prior achievements), but only for metropolitan men who were reared in farm or rural areas. We suggest this interpretation tentatively, owing to our small subgroups; if substantiated, this provisional finding

would supplement the earlier discussion of the importance of family size in explaining the effect of farm or rural background on career achievements in metropolitan stratification systems.

Marital Fertility and Socioeconomic Achievements for Residential Context of Rearing Subgroups

The idea that the structure of the family bears some impact on social mobility is not new (Dumont, 1890). So far we have discussed only the relationship of the size of the family of orientation to socioeconomic achievements over a portion of the life cycle, but data exist on cumulative marital fertility (number of live births) as it affects the occupational and economic attainments of the PFS men. These data speak to the relationship between the structure of the family of procreation and social mobility. A more detailed examination of the "mobility hypothesis" (linking social mobility to smaller family sizes) appear elsewhere (Featherman, 1970). Suffice it to say that the PFS data do not support the traditional forms of the mobility hypothesis, as it has been interpreted in contemporary research (Westoff, 1953).

The relevance of the mobility hypothesis may be doubted when applied in contemporary, largely metropolitan, populations. These lack the greater farm to nonfarm mobility observed in earlier decades and which quite likely constituted an important source of the asserted negative relationship between fertility and mobility (Blau and Duncan, 1967: 390). We might argue, however, that metropolitan males reared in nonurban contexts were exposed to different family size norms and selective factors associated with migration to urban areas. These could result in an alteration on the role of fertility in the process of achievement, when compared to the relationship of fertility to mobility for the urban-reared.

In Table 6 we note rather different patterns of correlation coefficients between cumulative marital fertility to panel III and background and achievement variables for the three residential context of rearing subgroups. It is only for men of farm or rural rearing backgrounds that the customary negative correlations between fertility and social status background, education, and socioeconomic achievement are observed to be statistically significant.⁶ An exception arises for the occupational variable at panel III (Y3), where fertility and occupational achievement are positively related among men of urban backgrounds (SG₁ and SG₂ in Table 6).

[Table 6 about here]

Additional data are recorded in Table 7, where marital fertility is interpreted as a determinant of both occupational and economic statuses at panel III. We see that the partial regression coefficients for the net effect of fertility on occupation are positive for urban subgroups 1 (.19) and 2 (.20) and negative for the farm or rural subgroup (-.08), although the latter is not significant. From both the small increments to the R² values and the lack of much change in the pattern of other coefficients, when compared to corresponding values in Tables 3 - 5, we would conclude that the net role of fertility in occupational achievements is minor. No statistically significant net effects of fertility are observed in the data for economic attainments.

While the net impact of marital fertility on socioeconomic achievement is not large, the direct effects on occupation do indicate some interaction for the residential background subgroups. Both Tables 6 and 7 imply that additional children

6. For this sample size, a correlation of about .10 is significant at the .05 level of confidence.

[Table 7 about here]

mean quite different things in the occupational careers of metropolitan males who were reared in urban as opposed to farm or rural areas. For those reared in urban contexts, additional children induce slightly greater occupational achievement than otherwise expected. At the same time, the sociological and demographic indicators of social background and early career achievements (Table 6) are virtually unrelated to marital fertility up to panel III. Among men reared on farms or in rural contexts but currently living in metropolitan areas, fertility is greater (especially) for those with fewer years of schooling and lower incomes at panels I and II. While the zero-order relationships are also negative with panel III statuses, the net effects of the latter are not interpretable, after controlling for social background and prior achievements.

How might we account for these interactions involving fertility and socio-economic achievements, especially for the positive net effect among the urban subgroup? Present tabulations do not suggest an answer, but we can rule out some possibilities. The differential impact of marital fertility on the career does not appear to be attributable to differential fertility itself, as suggested by the following means and standard deviations for the two urban (SG_1 and SG_2) and the rural or farm (SG_3) subgroups:

	SG_1	SG_2	SG_3
Means	3.2	3.0	3.3
S.D.	1.3	1.1	1.2

Nor does the timing or spacing of children seem to account for our results, since the correlations between the residential context variable (F) and the first and

second birth intervals and between variable F and age at marriage and duration of marriage are all non-significant (in the range .07 to -.01).

We are left with the fact that the impact of marital fertility on the socioeconomic careers of metropolitan males is different for men reared in urban contexts and for men reared in farm or rural areas. Although each subgroup sires nearly equal numbers of children at nearly the same rate at comparable points in the life cycle, the meaning of offspring is different for these residential background subgroups and results in a net positive influence in the careers of urban males.

Summary

We have observed similar details about the permeability of the American occupational structure and the process of stratification to those reported by Blau and Duncan (1967) in cross-sectional data. In longitudinal data referring to a portion of the white metropolitan male population between 1957-67, we have examined the processes of educational, occupational, and economic achievement, paying particular attention both to the permeability of urban stratification systems and to the processes of achievement for urban subgroups from different residential contexts of rearing.

Education is the key variable in the process of stratification. It functions both as an intervening variable whereby social origins and social destinations are connected and as the most major single influence (of those included in our models) on the socioeconomic career. It is differential education which appears to explain the residential background variation in achievement; educational completion, in turn, seems to hinge on both paternal occupation and the size of

the family of orientation. From both lower paternal status and greater numbers of siblings, the men with farm or rural rearing suffer handicaps in the educational component of occupational achievement.

In a brief analysis of potential interactions, we conclude rather tentatively that the status and structure of the family of origin may affect educational and middle-career economic achievements (respectively) differentially, on the basis of one's residential rearing context. Moreover, the structure of the family of procreation (marital fertility) also appears to influence socio-economic achievement in ways characteristic of residential context of rearing subgroups. Further research on the effects of the structure of the family (of socialization and of procreation) on the process of stratification for urban and rural men seems desirable.

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TABLE 1

Residential Context of Rearing

Score	Category	No. of Men with Fathers in Farm Occupations	Subset Total
1	Neither husband's parents reared in rural area; husband not reared in rural or farm area	2	274
2	Either husband's father or mother reared in rural area; husband not reared in rural area	0	60
3	Both parents reared in rural area; husband not reared in rural area	0	176
4	Neither parent reared in rural area; husband reared in rural area	0	13
5	Either parent reared in rural area; husband reared in rural area	2	20
6	Both parents reared in rural area; husband also reared in rural area	22	140
NR		0	32
TOTAL		26	715

TABLE 2

Partial Regression (Beta) Coefficients in Standard Form for Educational, Occupational, and Economic Achievement at Four Points in the Life Cycle on Three Background Variables and Antecedent Achievements for the PFS Subset, 1957-1967

Independent Variables	Dependent Variables (see stub)									
	U	W	IM	Y ₁	I ₁	Y ₂	I ₂	Y ₃	I ₃	
Father's Occ-NORC	X	.302	.098	.030 ^a	.077	.003 ^a	.011 ^a	.016 ^a	.084	.015 ^a
No. Siblings	S	-.236	-.049 ^a	.021 ^a	-.028 ^a	-.015 ^a	.012 ^a	.021 ^a	-.042 ^a	-.028 ^a
Farm-Rural Background	F	-.044 ^a	.032 ^a	-.058 ^a	.022 ^a	.050 ^a	-.040 ^a	.017 ^a	.022 ^a	-.016 ^a
Son's Education	U552	.009 ^a	.291	.122	.235	.132	.113	.127
Occ-NORC, Marr.	W209	.541	.137	.106	.068 ^a	.042 ^a	-.023 ^a
Income, Marr.	IM	-.060	.283	-.054	.025 ^a	-.002 ^a	.052 ^a
Occ-NORC, I	Y ₁128	.539	.048 ^a	.196	.058 ^a
Income, I	I ₁035 ^a	.491	-.032 ^a	.089
Occ-NORC, II	Y ₂122	.442	.000 ^a
Income, II	I ₂046 ^a	.271
Occ-NORC, III	Y ₃262
Income, III	I ₃
Coefficient of Determination	R ²	.176	.365	.053	.604	.232	.650	.476	.600	.410

^a Coefficient less than twice its standard error.

Table 3

Partial Regression Coefficients for Career Achievements on Background
and Antecedent Achievement Variables: Men Reared in Urban Areas

Independent Variables	Education	Occ-NORC Marr.	Income Marr.	Occ-NORC I	Income I	Occ-NORC II	Income II	Occ-NORC III	Income III
Father's Occ-NORC	.46	.16	49. ^a	.10	-24. ^a	-.00 ^a	202. ^a	.06 ^a	-97. ^a
No. Siblings	-.27	-.10 ^a	-12. ^a	-.07 ^a	35. ^a	.03 ^a	51. ^a	-.08 ^a	96. ^a
Education43	87. ^a	.26	268.	.18	210.	.10	461.
Occ-NORC, Marr.	144.	.44	248. ^a	.12	-45. ^a	.03 ^a	226. ^a
Income, Marr.	-.00 ^a	1.	-.00 ^a	0.	.00 ^a	0.
Occ-NORC, I	49. ^a	.52	-52. ^a	.19	69. ^a
Income I	-.00 ^a	1.	-.00 ^a	-0. ^a
Occ-NORC II	416.	.49	-45. ^a
Income II00 ^a	0.
Occ-NORC III	802.
Coefficient of Determination	.18	.36	.07	.58	.31	.62	.57	.65	.43

^a Coefficient less than twice absolute size of standard error.

Table 4

Partial Regression Coefficients for Career Achievements on Background
and Antecedent Achievement Variables: Men Reared in Urban Areas
and Parent(s) in Farm or Rural Areas

Independent Variables	Education	Occ-NORC Marr.	Income Marr.	Occ-NORC I	Income I	Occ-NORC II	Income II	Occ-NORC III	Income III
Father's Occ-Norc	.36	.04 ^a	53. a	.04 ^a	-55. a	.03a	142. a	.05 ^a	-29. a
No. Siblings	-.34	-.04 ^a	-20. a	-.09 ^a	79. a	-.03 ^a	20. a	-.04 ^a	56. a
Education45	-119. a	.20	156. a	.12	195. a	.10	507.
Occ-NORC, Marr.	252.	.52	230. a	.01 ^a	148. a	.10 ^a	-219. a
Income, Marr.	-.00 ^a	0.	-.00 ^a	0. a	.00 ^a	0.
Occ-NORC I	445.	.60	266. a	.16 ^a	56. a
Income I00 ^a	1.	-.00	0. a
Occ-NORC II	211. a	.36	169. a
Income II00	0.
Occ-NORC III	647.
Coefficient of Determination	.19	.36	.04	.59	.23	.65	.45	.56	.40

^aCoefficient less than twice absolute size of standard error.

TABLE 5

Partial Regression Coefficients for Career Achievements on Background
and Achievement Variables: Men Reared in Farm or Rural Areas

Independent Variables	Education	Occ-NORC Marr.	Income Marr.	Occ-NORC I	Income I	Occ-NORC II	Income II	Occ-NORC III	Income III
Father's Occ-NORC	.28	.10 ^a	8. ^a	.08	70. ^a	.02 ^a	-237. ^a	.09	152. ^a
No. Siblings	-.33	-.02 ^a	104. ^a	.04 ^a	-140. ^a	.05 ^a	50. ^a	-.00 ^a	-476.
Education49	22. ^a	.19	-14. ^a	.25	346.	.07 ^a	131. ^a
Occ-NORC, Marr.	163.	.66	216. ^a	.17	397. ^a	-.02 ^a	-661.
Income, Marr.	-.00 ^a	0.	-.00 ^a	-0. ^a	-.00 ^a	0. ^a
Occ-NORC I	Y ₁	328. ^a	.39	144. ^a	.26	955.
Income I	I ₁00 ^a	1.	.00 ^a	0. ^a
Occ-NORC II	Y ₂	198. ^a	.51	-97. ^a
Income II	I ₂	-.00 ^a	0.
Occ-NORC III	Y ₃	800.
Coefficient of Determination	R ²	.38	.04	.70	.19	.70	.44	.63	.50

^a Coefficient less than twice absolute size of standard error.

TABLE 6

Correlation Coefficients Between
Cumulative Fertility at Panel III and
Selected Background and Achievement Variables
for Residential Background Subgroups

Variables		Residential Backgrounds Subgroups			Total
		SG ₁ ^a	SG ₂	SG ₃	
Father's Occ-NORC	X	-.04	-.00	-.10	-.04
No. Siblings	S	.02	-.00	.14	.06
Education	U	.04	-.01	-.21	-.04
Occ-NORC, Marriage	W	-.02	-.08	-.08	-.05
Income, Marriage	IM	-.00	.03	-.04	-.00
Occ-NORC I	Y ₁	.00	-.06	-.13	-.05
Income I	I ₁	-.09	-.11	-.20	-.13
Occ-NORC II	Y ₂	-.00	.02	-.11	-.02
Income II	I ₂	-.05	.04	-.21	-.07
Occ-NORC III	Y ₃	.11	.11	-.14	.05
Income III	I ₃	.05	.06	-.15	.01

^a SG₁ Men reared in urban areas

SG₂ Men reared in urban areas, but parent(s) reared in farm or rural areas

SG₃ Men reared in farm or rural areas

TABLE 7

Partial Regression Coefficients for Effects of Background,
 Prior Achievement, and Cumulative Fertility Variables on Panel III
 Socioeconomic Achievements for PFS Subset and Residential Background

Subgroups

Independent Variables	Dependent Variables							
	Occ-NORC III			Income III				
	TOTAL	SG ₁ ^a	SG ₂	TOTAL	SG ₁ ^a	SG ₂		
Father's Occ-NORC	.08 ^a	.06 ^a	.06 ^a	.09	47. ^a	-91. ^a	-25. ^a	156. ^a
No. Siblings	-.04 ^a	-.08	-.04 ^a	.00 ^a	-97. ^a	91. ^a	55. ^a	-484.
Education	.09	.09	.10	.07 ^a	370.	456.	507.	149. ^a
Occ-NORC, Marriage	.04 ^a	.03 ^a	.12 ^a	-.01 ^a	-106. ^a	228. ^a	-202. ^a	-680.
Income, Marriage	-.00 ^a	.00 ^a	-.00 ^a	-.00 ^a	0.	0.	0. ^a	0.
Occ-NORC I	.21	.19	.18	.25	282. ^a	72. ^a	77. ^a	964.
Income I	-.00 ^a	-.00 ^a	-.00	.00 ^a	0.	-0. ^a	0.	0.
Occ-NORC II	.44	.50	.34	.51	14. ^a	-29. ^a	156. ^a	-116. ^a
Income II	.00 ^a	.00 ^a	.00	-.00 ^a	0.	0.	0.	0.
Cumulative Fertility	.14	.19	.20	-.08 ^a	192. ^a	136. ^a	230. ^a	204. ^a
Occ-NORC III	741.	775.	620.	811.
Coefficient of Determination	.61	.66	.57	.63	.42	.43	.40	.50

^a SG₁ Men reared in urban areas

SG₂ Men reared in urban areas, but parent(s) reared in farm or rural areas

SG₃ Men reared in farm or rural areas

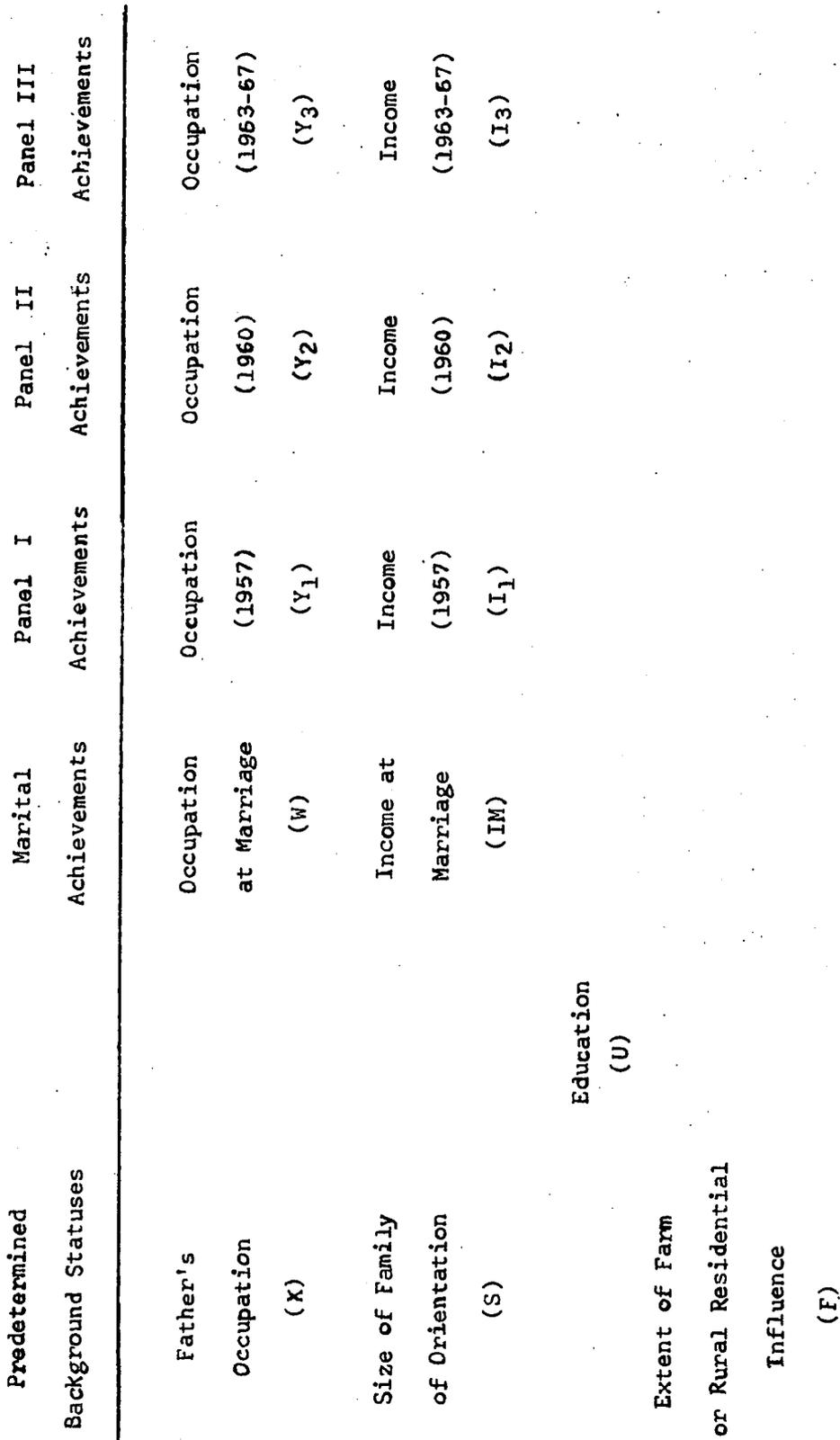


Figure 1--Processual Model of the Achievements in the
Careers of a Subset of Metropolitan Males

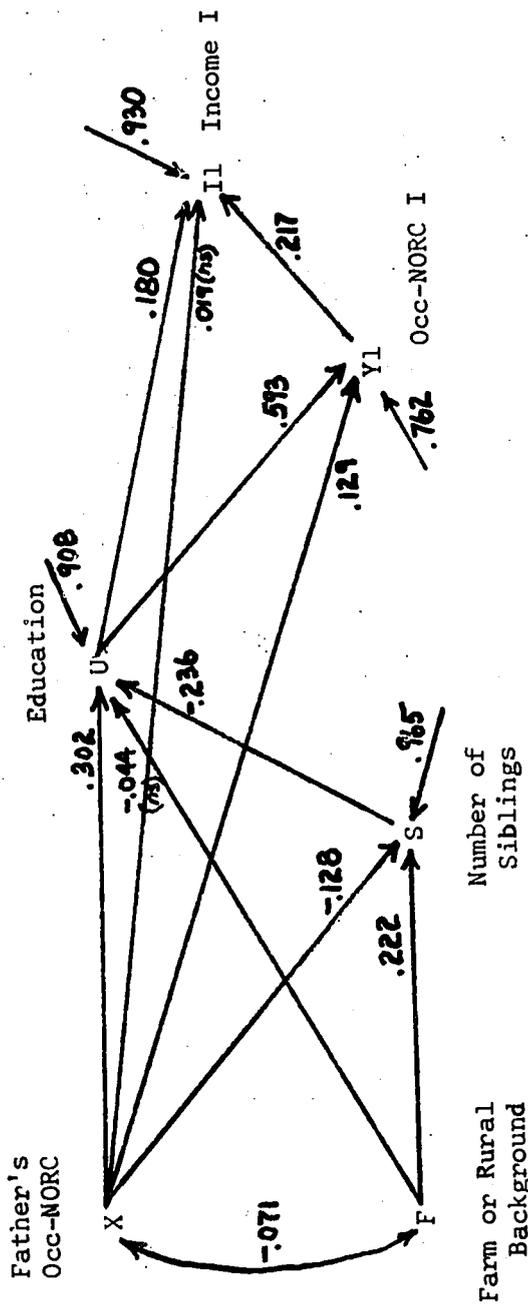


Figure 2--Representation of the Hypothesis of Indirect Effects of Size of Family of Orientation, Extent of Farm or Rural Background, and Paternal Occupation on Socioeconomic Achievement at Panel I

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